# Non-degenerate Hilbert cubes in random sets 

par CsABA SÁNDOR


#### Abstract

Résumé. Une légère modification de la démonstration du lemme des cubes de Szemerédi donne le résultat plus précis suivant: si une partie $S$ de $\{1, \ldots, n\}$ vérifie $|S| \geq \frac{n}{2}$, alors $S$ contient un cube de Hilbert non dégénéré de dimension $\left\lfloor\log _{2} \log _{2} n-3\right\rfloor$. Dans cet article nous montrons que dans un ensemble aléatoire avec les probabilités $\operatorname{Pr}\{s \in S\}=1 / 2$ indépendantes pour $1 \leq s \leq n$, la plus grande dimension d'un cube de Hilbert non dégénéré est proche de $\log _{2} \log _{2} n+\log _{2} \log _{2} \log _{2} n$ presque sûrement et nous déterminons la fonction seuil pour avoir un $k$-cube non dégénéré.


Abstract. A slight modification of the proof of Szemerédi's cube lemma gives that if a set $S \subset[1, n]$ satisfies $|S| \geq \frac{n}{2}$, then $S$ must contain a non-degenerate Hilbert cube of dimension $\left\lfloor\log _{2} \log _{2} n\right.$ $3\rfloor$. In this paper we prove that in a random set $S$ determined by $\operatorname{Pr}\{s \in S\}=\frac{1}{2}$ for $1 \leq s \leq n$, the maximal dimension of nondegenerate Hilbert cubes is a.e. nearly $\log _{2} \log _{2} n+\log _{2} \log _{2} \log _{2} n$ and determine the threshold function for a non-degenerate $k$-cube.

## 1. Introduction

Throughout this paper we use the following notations: let $[1, n]$ denote the first $n$ positive integers. The coordinates of the vector $\mathbf{A}^{(k, n)}=$ $\left(a_{0}, a_{1}, \ldots, a_{k}\right)$ are selected from the positive integers such that $\sum_{i=0}^{k} a_{i} \leq$ $n$. The vectors $\mathbf{B}^{(k, n)}, \mathbf{A}_{\mathbf{i}}^{(k, n)}$ are interpreted similarly. The set $S_{n}$ is a subset of $[1, n]$. The notations $f(n)=o(g(n))$ means $\lim _{n \rightarrow \infty} \frac{f(n)}{g(n)}=0$. An arithmetic progression of length $k$ is denoted by $A P_{k}$. The rank of a matrix $A$ over the field $\mathbb{F}$ is denoted by $r_{\mathbb{F}}(A)$. Let $\mathbb{R}$ denote the set of real numbers, and let $\mathbb{F}_{2}$ be the finite field of order 2.

Let $n$ be a positive integer, $0 \leq p_{n} \leq 1$. The random set $S\left(n, p_{n}\right)$ is the random variable taking its values in the set of subsets of $[1, n]$ with the law determined by the independence of the events $\left\{k \in S\left(n, p_{n}\right)\right\}, 1 \leq k \leq n$ with the probability $\operatorname{Pr}\left\{k \in S\left(n, p_{n}\right)\right\}=p_{n}$. This model is often used for

[^0]proving the existence of certain sequences. Given any combinatorial number theoretic property $P$, there is a probability that $S\left(n, p_{n}\right)$ satisfies $P$, which we write $\operatorname{Pr}\left\{S\left(n, p_{n}\right) \models P\right)$. The function $r(n)$ is called a threshold function for a combinatorial number theoretic property $P$ if
(i) When $p_{n}=o(r(n)), \lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, p_{n}\right) \mid=P\right\}=0$,
(ii) When $r(n)=o(p(n)), \lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, p_{n}\right) \mid=P\right\}=1$, or visa versa. It is clear that threshold functions are not unique. However, threshold functions are unique within factors $m(n), 0<\liminf _{n \rightarrow \infty} m(n) \leq$ $\limsup _{n \rightarrow \infty} m(n)<\infty$, that is if $p_{n}$ is a threshold function for $P$ then $p_{n}^{\prime}$ is also a threshold function iff $p_{n}=O\left(p_{n}^{\prime}\right)$ and $p_{n}^{\prime}=O\left(p_{n}\right)$. In this sense we can speak of the threshold function of a property.

We call $H \subset[1, n]$ a Hilbert cube of dimension $k$ or, simply, a $k$-cube if there is a vector $\mathbf{A}^{(k, n)}$ such that

$$
H=\mathbf{H}_{\mathbf{A}^{(k, n)}}=\left\{a_{0}+\sum_{i=1}^{k} \epsilon_{i} a_{i}: \epsilon_{i} \in\{0,1\}\right\}
$$

The positive integers $a_{1}, \ldots, a_{k}$ are called the generating elements of the Hilbert cube. The $k$-cube is non-degenerate if $|H|=2^{k}$ i.e. the vertices of the cube are distinct, otherwise it is called degenerate. The maximal dimension of a non-degenerate Hilbert cube in $S_{n}$ is denoted by $H_{\max }\left(S_{n}\right)$, i.e. $H_{\max }\left(S_{n}\right)$ is the largest integer $l$ such that there exists a vector $\mathbf{A}^{(l, n)}$ for which the non-degenerate Hilbert cube $\mathbf{H}_{\mathbf{A}^{(l, n)}} \subset S_{n}$.

Hilbert originally proved that if the positive integers are colored with finitely many colors then one color class contains a $k$-cube. The density version of theorem was proved by Szemerédi and has since become known as "Szemerédi's cube lemma". The best known result is due to Gunderson and Rödl (see [3]):

Theorem 1.1 (Gunderson and Rödl). For every $d \geq 3$ there exists $n_{0} \leq$ $\left(2^{d}-2 / \ln 2\right)^{2}$ so that, for every $n \geq n_{0}$, if $A \subset[1, n]$ satisfies $|A| \geq$ $2 n^{1-\frac{1}{2^{d-1}}}$, then $A$ contains a d-cube.

A direct consequence is the following:
Corollary 1.2. Every subset $S_{n}$ such that $\left|S_{n}\right| \geq \frac{n}{2}$ contains a $\left\lfloor\log _{2} \log _{2} n\right\rfloor$ cube.

A slight modification of the proof gives that the above set $S_{n}$ must contain a non-degenerate $\left\lfloor\log _{2} \log _{2} n-3\right\rfloor$-cube.

Obviously, a sequence $S$ has the Sidon property (that is the sums $s_{i}+$ $s_{j}, s_{i} \leq s_{j}, s_{i}, s_{j} \in S$ are distinct) iff $S$ contains no 2-cube. Godbole, Janson, Locantore and Rapoport studied the threshold function for the Sidon property and gave the exact probability distribution in 1999 (see [2]):

Theorem 1.3 (Godbole, Janson, Locantore and Rapoport). Let $c>0$ be arbitrary. Let $P$ be the Sidon property. Then with $p_{n}=c n^{-3 / 4}$,

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, p_{n}\right) \models P\right\}=e^{-\frac{c^{4}}{12}}
$$

Clearly, a subset $H \subset[1, n]$ is a degenerate 2-cube iff it is an $A P_{3}$. Moreover, an easy argument gives that the threshold function for the event " $A P_{3}$-free" is $p_{n}=n^{-2 / 3}$. Hence

Corollary 1.4. Let $c>0$ be arbitrary. Then with $p_{n}=c n^{-3 / 4}$,

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, p_{n}\right) \text { contains no non-degenerate } 2 \text {-cube }\right\}=e^{-\frac{c^{4}}{12}}
$$

In Theorem 1.5 we extend the previous Corollary.
Theorem 1.5. For any real number $c>0$ and any integer $k \geq 2$, if $p_{n}=c n^{-\frac{k+1}{2^{k}}}$,
$\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, p_{n}\right)\right.$ contains no non-degenerate $k$-cube $\}=e^{-\frac{c^{2^{k}}}{(k+1)!k!}}$.
In the following we shall find bounds on the maximal dimension of nondegenerate Hilbert cubes in the random set $S\left(n, \frac{1}{2}\right)$. Let

$$
D_{n}(\epsilon)=\left\lfloor\log _{2} \log _{2} n+\log _{2} \log _{2} \log _{2} n+\frac{(1-\epsilon) \log _{2} \log _{2} \log _{2} n}{\log 2 \log _{2} \log _{2} n}\right\rfloor
$$

and

$$
E_{n}(\epsilon)=\left[\log _{2} \log _{2} n+\log _{2} \log _{2} \log _{2} n+\frac{(1+\epsilon) \log _{2} \log _{2} \log _{2} n}{\log 2 \log _{2} \log _{2} n}\right\rfloor
$$

The next theorem implies that for almost all $n, H_{\max }\left(S\left(n, \frac{1}{2}\right)\right)$ concentrates on a single value because for every $\epsilon>0, D_{n}(\epsilon)=E_{n}(\epsilon)$ except for a sequence of zero density.

Theorem 1.6. For every $\epsilon>0$

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{D_{n}(\epsilon) \leq H_{\max }\left(S\left(n, \frac{1}{2}\right)\right) \leq E_{n}(\epsilon)\right\}=1 .
$$

## 2. Proofs

In order to prove the theorems we need some lemmas.
Lemma 2.1. For $k_{n}=o\left(\frac{\log n}{\log \log n}\right)$ the number of non-degenerate $k_{n}$-cubes in $[1, n]$ is $(1+o(1))\binom{n}{k_{n}+1} \frac{1}{k_{n}!}$, as $n \rightarrow \infty$.

Proof. All vectors $\mathbf{A}^{\left(k_{n}, n\right)}$ are in 1-1 correspondence with all vectors $\left(v_{0}, v_{1}\right.$, $\ldots, v_{k_{n}}$ ) with $1 \leq v_{1}<v_{2}<\cdots<v_{k_{n}} \leq n$ in $\mathbb{R}^{k_{n}+1}$ according to the formulas $\left(a_{0}, a_{1}, \ldots, a_{k_{n}}\right) \mapsto\left(v_{0}, v_{1}, \ldots, v_{k_{n}}\right)=\left(a_{0}, a_{0}+a_{1}, \ldots, a_{0}+a_{1}+\right.$ $\left.\cdots+a_{k_{n}}\right)$; and $\left(v_{0}, v_{1}, \ldots, v_{k_{n}}\right) \mapsto\left(a_{0}, a_{1}, \ldots, a_{k_{n}}\right)=\left(v_{0}, v_{1}-v_{0}, \ldots, v_{k_{n}}-\right.$ $\left.v_{k_{n}-1}\right)$. Consequently,

$$
\begin{aligned}
\binom{n}{k_{n}+1}= & \mid\left\{\mathbf{A}^{\left(k_{n}, n\right)}: \mathbf{H}_{\mathbf{A}^{\left(k_{n}, n\right)}} \text { is non-degenerate }\right\} \mid \\
& +\mid\left\{\mathbf{A}^{\left(k_{n}, n\right)}: \mathbf{H}_{\mathbf{A}^{\left(k_{n}, n\right)}} \text { is degenerate }\right\} \mid
\end{aligned}
$$

By the definition of a non-degenerate cube the cardinality of the set $\left\{\mathbf{A}^{\left(k_{n}, n\right)}: \mathbf{H}_{\mathbf{A}^{\left(k_{n}, n\right)}}\right.$ is non-degenerate $\}$ is equal to

$$
k_{n}!\mid\left\{\text { non-degenerate } k_{n} \text {-cubes in }[1, n]\right\} \mid,
$$

because permutations of $a_{1}, \ldots, a_{k}$ give the same $k_{n}$-cube. It remains to verify that the number of vectors $\mathbf{A}^{\left(k_{n}, n\right)}$ which generate degenerate $k_{n^{-}}$ cubes is $o\left(\binom{n}{k_{n}+1}\right)$. Let $\mathbf{A}^{\left(k_{n}, n\right)}$ be a vector for which $\mathbf{H}_{\mathbf{A}^{\left(k_{n}, n\right)}}$ is a degenerate $k_{n}$-cube. Then there exist integers $1 \leq u_{1}<u_{2}<\ldots<u_{s} \leq k_{n}, 1 \leq v_{1}<$ $v_{2}<\ldots<v_{t} \leq k_{n}$ such that

$$
a_{0}+a_{u_{1}}+\ldots+a_{u_{s}}=a_{0}+a_{v_{1}}+\ldots+a_{v_{t}}
$$

where we may assume that the indices are distinct, therefore $s+t \leq k_{n}$. Then the equation

$$
x_{1}+x_{2}+\ldots+x_{s}-x_{s+1}-\ldots-x_{s+t}=0
$$

can be solved over the set $\left\{a_{1}, a_{2} \ldots, a_{k_{n}}\right\}$. The above equation has at most $n^{s+t-1} \leq n^{k_{n}-1}$ solutions over $[1, n]$. Since we have at most $k_{n}^{2}$ possibilities for $(s, t)$ and at most $n$ possibilities for $a_{0}$, therefore the number of vectors $\mathbf{A}^{\left(k_{n}, n\right)}$ for which $\mathbf{H}_{\mathbf{A}^{\left(k_{n}, n\right)}}$ is degenerate is at most $k_{n}^{2} n^{k_{n}}=o\left(\binom{n}{k_{n}+1}\right)$.

In the remaining part of this section the Hilbert cubes are nondegenerate.

The proofs of Theorem 1.5 and 1.6 will be based on the following definition. For two intersecting $k$-cubes $\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}$ let $\mathbf{H}_{\mathbf{A}^{(k, n)}} \cap \mathbf{H}_{\mathbf{B}^{(k, n)}}=$ $\left\{c_{1}, \ldots, c_{m}\right\}$ with $c_{1}<\ldots<c_{m}$, where

$$
c_{d}=a_{0}+\sum_{l=1}^{k} \alpha_{d, l} a_{l}=b_{0}+\sum_{l=1}^{k} \beta_{d, l} b_{l}, \quad \alpha_{d, l}, \beta_{d, l} \in\{0,1\}
$$

for $1 \leq d \leq m$ and $1 \leq l \leq k$. The rank of the intersection of two $k$-cubes $\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}$ is defined as follows: we say that $r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(\mathrm{s}, \mathrm{t})$ if for the matrices $A=\left(\alpha_{d, l}\right)_{m \times k}, B=\left(\beta_{d, l}\right)_{m \times k}$ we have $r_{\mathbb{R}}(A)=s$ and $r_{\mathbb{R}}(B)=t$. The matrices $A$ and $B$ are called matrices of the common vertices of $\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}$.

Lemma 2.2. The condition $r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(s, t)$ implies that

$$
\left|\mathbf{H}_{\mathbf{A}^{(k, n)}} \cap \mathbf{H}_{\mathbf{B}^{(k, n)}}\right| \leq 2^{\min \{s, t\}}
$$

Proof. We may assume that $s \leq t$. The inequality $\left|\mathbf{H}_{\mathbf{A}^{(k, n)}} \cap \mathbf{H}_{\mathbf{B}^{(k, n)}}\right| \leq 2^{s}$ is obviously true for $s=k$. Let us suppose that $s<k$ and the number of common vertices is greater than $2^{s}$. Then the corresponding ( $0-1$ )-matrices $A$ and $B$ have more than $2^{s}$ different rows, therefore $r_{\mathbb{F}_{2}}(A)>s$, but we know from elementary linear algebra that for an arbitrary $(0-1)$-matrix $M$ we have $r_{\mathbb{F}_{2}}(M) \geq r_{\mathbb{R}}(M)$, which is a contradiction.
Lemma 2.3. Suppose that the sequences $\mathbf{A}^{(k, n)}$ and $\mathbf{B}^{(k, n)}$ generate nondegenerate $k$-cubes. Then
(1) $\left|\left\{\left(\mathbf{A}^{(k, n)}, \mathbf{B}^{(k, n)}\right): r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(s, t)\right\}\right|$

$$
\leq 2^{2 k^{2}}\binom{n}{k+1} n^{k+1-\max \{s, t\}}
$$

for all $0 \leq s, t \leq k$;
(2) $\mid\left\{\left(\mathbf{A}^{(k, n)}, \mathbf{B}^{(k, n)}\right): r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(r, r),\left|\mathbf{H}_{\mathbf{A}^{(k, n)}} \cap \mathbf{H}_{\mathbf{B}^{(k, n)}}\right|=\right.$ $\left.2^{r}\right\} \left\lvert\, \leq 2^{2 k^{2}}\binom{n}{k+1} n^{k-r}\right.$
for all $0 \leq r<k$;
(3) $\mid\left\{\left(\mathbf{A}^{(k, n)}, \mathbf{B}^{(k, n)}\right): r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(k, k),\left|\mathbf{H}_{\mathbf{A}^{(k, n)}} \cap \mathbf{H}_{\mathbf{B}^{(k, n)}}\right|>\right.$ $\left.2^{k-1}\right\} \left\lvert\, \leq 2^{2 k^{2}+2 k}\binom{n}{k+1}\right.$.
Proof. (1) We may assume that $s \leq t$. In this case we have to prove that the number of corresponding pairs $\left(\mathbf{A}^{(k, n)}, \mathbf{B}^{(k, n)}\right)$ is at most $\binom{n}{k+1} 2^{2 k^{2}} n^{k+1-t}$. We have already seen in the proof of Lemma 2.1 that the number of vectors $\mathbf{A}^{(k, n)}$ is at most $\binom{n}{k+1}$. Fix a vector $\mathbf{A}^{(k, n)}$ and count the suitable vectors $\mathbf{B}^{(k, n)}$. Then the matrix $B$ has $t$ linearly independent rows, namely $r_{\mathbb{R}}\left(\left(\beta_{d_{i}, l}\right)_{t \times k}\right)=t$, for some $1 \leq d_{1}<\cdots<d_{t} \leq m$, where

$$
a_{0}+\sum_{l=1}^{k} \alpha_{d_{i}, l} a_{l}=b_{0}+\sum_{l=1}^{k} \beta_{d_{i}, l} b_{l}, \quad \alpha_{d_{i}, l}, \beta_{d_{i}, l} \in\{0,1\} \quad \text { for } 1 \leq i \leq t .
$$

The number of possible $b_{0}$ s is at most $n$. For fixed $b_{0}, \alpha_{d_{i}, l}, \beta_{d_{i}, l}$ let us study the system of equations

$$
a_{0}+\sum_{l=1}^{k} \alpha_{d_{i}, l} a_{l}=b_{0}+\sum_{l=1}^{k} \beta_{d_{i}, l} x_{l}, \quad \alpha_{d_{i}, l}, \beta_{d_{i}, l} \in\{0,1\} \quad \text { for } 1 \leq i \leq t .
$$

The assumption $r_{\mathbb{R}}\left(\beta_{d_{i}, l}\right)_{t \times k}=t$ implies that the number of solutions over $[1, n]$ is at most $n^{k-t}$. Finally, we have at most $2^{k t}$ possibilities on the left-hand side for $\alpha_{d_{i}, l}$ s and, similarly, we have at most $2^{k t}$ possibilities on the right-hand side for $\beta_{d_{i}, l} \mathrm{~s}$, therefore the number of possible systems of equations is at most $2^{2 k^{2}}$.
(2) The number of vectors $\mathbf{A}^{(k, n)}$ is $\binom{n}{k+1}$ as in (1). Fix a vector $\mathbf{A}^{(k, n)}$ and count the suitable vectors $\mathbf{B}^{(k, n)}$. It follows from the assumptions
$r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(r, r),\left|\mathbf{H}_{\mathbf{A}^{(k)}} \cap \mathbf{H}_{\mathbf{B}^{(k)}}\right|=2^{r}$ that the vectors $\left(\alpha_{d, 1}, \ldots\right.$, $\left.\alpha_{d, k}\right), d=1, \ldots, 2^{r}$ and the vectors $\left(\beta_{d, 1}, \ldots, \beta_{d, k}\right), d=1, \ldots, 2^{r}$, respectively form $r$-dimensional subspaces of $\mathbb{F}_{2}^{k}$. Considering the zero vectors of these subspaces we get $a_{0}=b_{0}$. The integers $b_{1}, \ldots, b_{k}$ are solutions of the system of equations

$$
a_{0}+\sum_{l=1}^{k} \alpha_{d, l} a_{l}=b_{0}+\sum_{l=1}^{k} \beta_{d, l} x_{l} \quad \alpha_{d, l}, \beta_{d, l} \in\{0,1\} \quad \text { for } 1 \leq d \leq 2^{r} .
$$

Similarly to the previous part this system of equation has at most $n^{k-r}$ solutions over $[1, n]$ and the number of choices for the $r$ linearly independent rows is at most $2^{2 k^{2}}$.
(3) Fix a vector $\mathbf{A}^{(k, n)}$. Let us suppose that for a vector $\mathbf{B}^{(k, n)}$ we have $r\left(\mathbf{H}_{\mathbf{A}^{(k, n)}}, \mathbf{H}_{\mathbf{B}^{(k, n)}}\right)=(k, k)$ and $\left|\mathbf{H}_{\mathbf{A}^{(k, n)}} \cap \mathbf{H}_{\mathbf{B}^{(k, n)}}\right|>2^{k-1}$. Let the common vertices be

$$
a_{0}+\sum_{l=1}^{k} \alpha_{d, l} a_{l}=b_{0}+\sum_{l=1}^{k} \beta_{d, l} b_{l}, \quad \alpha_{d, l}, \beta_{d, l} \in\{0,1\} \quad \text { for } 1 \leq d \leq m,
$$

where we may assume that the rows $d_{1}, \ldots, d_{k}$ are linearly independent, i.e. the matrix $B_{k}=\left(\beta_{d_{i}, l}\right)_{k \times k}$ is regular. Write the rows $d_{1}, \ldots, d_{k}$ in matrix form as

$$
\begin{equation*}
\underline{a}=b_{0} \underline{1}+B_{k} \underline{b}, \tag{1}
\end{equation*}
$$

with vectors $\underline{a}=\left(a_{0}+\sum_{l=1}^{k} \alpha_{d_{i}, l} a_{l}\right)_{k \times 1}, \underline{1}=(1)_{k \times 1}$ and $\underline{b}=\left(b_{i}\right)_{k \times 1}$. It follows from (1) that

$$
\underline{b}=B_{k}^{-1}\left(\underline{a}-b_{0} \underline{1}\right)=B_{k}^{-1} \underline{a}-b_{0} B_{k}^{-1} \underline{1} .
$$

Let $B_{k}^{-1} \underline{1}=\left(d_{i}\right)_{k \times 1}$ and $B_{k}^{-1} \underline{a}=\left(c_{i}\right)_{k \times 1}$. Obviously, the number of subsets $\left\{i_{1}, \ldots i_{l}\right\} \subset\{1, \ldots, k\}$ for which $d_{i_{1}}+\ldots+d_{i_{l}} \neq 1$ is at least $2^{k-1}$, therefore there exist $1 \leq u_{1}<\ldots<u_{s} \leq k$ and $1 \leq v_{1}<\ldots<v_{t} \leq k$ such that $a_{0}+a_{u_{1}}+\ldots+a_{u_{s}}=b_{0}+b_{v_{1}}+\ldots+b_{v_{t}}$, and $d_{v_{1}}+\ldots+d_{v_{t}} \neq 1$. Hence $a_{0}+a_{u_{1}}+\ldots+a_{u_{s}}=b_{0}+b_{v_{1}}+\ldots+b_{v_{t}}=b_{0}+c_{v_{1}}+\ldots+c_{v_{t}}-b_{0}\left(d_{v_{1}}+\ldots+d_{v_{t}}\right)$

$$
b_{0}=\frac{a_{0}+a_{u_{1}}+\ldots+a_{u_{s}}-c_{v_{1}}-\ldots-c_{v_{t}}}{1-\left(d_{v_{1}}+\ldots+d_{v_{t}}\right)}
$$

To conclude the proof we note that the number of sets $\left\{u_{1}, \ldots, u_{s}\right\}$ and $\left\{v_{1}, \ldots, v_{t}\right\}$ is at most $2^{2 k}$ and there are at most $2^{k^{2}}$ choices for $B_{k}$ and $\underline{a}$, respectively. Finally, for given $B_{k}, \underline{a}, b_{0}, 1 \leq u_{1}<\ldots<u_{s} \leq k$ and $1 \leq v_{1}<\ldots<v_{t} \leq k$, the vector $\mathbf{B}^{(k, n)}$ is determined uniquely.

In order to prove the theorems we need two lemmas from probability theory (see e.g. [1] p. 41, 95-98.). Let $X_{i}$ be the indicator function of the event $A_{i}$ and $S_{N}=X_{1}+\ldots+X_{N}$. For indices $i, j$ write $i \sim j$ if $i \neq j$ and
the events $A_{i}, A_{j}$ are dependant. We set $\Gamma=\sum_{i \sim j} \operatorname{Pr}\left\{A_{i} \cap A_{j}\right\}$ (the sum over ordered pairs).

Lemma 2.4. If $E\left(S_{n}\right) \rightarrow \infty$ and $\Gamma=o\left(E\left(S_{n}\right)^{2}\right)$, then $X>0$ a.e.
In many instances, we would like to bound the probability that none of the bad events $B_{i}, i \in I$, occur. If the events are mutually independent, then $\operatorname{Pr}\left\{\cap_{i \in I} \overline{B_{i}}\right\}=\prod_{i \in I} \operatorname{Pr}\left\{\overline{B_{i}}\right\}$. When the $B_{i}$ are "mostly" independent, the Janson's inequality allows us, sometimes, to say that these two quantities are "nearly" equal. Let $\Omega$ be a finite set and $R$ be a random subset of $\Omega$ given by $\operatorname{Pr}\{r \in R\}=p_{r}$, these events being mutually independent over $r \in \Omega$. Let $E_{i}, i \in I$ be subsets of $\Omega$, where $I$ a finite index set. Let $B_{i}$ be the event $E_{i} \subset R$. Let $X_{i}$ be the indicator random variable for $B_{i}$ and $X=\sum_{i \in I} X_{i}$ be the number of $E_{i} \mathrm{~s}$ contained in $R$. The event $\cap_{i \in I} \overline{B_{i}}$ and $X=0$ are then identical. For $i, j \in I$, we write $i \sim j$ if $i \neq j$ and $E_{i} \cap E_{j} \neq \emptyset$. We define $\Delta=\sum_{i \sim j} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\}$, here the sum is over ordered pairs. We set $M=\prod_{i \in I} \operatorname{Pr}\left\{\overline{B_{i}}\right\}$.

Lemma 2.5 (Janson's inequality). Let $\varepsilon \in] 0,1\left[\right.$ and let $B_{i}, i \in I, \Delta, M$ be as above and assume that $\operatorname{Pr}\left\{B_{i}\right\} \leq \varepsilon$ for all $i$. Then

$$
M \leq \operatorname{Pr}\left\{\cap_{i \in I} \overline{B_{i}}\right\} \leq M e^{\frac{1}{1-\varepsilon} \frac{\Delta}{2}}
$$

Proof of Theorem 1.5. Let $\mathbf{H}_{\mathbf{A}_{1}^{(k, n)}, \ldots, \mathbf{H}_{\mathbf{A}_{\mathrm{N}}^{(k, n)}} \text { be the distinct non-degen- }}$ erate $k$-cubes in $[1, n]$. Let $B_{i}$ be the event $\mathbf{H}_{\mathbf{A}_{\mathrm{i}}^{(k, n)}} \subset S\left(n, c n^{-\frac{k+1}{2^{k}}}\right)$. Then $\operatorname{Pr}\left\{B_{i}\right\}=c^{2^{k}} n^{-k-1}=o(1)$ and $N=(1+o(1))\binom{n}{k+1} \frac{1}{k!}$. It is enough to prove

$$
\Delta=\sum_{i \sim j} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\}=o(1)
$$

since then Janson's inequality implies

$$
\begin{aligned}
\operatorname{Pr}\left\{S\left(n, c n^{-\frac{k+1}{2^{k}}}\right)\right. & \text { does not contain any } k \text {-cubes }\} \\
& =\operatorname{Pr}\left\{\cap_{i=1}^{N} \overline{B_{i}}\right\} \\
& =(1+o(1))\left(1-\left(c n^{-\frac{k+1}{2^{k}}}\right)^{2^{k}}\right)^{(1+o(1))\binom{n}{k+1} \frac{1}{k!}} \\
& =(1+o(1)) e^{-\frac{c^{k}}{(k+1) \cdot k!}} .
\end{aligned}
$$

It remains to verify that $\sum_{i \sim j} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\}=o(1)$. We split this sum according to the ranks in the following way

$$
\begin{aligned}
& \sum_{i \sim j} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\}=\sum_{s=0}^{k} \sum_{t=0}^{k} \sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right)=(s, t)}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& =2 \sum_{s=1}^{k} \sum_{t=0}^{s-1} \sum_{i \sim j} \sum_{\substack{\left.i \sim j \\
\mathbf{A}_{\mathbf{i}}^{(k, n)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}(k, n)}\right)=(s, t)}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& +\sum_{r=0}^{k-1} \sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}(k, n)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right)=(r, r)}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& +\sum_{r=1}^{k-1} \sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}(k, n)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right)=(r, r)}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& \left|\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}} \cap \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right|<2^{r}
\end{aligned}
$$

The first sum can be estimated by Lemmas 2 and 2.3(1)

$$
\begin{aligned}
& \left.\sum_{s=1}^{k} \sum_{t=0}^{s-1} \sum_{i \sim j} \operatorname{Pr}_{r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right)=(s, t)} \operatorname{Pr} B_{i} \cap B_{j}\right\} \\
& \leq \sum_{s=1}^{k} \sum_{t=0}^{s-1} 2^{2 k^{2}}\binom{n}{k+1} n^{k+1-s}\left(c n^{-\frac{k+1}{2^{k}}}\right)^{2 \cdot 2^{k}-2^{t}} \\
& =n^{o(1)} \sum_{s=1}^{k} n^{2^{s-1} \frac{k+1}{2^{k}}-s}=n^{o(1)}\left(n^{\frac{k+1}{2^{k}}-1}+n^{\frac{k+1}{2}-k}\right)=o(1)
\end{aligned}
$$

since the sequence $a_{s}=2^{s-1} \frac{k+1}{2^{k}}-s$ is decreasing for $1 \leq s \leq k+1-$ $\log _{2}(k+1)$ and increasing for $k+1-\log _{2}(k+1)<s \leq k$.

To estimate the second sum we apply Lemma 2.3(2)

$$
\begin{aligned}
& \sum_{r=0}^{k-1} \sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}(k, n)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right)=(r, r) \\
\left|\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}} \cap \mathbf{H}_{\mathbf{A}_{\mathbf{j}}(k, n)}\right|=2^{r}}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& \leq \sum_{r=0}^{k-1} 2^{2 k^{2}}\binom{n}{k+1} n^{k-r}\left(c n^{-\frac{k+1}{2^{k}}}\right)^{2 \cdot 2^{k}-2^{r}} \\
& \quad=n^{-1+o(1)} \sum_{r=0}^{k-1} n^{2^{r} \frac{k+1}{2^{k}}-r}=n^{-1+o(1)}\left(n^{\frac{k+1}{2^{k}}}+n^{\frac{k+1}{2}-(k-1)}\right)=o(1) .
\end{aligned}
$$

The third sum can be bounded using Lemma 2.3(1):

$$
\begin{aligned}
& \sum_{r=1}^{k-1} \sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}}^{(k, n)}\right)=(r, r) \\
\left|\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}} \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right|<2^{r}}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& \leq \sum_{r=1}^{k-1} 2^{2 k^{2}}\binom{n}{k+1} n^{k+1-r}\left(c n^{\left.-\frac{k+1}{2^{k}}\right)^{2 \cdot 2^{k}-2^{r}+1}}\right. \\
& \quad \leq n^{o(1)-\frac{k+1}{2^{k}}} \sum_{r=1}^{k-1} n^{2^{r} \frac{k+1}{2^{k}-r}}=n^{o(1)-\frac{k+1}{2^{k}}}\left(n^{2 \frac{k+1}{2^{k}-1}}+n^{\frac{k+1}{2}-(k-1)}\right)=o(1)
\end{aligned}
$$

Similarly, for the fourth sum we apply Lemma 2.3(1)

$$
\sum_{\substack{i \sim j \\ r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}(k, n)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right)=(k, k) \\\left|\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}}^{(k)} \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right| \leq 2^{k-1}}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \leq n^{o(1)} n^{k+2}\left(c n^{-\frac{k+1}{2^{k}}}\right)^{1 \cdot 5 \cdot 2^{k}}=o(1) .
$$

To estimate the fifth sum we note that $\left|\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{(k, n)}} \cup \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{(k, n)}}\right| \geq 2^{k}+1$. It follows from Lemma 2.3(3) that

$$
\sum_{\substack{i \sim j \\ r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}(k, n)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}(k, n)}\right)=(k, k) \\\left|\mathbf{H}_{\mathbf{A}}(k, n) \cap \mathbf{H}_{\mathbf{\Delta}}(k, n)\right|>2^{k-1}}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \leq 2^{2 k^{2}+2 k} n^{k+1}\left(c n^{-\frac{k+1}{2^{k}}}\right)^{2^{k}+1}=o(1)
$$

which completes the proof.

Proof of Theorem 1.6. Let $\epsilon>0$ and for simplicity let $D_{n}=D_{n}(\epsilon)$ and $E_{n}=E_{n}(\epsilon)$. In the proof we use the estimations

$$
\begin{aligned}
2^{2^{D_{n}}} & \leq 2^{2^{\log _{2} \log _{2} n+\log _{2} \log _{2} \log _{2} n+\frac{(1-\epsilon) \log _{2} \log _{2} \log _{2} n}{\log _{2 \log _{2} \log _{2} n}}}}=n^{\log _{2} \log _{2} n+(1-\epsilon+o(1)) \log _{2} \log _{2} \log _{2} n}
\end{aligned}
$$

and

$$
\begin{aligned}
2^{2^{E_{n}+1}} & \geq 2^{2^{\log _{2} \log _{2} n+\log _{2} \log _{2} \log _{2} n+\frac{(1+\epsilon) \log _{2} \log _{2} \log _{2} n}{\log _{2} \log _{2} \log _{2} n}}} \\
& =n^{\log _{2} \log _{2} n+(1+\epsilon+o(1)) \log _{2} \log _{2} \log _{2} n}
\end{aligned}
$$

In order to verify Theorem 1.6 we have to show that

$$
\begin{equation*}
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, \frac{1}{2}\right) \text { contains a } D_{n} \text {-cube }\right\}=1 \tag{2}
\end{equation*}
$$

and

$$
\begin{equation*}
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, \frac{1}{2}\right) \text { contains an }\left(E_{n}+1\right) \text {-cube }\right\}=0 . \tag{3}
\end{equation*}
$$

To prove the limit in (4) let $\mathbf{H}_{\mathbf{A}_{1}^{\left(D_{n}, n\right)}, \ldots, \mathbf{H}_{\mathbf{A}_{\mathrm{N}}^{\left(D_{n}, n\right)}} \text { be the different non- }}$ degenerate $D_{n}$-cubes in $[1, n], B_{i}$ be the event $H_{\mathbf{A}_{\mathbf{i}}^{\left(D_{n}, n\right)}} \subset S\left(n, \frac{1}{2}\right), X_{i}$ be the indicator random variable for $B_{i}$ and $S_{N}=X_{1}+\ldots+X_{N}$ be the number of $\mathbf{H}_{\mathbf{A}_{\mathbf{i}}\left(D_{n}, n\right)} \subset S\left(n, \frac{1}{2}\right)$. The linearity of expectation gives by Lemma 2.1 and inequality (2)

$$
\begin{aligned}
E\left(S_{N}\right) & =N E\left(X_{i}\right)=(1+o(1))\binom{n}{D_{n}+1} \frac{1}{D_{n}!} 2^{-2^{D_{n}}} \\
& \geq n^{\log _{2} \log _{2} n+(1+o(1)) \log _{2} \log _{2} \log _{2} n} n^{-\log _{2} \log _{2} n-(1-\epsilon+o(1)) \log _{2} \log _{2} \log _{2} n} \\
& =n^{(\epsilon+o(1)) \log _{2} \log _{2} \log _{2} n} .
\end{aligned}
$$

Therefore $E\left(S_{N}\right) \rightarrow \infty$, as $n \rightarrow \infty$. By Lemma 2.4 it remains to prove that

$$
\sum_{i \sim j} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\}=o\left(E\left(S_{N}\right)^{2}\right)
$$

where $i \sim j$ means that the events $B_{i}, B_{j}$ are not independent i.e. the cubes $\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{\left(D_{n}, n\right)}}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{\left(D_{n}, n\right)}}$ have common vertices. We split this sum according to
the ranks

$$
\begin{aligned}
\sum_{i \sim j} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\}= & \sum_{s=0}^{D_{n}} \sum_{t=0}^{D_{n}} \sum_{r\left(\mathbf{H}_{\left.\mathbf{A}_{\mathbf{i}}^{\left(D_{n}, n\right)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}}^{\left(D_{n}, n\right)}\right)=(s, t)}\right.} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
\leq & \sum_{i \sim j}^{i\left(\mathbf{H}_{\left.\mathbf{A}_{\mathbf{i}}^{\left(D_{n}, n\right)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}}^{\left(D_{n}, n\right)}\right)=(0,0)}\right.} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& +2 \sum_{s=1}^{\sum_{n}} \sum_{t=0}^{s} \sum_{r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}\left(D_{n}, n\right)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}}^{\left(\mathbf{D}_{n}, n\right)}\right.} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} .
\end{aligned}
$$

The condition $r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{\left(D_{n}, n\right)}}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{\left(D_{n}, n\right)}}\right)=(0,0)$ implies that

$$
\left|\mathbf{H}_{\mathbf{A}_{\mathbf{i}}^{\left(D_{n}, n\right)}} \cup \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{\left(D_{n}, n\right)}}\right|=2^{D_{n}+1}-1,
$$

thus by Lemma 2.3(2)

$$
\begin{aligned}
\sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}}^{\left(D_{n}, n\right)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{\left(D_{n}, n\right)}}\right)=(0,0)}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} & \leq 2^{2 D_{n}^{2}}\binom{n}{D_{n}+1} n^{D_{n}} 2^{-2^{D_{N}+1}+1} \\
& =o\left(\left(\binom{n}{D_{n}+1} \frac{1}{D_{n}!} 2^{-2^{D_{n}}}\right)^{2}\right) \\
& =o\left(E\left(S_{N}\right)^{2}\right) .
\end{aligned}
$$

In the light of Lemmas 2 and 2.3(1) the second term in (6) can be estimated as

$$
\begin{aligned}
& \sum_{s=1}^{D_{n}} \sum_{t=0}^{s} \sum_{\substack{i \sim j \\
r\left(\mathbf{H}_{\mathbf{A}_{\mathbf{i}}}^{\left(D_{n}, n\right)}, \mathbf{H}_{\mathbf{A}_{\mathbf{j}}^{\left(D_{n}, n\right)}}\right)=(s, t)}} \operatorname{Pr}\left\{B_{i} \cap B_{j}\right\} \\
& \leq \sum_{s=1}^{D_{n}} \sum_{t=0}^{s}\binom{n}{D_{n}+1} 2^{2 D_{n}^{2}} n^{D_{n}+1-s} 2^{-2 \cdot 2^{D_{n}}+2^{t}} \\
& \quad=\left(\binom{n}{D_{n}+1} \frac{1}{D_{n}!} 2^{-2^{D_{n}}}\right)^{2} n^{o(1)} \sum_{s=1}^{D_{n}} \sum_{t=0}^{s} \frac{2^{2^{t}}}{n^{s}} \\
& \quad=\left(\binom{n}{D_{n}+1} \frac{1}{D_{n}!} 2^{-2^{D_{n}}}\right)^{2} n^{o(1)} \sum_{s=1}^{D_{n}} \frac{2^{2^{s}}}{n^{s}}
\end{aligned}
$$

Finally, the function $f(x)=\frac{2^{2^{x}}}{n^{x}}$ decreases on $\left(-\infty, \log _{2} \log n-2 \log _{2} \log 2\right]$ and increases on $\left[\log _{2} \log n-2 \log _{2} \log 2, \infty\right)$, therefore by (2)

$$
\sum_{s=1}^{D_{n}} \frac{2^{2^{s}}}{n^{r}}=n^{o(1)}\left(\frac{4}{n}+\frac{2^{2^{D_{n}}}}{n^{D_{n}}}\right)=n^{-1+o(1)}
$$

which proves the limit in (4).
In order to prove the limit in (5) let $\mathbf{H}_{\mathbf{C}_{\mathbf{1}}^{\left(E_{n}+1, n\right)}}, \ldots, \mathbf{H}_{\mathbf{C}_{\mathbf{K}}^{\left(E_{n}+1, n\right)}}$ be the distinct $\left(E_{n}+1\right)$-cubes in $[1, n]$ and let $F_{i}$ be the event $\mathbf{H}_{\mathbf{C}_{\mathbf{i}}^{\left(E_{n}+1, n\right)}} \subset S\left(n, \frac{1}{2}\right)$. By (3) we have

$$
\begin{gathered}
\operatorname{Pr}\left\{S_{n} \text { contains an }\left(E_{n}+1\right) \text {-cube }\right\}=\operatorname{Pr}\left\{\cup_{i=1}^{K} F_{i}\right\} \leq \sum_{i=1}^{K} \operatorname{Pr}\left\{F_{i}\right\} \leq \\
\binom{n}{E_{n}+2} 2^{-2^{E_{n}+1}} \leq \frac{n^{\log _{2} \log _{2} n+(1+o(1)) \log _{2} \log _{2} \log _{2} n}}{n^{\log _{2} \log _{2} n+(1+\epsilon+o(1)) \log _{2} \log _{2} \log _{2} n}}=o(1)
\end{gathered}
$$

which completes the proof.

## 3. Concluding remarks

The aim of this paper is to study non-degenerate Hilbert cubes in a random sequence. A natural problem would be to give analogous theorems for Hilbert cubes, where degenerate cubes are allowed. In this situation the dominant terms may come from arithmetic progressions. An $A P_{k+1}$ forms a $k$-cube. One can prove by the Janson inequality (see Lemma 2.5) that for a fixed $k \geq 2$

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, c n^{-\frac{2}{k+1}}\right) \text { contains no } A P_{k+1}\right\}=e^{-\frac{c^{k+1}}{2 k}}
$$

An easy argument shows (using Janson's inequality again) that for all $c>0$, with $p_{n}=c n^{-2 / 5}$

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, p_{n}\right) \text { contains no } 4 \text {-cubes }\right\}=e^{-\frac{c^{5}}{8}}
$$

Conjecture 3.1. For $k \geq 4$

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{S\left(n, c n^{-\frac{2}{k+1}}\right) \text { contains no } k \text {-cubes }\right\}=e^{-\frac{c^{k+1}}{2 k}}
$$

A simple calculation implies that in the random sequence $S\left(n, \frac{1}{2}\right)$ the length of the longest arithmetic progression is a.e. nearly $2 \log _{2} n$, therefore it contains a Hilbert cube of dimension $(2-\epsilon) \log _{2} n$.

Conjecture 3.2. For every $\epsilon>0$

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}\left\{\begin{array}{l}
\text { the maximal dimension of Hilbert cubes } \\
\text { in } S\left(n, \frac{1}{2}\right) \text { is }<(2+\epsilon) \log _{2} n
\end{array}\right\}=1
$$

N. Hegyvári (see [5]) studied the special case where the generating elements of Hilbert cubes are distinct. He proved that in this situation the maximal dimension of Hilbert cubes is a.e. between $c_{1} \log n$ and $c_{2} \log n \log \log n$. In this problem the lower bound seems to be the correct magnitude.

## References

[1] N. Alon, J. Spencer, The Probabilistic Method. Wiley-Interscience, Series in Discrete Math. and Optimization, 1992.
[2] A. Godbole, S. Janson, N. Locantore, R. Rapoport, Random Sidon Seqence. J. Number Theory 75 (1999), no. 1, 7-22.
[3] D. S. Gunderson, V. Rödl, Extremal problems for Affine Cubes of Integers. Combin. Probab. Comput 7 (1998), no. 1, 65-79.
[4] R. L. Graham, B. L. Rothchild, J. Spencer, Ramsey Theory. Wiley-Interscience, Series in Discrete Math. and Optimization, 1990.
[5] N. Hegyvári, On the dimension of the Hilbert cubes. J. Number Theory 77 (1999), no. 2, 326-330.

## Csaba SÁndor

Institute of Mathematics
Budapest University of Technology and Economics
Egry J. u. 1., H-1111 Budapest, Hungary
E-mail : csandor@math.bme.hu


[^0]:    Manuscrit reçu le 30 novembre 2005.
    Supported by Hungarian National Foundation for Scientific Research, Grant No. T 49693 and 61908.

