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# On the Increments of the Principal Value of Brownian Local Time

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**Abstract:** Let W be a one-dimensional Brownian motion starting from 0. Define  $Y(t) = \int_0^t \frac{ds}{W(s)} := \lim_{\epsilon \to 0} \int_0^t \mathbb{1}_{\{|W(s)| > \epsilon\}} \frac{ds}{W(s)}$  as Cauchy's principal value related to local time. We prove limsup and limit results for the increments of Y.

Keywords: Brownian motion, local time, principal value, large increments.

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### 1. Introduction

Let  $\{W(t); t \ge 0\}$  be a one-dimensional standard Brownian motion with W(0) = 0, and let  $\{L(t,x); t \ge 0, x \in \mathbb{R}\}$  denote its jointly continuous local time process. That is, for any Borel function  $f \ge 0$ ,

$$\int_0^t f(W(s)) \,\mathrm{d}s = \int_{-\infty}^\infty f(x) L(t, x) \,\mathrm{d}x, \qquad t \ge 0.$$

We are interested in the process

(1.1) 
$$Y(t) := \int_0^t \frac{\mathrm{d}s}{W(s)}, \qquad t \ge 0$$

Rigorously speaking, the integral  $\int_0^t ds / W(s)$  should be considered in the sense of Cauchy's principal value, i.e., Y(t) is defined by

(1.2) 
$$Y(t) := \lim_{\varepsilon \to 0^+} \int_0^t \frac{\mathrm{d}s}{W(s)} \mathbf{1}_{\{|W(s)| \ge \varepsilon\}} = \int_0^\infty \frac{L(t,x) - L(t,-x)}{x} \,\mathrm{d}x.$$

Since  $x \mapsto L(t,x)$  is Hölder continuous of order  $\nu$ , for any  $\nu < 1/2$ , the integral on the extreme right in (1.2) is almost surely absolutely convergent for all t > 0. The process  $\{Y(t), t \ge 0\}$  is called the principal value of Brownian local time.

It is easily seen that  $Y(\cdot)$  inherits a scaling property from Brownian motion, namely, for any fixed a > 0,  $t \mapsto a^{-1/2}Y(at)$  has the same law as  $t \mapsto Y(t)$ . Although some properties distinguish  $Y(\cdot)$  from Brownian motion (in particular,  $Y(\cdot)$  is not a semimartingale), it is a kind of folklore that the asymptotic behaviors of Y are somewhat like that of a Brownian motion. For detailed studies and surveys on principal value, and relation to Hilbert transform see Biane and Yor [4], Fitzsimmons and Getoor [13], Bertoin [2], [3], Yamada [20], Boufoussi et al. [5], Ait Ouahra and Eddahbi [1], Csáki et al. [11] and a collection of papers [22] together with their references. Biane and Yor [4] presented a detailed study on Y and determined a number of distributions for principal values and related processes.

Concerning almost sure limit theorems for Y and its increments, we summarize the relevant results in the literature. It was shown in [17] that the following law of the iterated logarithm holds:

**Theorem A.** (Hu and Shi [17])

(1.3) 
$$\limsup_{T \to \infty} \frac{Y(T)}{\sqrt{T \log \log T}} = \sqrt{8}, \qquad \text{a.s.}$$

This was extended in [10] to a Strassen-type [18] functional law of the iterated logarithm.

Theorem B. (Csáki et al. [10]) With probability one the set

(1.4) 
$$\left\{\frac{Y(xT)}{\sqrt{8T\log\log T}}, \ 0 \le x \le 1\right\}_{T \ge 3}$$

is relatively compact in C[0,1] with limit set equal to

(1.5) 
$$\mathcal{S} := \left\{ f \in C[0,1] : f(0) = 0, f \text{ is absolutely continuous and } \int_0^1 (f'(x))^2 \, \mathrm{d}x \le 1 \right\}.$$

Concerning Chung-type law of the iterated logarithm, we have the following result:

**Theorem C.** (Hu [16])

(1.6) 
$$\liminf_{T \to \infty} \sqrt{\frac{\log \log T}{T}} \sup_{0 \le s \le T} |Y(s)| = K_1, \quad \text{a.s}$$

with some (unknown) constant  $K_1 > 0$ .

The large increments were studied in [7] and [8]: **Theorem D.** (Csáki et al. [7]) Under the conditions

(1.7) 
$$\begin{cases} 0 < a_T \le T, \\ T \mapsto a_T \text{ and } T \mapsto T/a_T \text{ are both non-decreasing,} \\ \lim_{T \to \infty} \frac{\log(T/a_T)}{\log \log T} = \infty, \end{cases}$$

we have

(1.8) 
$$\lim_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{a_T \log(T/a_T)}} = 2, \quad \text{a.s.}$$

Wen [19] studied the lag increments of Y and among others proved the following results. **Theorem E.** (Wen [19])

(1.9) 
$$\limsup_{T \to \infty} \sup_{0 \le t \le T} \frac{\sup_{t \le s \le T} |Y(s) - Y(s - t)|}{\sqrt{t(\log(T/t) + 2\log\log t)}} = 2, \quad \text{a.s.}$$

Under the conditions  $0 < a_T \leq T$ ,  $a_T \to \infty$  as  $T \to \infty$ , we have

(1.10) 
$$\limsup_{T \to \infty} \sup_{0 \le t \le T - a_T} \frac{\sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{a_T(\log((t+a_T)/a_T) + 2\log\log a_T))}} \le 2, \quad \text{a.s.}$$

If  $a_T$  is onto, then we have equality in (1.10).

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In this note our aim is to investigate further limsup and liminf behaviors of the increments of Y.

**Theorem 1.1.** Assume that  $T \mapsto a_T$  is a function such that  $0 < a_T \leq T$ , and both  $a_T$  and  $T/a_T$  are non-decreasing. Then

(i)

(1.11) 
$$\limsup_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{a_T \left(\log \sqrt{T/a_T} + \log \log T\right)}} = \sqrt{8}, \quad \text{a.s}$$

(iia) If  $a_T > T(\log T)^{-\alpha}$  for some  $\alpha < 2$ , then

(1.12) 
$$\liminf_{T \to \infty} \sqrt{\frac{\log \log T}{a_T}} \sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| = K_2, \quad \text{a.s}$$

(iib) If  $a_T \leq T(\log T)^{-\alpha}$  for some  $\alpha > 2$ , then

(1.13) 
$$\liminf_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{a_T \log(T/a_T)}} = K_3, \quad \text{a.s.}$$

with some positive constants  $K_2, K_3$ . If, moreover,

$$\lim_{T \to \infty} \frac{\log(T/a_T)}{\log \log T} = \infty,$$

then  $K_3 = 2$ .

**Theorem 1.2.** Assume that  $T \mapsto a_T$  is a function such that  $0 < a_T \leq T$ , and both  $a_T$  and  $T/a_T$  are non-decreasing. Then

(1.14) 
$$\liminf_{T \to \infty} \frac{\sqrt{T \log \log T}}{a_T} \inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| = K_4, \quad \text{a.s.}$$

with some positive constant  $K_4$ . If  $\lim_{T\to\infty} (a_T/T) = 0$ , then  $K_4 = 1/\sqrt{2}$ .

(iia) If 
$$0 < \lim_{T \to \infty} (a_T/T) = \rho \le 1$$
, then  
(1.15) 
$$\limsup_{T \to \infty} \frac{\inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{T \log \log T}} = \rho \sqrt{8},$$

(iib) If

$$\lim_{T \to \infty} \frac{a_T (\log \log T)^2}{T} = 0,$$

a.s.

then

(1.16) 
$$\limsup_{T \to \infty} \frac{\sqrt{T}}{a_T \sqrt{\log \log T}} \inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| = K_5, \quad \text{a.s.}$$

with some positive constant  $K_5$ .

**Remark 1.** The exact values of the constants  $K_i$ , i = 2, 3, 4, 5 are unknown in general and it seems difficult to determine them except in certain particular cases. In the proofs we establish different upper and lower bounds. It follows however by 0-1 law for Brownian motion that the limsup's and liminf's considered here are non-random constants.

**Remark 2.** Plainly we recover some previous results on the path properties of Y by considering particular cases of Theorems 1.1 and 1.2. For instance, Theorems A and C follow from (1.11) and (1.12) respectively by taking  $a_T = T$ , and (1.8) follows from (1.11) combining with (1.13). However in Theorem 1.1(ii) and Theorem 1.2(ii) there are still small gaps in  $a_T$ .

The organization of the paper is as follows: In Section 2 some facts are presented needed in the proofs. Section 3 contains the necessary probability estimates. Theorem 1.1(i) and Theorem 1.1(iia,b) are proved in Sections 4 and 5, resp., while Theorem 1.2(i) and Theorem 1.2(iia,b) are proved in Sections 6 and 7, resp.

Throughout the paper, the letter K with subscripts will denote some important but unknown finite positive constants, while the letter c with subscripts denotes some finite and positive universal constants not important in our investigations. When the constants depend on a parameter, say  $\delta$ , they are denoted by  $c(\delta)$  with subscripts.

#### 2. Facts

Let  $\{W(t), t \ge 0\}$  be a standard Brownian motion and define the following objects:

(2.1) 
$$g := \sup\{t : t \le 1, W(t) = 0\}$$

(2.2) 
$$B(s) := \frac{W(sg)}{\sqrt{g}}, \qquad 0 \le s \le 1,$$

(2.3) 
$$m(s) := \frac{|W(g + s(1 - g))|}{\sqrt{1 - g}}, \qquad 0 \le s \le 1.$$

Here we summarize some well-known facts needed in our proofs.

Fact 2.1. (Biane and Yor [4])

(2.4) 
$$\frac{\mathbb{P}(Y(1) \in \mathrm{d}x)}{\mathrm{d}x} = \sqrt{\frac{2}{\pi^3}} \sum_{k=0}^{\infty} (-1)^k \exp\left(-\frac{(2k+1)^2 x^2}{8}\right), \quad x \in \mathbb{R}.$$

Consequently we have the estimate: for  $\delta > 0$ 

(2.5) 
$$c_1 \exp\left(-\frac{z^2}{8(1-\delta)}\right) \le \mathbb{P}(Y(1) \ge z) \le \exp\left(-\frac{z^2}{8}\right), \qquad z \ge 1$$

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with some positive constant  $c_1 = c_1(\delta)$ . Moreover, g,  $\{B(s), 0 \le s \le 1\}$  and  $\{m(s), 0 \le s \le 1\}$  are independent, g has arcsine distribution, B is a Brownian bridge and m is a Brownian meander.

(2.6)  
$$\mathbb{P}\left(\int_{0}^{1} \frac{\mathrm{d}v}{m(v)} < z \mid m(1) = 0\right)$$
$$= \sum_{k=-\infty}^{\infty} (1 - k^{2}z^{2}) \exp\left(-\frac{k^{2}z^{2}}{2}\right) = \frac{8\pi^{2}\sqrt{2\pi}}{z^{3}} \sum_{k=1}^{\infty} \exp\left(-\frac{2k^{2}\pi^{2}}{z^{2}}\right), \quad z > 0.$$

(2.7) 
$$\mathbb{P}(m(1) > x) = e^{-x^2/2}, \qquad x > 0.$$

Fact 2.2. (Yor [21, Exercise 3.4 and pp. 44]) Let  $Q_{x\to 0}^{\delta}$  be the law of the square of a Bessel bridge from x to 0 of dimension  $\delta > 0$  during time interval [0,1]. The process  $(m^2(1-v), 0 \le v \le 1)$ conditioned on  $\{m^2(1) = x\}$  is distributed as  $Q_{x\to 0}^3$ . Furthermore, we have

(2.8) 
$$Q_{x\to 0}^{\delta} = Q_{0\to 0}^{\delta} * Q_{x\to 0}^{0}, \quad \forall \, \delta > 0, \, x > 0,$$

where \* denotes convolution operator. Consequently, for any x > 0

(2.9) 
$$\mathbb{P}\left(\int_0^1 \frac{\mathrm{d}v}{m(v)} < z \,\Big|\, m(1) = x\right) \ge \mathbb{P}\left(\int_0^1 \frac{\mathrm{d}v}{m(v)} < z \,\Big|\, m(1) = 0\right).$$

Fact 2.3. (Hu [16]) For  $0 < z \le 1$ 

(2.10) 
$$c_2 \exp\left(-\frac{c_3}{z^2}\right) \le \mathbb{P}(\sup_{0 \le s \le 1} |Y(s)| < z) \le c_4 \exp\left(-\frac{c_5}{z^2}\right)$$

with some positive constants  $c_2, c_3, c_4, c_5$ .

**Fact 2.4.** (Csörgő and Révész [12]) Assume that  $T \mapsto a_T$  is a function such that  $0 < a_T \leq T$ , and both  $a_T$  and  $T/a_T$  are non-decreasing. Then

(2.11) 
$$\limsup_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |W(t+s) - W(t)|}{\sqrt{a_T (\log(T/a_T) + \log \log T)}} = \sqrt{2}, \quad \text{a.s}$$

Fact 2.5. (Strassen [18]) If  $f \in S$  defined by (1.5), then for any partition  $x_0 = 0 < x_1 < \ldots < x_k < x_{k+1} = 1$  we have

(2.12) 
$$\sum_{i=1}^{k+1} \frac{(f(x_i) - f(x_{i-1}))^2}{x_i - x_{i-1}} \le 1.$$

Fact 2.6. (Chung [6])

(2.13) 
$$\liminf_{t \to \infty} \sqrt{\frac{\log \log t}{t}} \sup_{0 \le s \le t} |W(s)| = \frac{\pi}{\sqrt{8}}, \quad \text{a.s}$$

Define  $g(T) := \max\{s \leq T : W(s) = 0\}$ . A joint lower class result for g(T) and  $M(T) := \sup_{0 \leq s \leq T} |W(s)|$  reads as follows.

Fact 2.7. (Grill [15]) Let  $\beta(t)$ ,  $\gamma(t)$  be positive functions slowly varying at infinity, such that  $0 < \beta(t) \le 1, 0 < \gamma(t) \le 1, \beta(t)$  is non-increasing,  $\beta(t)\sqrt{t} \uparrow \infty, \gamma(t)$  is monotone,  $\gamma(t)t \uparrow \infty$ ,  $\gamma(t)/\beta^2(t)$  is monotone. Then

$$\mathbb{P}\left(M(T) \le \beta(T)\sqrt{T}, g(T) \le \gamma(T)T \quad \text{i.o.}\right) = 0 \quad \text{or} \quad 1$$

according as  $I(\beta, \gamma) < \infty$  or  $= \infty$ , where

$$I(\beta,\gamma) = \int_{1}^{\infty} \frac{1}{t\beta^{2}(t)} \left(1 + \frac{\beta^{2}(t)}{\gamma(t)}\right)^{-1/2} \exp\left(-\frac{(4 - 3\gamma(t))\pi^{2}}{8\beta^{2}(t)}\right) \,\mathrm{d}t.$$

Now define  $d(T) := \min\{s \ge T : W(s) = 0\}$ . Since  $\{d(T) > t\} = \{g(t) < T\}$ , we deduce from Fact 2.7 the following estimate on d(T) when  $T \to \infty$ .

Fact 2.8. With probability 1

$$d(T) = O(T(\log T)^3), \qquad T \to \infty.$$

### 3. Probability estimates

**Lemma 3.1.** For  $T \ge 1$ ,  $\delta, z > 0$  we have

(3.1)  
$$\mathbb{P}\left(\sup_{0\leq t\leq T-1}\sup_{0\leq s\leq 1}|Y(t+s)-Y(t)|>z\right)$$
$$\leq c_6\left(\sqrt{T}\exp\left(-\frac{z^2}{8(1+\delta)}\right)+T\exp\left(-\frac{z^2}{2(1+\delta)}\right)\right)$$

with some positive constant  $c_6 = c_6(\delta)$ .

For the proof see Csáki et al. [7], Lemma 2.8.

**Lemma 3.2.** For T > 1,  $0 < \delta < 1/2$ , z > 1 we have

(3.2)  

$$\mathbb{P}\left(\sup_{0\leq t\leq T-1} (Y(t+1) - Y(t)) \geq z\right) \\
\geq \min\left(\frac{1}{2}, \frac{c_7\sqrt{T-1}}{z} \exp\left(-\frac{z^2}{8(1-\delta)}\right)\right) - \exp\left(-z^2\right)$$

with some positive constant  $c_7 = c_7(\delta) > 0$ .

**Proof.** Let us construct an increasing sequence of stopping times by  $\eta_0 := 0$  and

$$\eta_{k+1} := \inf\{t > \eta_k + 1 : W(t) = 0\}, \qquad k = 0, 1, 2, \dots$$

Let

$$\nu_t := \min\{i \ge 1 : \eta_i > t\}$$
  
$$Z_i := Y(\eta_{i-1} + 1) - Y(\eta_{i-1}), \qquad i = 1, 2, \dots$$

Then  $(Z_i, \eta_i - \eta_{i-1})_{i \ge 1}$  are i.i.d. random vectors with

$$\eta_i - \eta_{i-1} \stackrel{\text{law}}{=} 1 + \tau^2, \qquad Z_i \stackrel{\text{law}}{=} Y(1),$$

where  $\tau$  has Cauchy distribution. Clearly, for t > 0,

$$\sup_{0 \le s \le t} (Y(s+1) - Y(s)) \ge \max_{1 \le i \le \nu_t} Z_i = \overline{Z}_{\nu_t},$$

with  $\overline{Z}_k := \max_{1 \le i \le k} Z_i$ . First consider the Laplace transform  $(\lambda > 0)$ :

$$\begin{split} \lambda \int_{0}^{\infty} e^{-\lambda u} \mathbb{P}\left(\overline{Z}_{\nu_{u}} < z\right) \, \mathrm{d}u \\ &= \lambda \sum_{k=1}^{\infty} \mathbb{E} \int_{0}^{\infty} e^{-\lambda u} \mathbf{1}_{\{\eta_{k-1} \leq u < \eta_{k}\}} \mathbf{1}_{\{\overline{Z}_{k} < z\}} \, \mathrm{d}u \\ &= \sum_{k=1}^{\infty} \mathbb{E}\left(\left[e^{-\lambda \eta_{k-1}} - e^{-\lambda \eta_{k}}\right] \mathbf{1}_{\{\overline{Z}_{k} < z\}}\right) \\ &= \sum_{k=1}^{\infty} \left(\mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k} < z\}} e^{-\lambda \eta_{k-1}}\right] - \mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k} < z\}} e^{-\lambda \eta_{k}}\right]\right) \\ &= \sum_{k=1}^{\infty} \left(\mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k-1} < z\}} e^{-\lambda \eta_{k-1}}\right] - \mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k-1} < z, Z_{k} \geq z\}} e^{-\lambda \eta_{k-1}}\right] - \mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k-2}\}} e^{-\lambda \eta_{k}}\right]\right) \\ &= 1 - \sum_{k=1}^{\infty} \mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k-1} < z, Z_{k} \geq z\}} e^{-\lambda \eta_{k-1}}\right] \\ &= 1 - \sum_{k=1}^{\infty} \mathbb{E}\left[\mathbf{1}_{\{\overline{Z}_{k-1} < z\}} e^{-\lambda \eta_{k-1}}\right] \mathbb{P}(Y(1) \geq z) \\ &= 1 - \sum_{k=1}^{\infty} \left(\mathbb{E}\left[\mathbf{1}_{\{Z_{1} < z\}} e^{-\lambda \eta_{1}}\right]\right)^{k-1} \mathbb{P}(Y(1) \geq z) \\ &= 1 - \frac{\mathbb{P}(Y(1) \geq z)}{\mathbf{1} - \mathbb{E}\left[\mathbf{1}_{\{Z_{1} < z\}} e^{-\lambda \eta_{1}}\right]}, \end{split}$$

i.e.,

(3.3) 
$$\lambda \int_0^\infty e^{-\lambda u} \mathbb{P}\left(\overline{Z}_{\nu_u} \ge z\right) \, \mathrm{d}u = \frac{\mathbb{P}(Y(1) \ge z)}{1 - \mathbb{E}\left[1_{\{Z_1 < z\}} e^{-\lambda \eta_1}\right]}$$

But (recalling that  $Z_1 = Y(1)$ )

$$1 - \mathbb{E}\Big[\mathbf{1}_{\{Z_1 < z\}} e^{-\lambda\eta_1}\Big] \le 1 - \mathbb{E}(e^{-\lambda\eta_1}) + \mathbb{P}(Y(1) \ge z)$$

and (cf. [14], 3.466/1)

$$1 - \mathbb{E}e^{-\lambda\eta_1} = 1 - \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{e^{-\lambda(1+x^2)}}{1+x^2} \, \mathrm{d}x = \frac{2}{\sqrt{\pi}} \int_{0}^{\sqrt{\lambda}} e^{-x^2} \, \mathrm{d}x \le 2\sqrt{\lambda},$$

hence

$$\lambda \int_0^\infty e^{-\lambda u} \mathbb{P}\left(\overline{Z}_{\nu_u} \ge z\right) \, \mathrm{d}u \ge \frac{\mathbb{P}(Y(1) \ge z)}{2\sqrt{\lambda} + \mathbb{P}(Y(1) \ge z)}.$$

On the other hand, for any  $u_0 > 0$  we have

$$\begin{split} \lambda \int_0^\infty e^{-\lambda u} \mathbb{P}\left(\overline{Z}_{\nu_u} \ge z\right) \ \mathrm{d}u &= \lambda \int_0^{u_0} e^{-\lambda u} \mathbb{P}\left(\overline{Z}_{\nu_u} \ge z\right) \ \mathrm{d}u + \lambda \int_{u_0}^\infty e^{-\lambda u} \mathbb{P}\left(\overline{Z}_{\nu_u} \ge z\right) \ \mathrm{d}u \\ &\leq \mathbb{P}\left(\overline{Z}_{\nu_{u_0}} \ge z\right) + e^{-\lambda u_0}. \end{split}$$

It turns out that

(3.4) 
$$\mathbb{P}\left(\overline{Z}_{\nu_{u_0}} \ge z\right) \ge \frac{\mathbb{P}(Y(1) \ge z)}{2\sqrt{\lambda} + \mathbb{P}(Y(1) \ge z)} - e^{-\lambda u_0} \ge \min\left(\frac{1}{2}, \frac{\mathbb{P}(Y(1) \ge z)}{4\sqrt{\lambda}}\right) - e^{-\lambda u_0},$$

where the inequality

$$\frac{x}{y+x} \ge \min\left(\frac{1}{2}, \frac{x}{2y}\right), \qquad x > 0, \ y > 0$$

was used. Choosing  $u_0 = T - 1$ ,  $\lambda = z^2/u_0$ , and applying (2.5) of Fact 2.1, we finally get

(3.5)  

$$\mathbb{P}\left(\sup_{0\leq t\leq T-1} (Y(t+1) - Y(t)) \geq z\right) \\
\geq \min\left(\frac{1}{2}, \frac{c_8(\delta)\sqrt{T-1}}{z} \exp\left(-\frac{z^2}{8(1-\delta)}\right)\right) - \exp\left(-z^2\right).$$

This proves Lemma 3.2.

**Lemma 3.3.** For  $T \ge 2, 0 \le \kappa < 1$  and  $\delta, z > 0$  we have

(3.6) 
$$\mathbb{P}\left(\sup_{0 \le t \le T-1} (Y(t+1) - Y(t)) < z\right) \le \frac{5}{T^{\kappa/2}} + \exp\left(-c_9 T^{(1-\kappa)/2} e^{-(1+\delta)z^2/8}\right)$$

with some positive constant  $c_9 = c_9(\delta)$ .

See Csáki et al. [7], Lemma 3.1.

**Lemma 3.4.** For  $T \ge 1, 0 < z \le 1/2$  we have

(3.7) 
$$\mathbb{P}\left(\sup_{0 \le t \le T-1} \sup_{0 \le s \le 1} |Y(t+s) - Y(t)| < z\right) \ge \frac{c_{10}}{\sqrt{T}} \exp\left(-\frac{c_{11}}{z^2}\right)$$

with some positive constants  $c_{10}, c_{11}$ .

**Proof.** Define the events

$$A := \left\{ \sup_{0 \le s \le 1} |Y(s)| < \frac{z}{4}, W(1) \ge \frac{4}{z}, \inf_{1 \le u \le T} W(u) \ge \frac{2}{z} \right\}$$

and

$$\widetilde{A} := \left\{ \sup_{0 \le t \le T-1} \sup_{0 \le s \le 1} |Y(t+s) - Y(t)| < z \right\}.$$

Then  $A \subset \widetilde{A}$ , since if A occurs and  $t < 1, t + s \le 1$ , then

$$|Y(t+s) - Y(t)| \le 2 \sup_{0 \le s \le 1} |Y(s)| \le \frac{z}{2} < z.$$

If A occurs and  $t < 1, s \le 1, 1 < t + s \le T$ , then

$$|Y(t+s) - Y(t)| \le Y(t+s) - Y(1) + |Y(t) - Y(1)| \le \int_1^{t+s} \frac{\mathrm{d}u}{W(u)} + \frac{z}{2} < z.$$

Moreover, if A occurs and  $1 \le t, s \le 1, t + s \le T$ , then

$$|Y(t+s) - Y(t)| = \int_{t}^{t+s} \frac{\mathrm{d}u}{W(u)} \le \frac{z}{2} < z$$

Hence  $A \subset \widetilde{A}$  as claimed. But by the Markov property of W,

$$(3.8) \quad \mathbb{P}(A) = \int_{4/z}^{\infty} \mathbb{P}\left(\sup_{0 \le s \le 1} |Y(s)| < \frac{z}{4} \mid W(1) = x\right) \mathbb{P}\left(\inf_{1 \le u \le T} W(u) \ge \frac{2}{z} \mid W(1) = x\right) \varphi(x) \, \mathrm{d}x,$$

where  $\varphi$  denotes the standard normal density function.

Using reflection principle and  $x \ge 4/z$ ,  $z \le 1/2$ , we get

(3.9)  
$$\mathbb{P}\left(\inf_{1\leq u\leq T}W(u)\geq \frac{2}{z} \mid W(1)=x\right) = 2\Phi\left(\frac{x-2/z}{\sqrt{T-1}}\right) - 1$$
$$\geq 2\Phi\left(\frac{2}{z\sqrt{T-1}}\right) - 1\geq 2\Phi\left(\frac{4}{\sqrt{T}}\right) - 1\geq \frac{c_{12}}{\sqrt{T}},$$

with some constant  $c_{12} > 0$ , where  $\Phi(\cdot)$  is the standard normal distribution function. Hence

(3.10) 
$$\mathbb{P}(\widetilde{A}) \ge \mathbb{P}(A) \ge \frac{c_{12}}{\sqrt{T}} \mathbb{P}\left(\sup_{0 \le s \le 1} |Y(s)| \le \frac{z}{4}, W(1) \ge \frac{4}{z}\right).$$

To get a lower bound of the probability on the right-hand side, define g,  $(m(v), 0 \le v \le 1)$ ,  $(B(u), 0 \le u \le 1)$  by (2.1), (2.2) and (2.3), respectively. Recall (see Fact 2.1) that these three objects are independent, g has arc sine distribution, m is a Brownian meander and B is a Brownian

bridge. Moreover, (g, m, B) are independent of sgn(W(1)) which is a Bernoulli variable. Observe that

$$\sup_{0 \le s \le g} |Y(s)| = \sqrt{g} \sup_{0 \le s \le 1} \left| \int_0^s \frac{\mathrm{d}u}{B(u)} \right|,$$
$$\sup_{g \le s \le 1} |Y(s)| = |Y(1) - Y(g)| = \sqrt{1 - g} \int_0^1 \frac{\mathrm{d}v}{m(v)},$$
$$|W(1)| = \sqrt{1 - g} m(1).$$

Then

$$\begin{split} & \mathbb{P}\left(\sup_{0 \le s \le 1} |Y(s)| \le \frac{z}{4}, W(1) \ge \frac{4}{z}\right) \\ & \ge \mathbb{P}\left(\sup_{0 \le s \le g} |Y(s)| \le \frac{z}{8}, Y(1) - Y(g) \le \frac{z}{8}, W(1) \ge \frac{4}{z}\right) \\ & \ge \mathbb{P}\left(\sqrt{g} \sup_{0 \le s \le 1} \left|\int_0^s \frac{\mathrm{d}u}{B(u)}\right| \le \frac{z}{8}, \sqrt{1-g} \int_0^1 \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8}, \sqrt{1-g} m(1) \ge \frac{4}{z}, W(1) > 0, g < z^2\right) \\ & \ge \mathbb{P}\left(\sup_{0 \le s \le 1} \left|\int_0^s \frac{\mathrm{d}u}{B(u)}\right| \le \frac{1}{8}, \int_0^1 \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8}, m(1) \ge \frac{4}{z\sqrt{1-z^2}}, W(1) > 0, g < z^2\right) \\ & = \mathbb{P}\left(\sup_{0 \le s \le 1} \left|\int_0^s \frac{\mathrm{d}u}{B(u)}\right| \le \frac{1}{8}\right) \mathbb{P}\left(\int_0^1 \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8}, m(1) \ge \frac{4}{z\sqrt{1-z^2}}\right) \mathbb{P}(W(1) > 0) \mathbb{P}(g < z^2) \\ & \ge c_{13}z \mathbb{P}\left(\int_0^1 \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8}, m(1) \ge \frac{4}{z\sqrt{1-z^2}}\right) \\ & = c_{13}z \int_{4/(z\sqrt{1-z^2})}^{\infty} \mathbb{P}\left(\int_0^1 \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8} \mid m(1) = x\right) \mathbb{P}(m(1) \in \mathrm{d}x). \end{split}$$

It follows from Facts 2.1 and 2.2 that for x > 0, z > 0

(3.11) 
$$\mathbb{P}\left(\int_{0}^{1} \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8} \, \middle| \, m(1) = x\right) \ge \mathbb{P}\left(\int_{0}^{1} \frac{\mathrm{d}v}{m(v)} \le \frac{z}{8} \, \middle| \, m(1) = 0\right) \ge \frac{c_{14}}{z^{3}} \exp\left(-\frac{c_{15}}{z^{2}}\right)$$

and

(3.12) 
$$\mathbb{P}\left(m(1) > \frac{4}{z\sqrt{1-z^2}}\right) = \exp\left(-\frac{8}{z^2(1-z^2)}\right).$$

Putting (3.10), (3.11), (3.12) together, we get (3.7).

**Lemma 3.5.** For  $T > 1, 0 < z \le 1/2, 0 < \delta \le 1/2$  we have

(3.13) 
$$\mathbb{P}\left(\inf_{0 \le t \le T-1} \sup_{0 \le s \le 1} |Y(t+s) - Y(t)| < z\right) \\
\le c_{16}\left(\exp\left(-\frac{(1-\delta)^2}{2(1+\delta)^2 z^2 T}\right) + \exp\left(-\frac{c_5\delta}{4(1+\delta)^2 z^2}\right) + \exp\left(\frac{c_{17}}{z^2} - \frac{c_{18}z^2}{T}e^{c_{19}/z^2}\right)\right)$$

with some positive constants  $c_{16}$ ,  $c_{17} = c_{17}(\delta)$ ,  $c_{18} = c_{18}(\delta)$ ,  $c_{19} = c_{19}(\delta)$ .

**Proof.** Consider a positive integer N to be given later, h = (T-1)/N,  $t_k = kh$ , k = 0, 1, 2, ..., N. Then for  $0 < \delta \le 1/2$  we have

$$\mathbb{P}\left(\inf_{0 \le t \le T-1} \sup_{0 \le s \le 1} |Y(t+s) - Y(t)| < z \right)$$

$$\le \mathbb{P}\left(\inf_{0 \le k \le N} \sup_{0 \le s \le 1} |Y(t_k+s) - Y(t_k)| \le (1+\delta)z \right) + \mathbb{P}\left(\sup_{0 \le t \le T-1} \sup_{0 \le s \le h} |Y(t+s) - Y(t)| > \delta z \right)$$

$$=: P_1 + P_2.$$

By scaling and Lemma 3.1

$$P_{2} = \mathbb{P}\left(\sup_{0 \le t \le (T-1)/h} \sup_{0 \le s \le 1} |Y(t+s) - Y(t)| > \frac{\delta z}{\sqrt{h}}\right)$$
  
$$\leq c_{6}\left(\sqrt{\frac{T-1}{h} + 1} \exp\left(-\frac{\delta^{2} z^{2}}{8h(1+\delta)}\right) + \left(\frac{T-1}{h} + 1\right) \exp\left(-\frac{\delta^{2} z^{2}}{2h(1+\delta)}\right)\right)$$
  
$$\leq 2c_{6}(N+1) \exp\left(-\frac{\delta^{2} z^{2}}{8h(1+\delta)}\right).$$

To bound  $P_1$ , we denote by  $d(t) := \inf\{s \ge t : W(s) = 0\}$  the first zero of W after t. Consider those k for which  $\sup_{0 \le s \le 1} |Y(t_k + s) - Y(t_k)| \le (1 + \delta)z$ . If, moreover,  $d(t_k) \ge t_k + 1 - \delta$ , which means that the Brownian motion W does not change sign over  $[t_k, t_k + 1 - \delta)$ , then

$$(1+\delta)z \ge |Y(t_k+1-\delta) - Y(t_k)| = \int_0^{1-\delta} \frac{\mathrm{d}s}{|W(t_k+s)|} \ge \frac{1-\delta}{\sup_{0 \le s \le T} |W(s)|}$$

and it follows that

$$P_1 \leq \mathbb{P}\left(\sup_{0\leq s\leq T} |W(s)| > \frac{(1-\delta)}{z(1+\delta)}\right)$$
  
+  $\mathbb{P}\left(\exists k \leq N : \sup_{0\leq s\leq 1} |Y(t_k+s) - Y(t_k)| \leq (1+\delta)z; d(t_k) < t_k + 1 - \delta\right)$   
$$\leq 4 \exp\left(-\frac{(1-\delta)^2}{2(1+\delta)^2 z^2 T}\right)$$
  
+  $\sum_{k=0}^N \mathbb{P}\left(\sup_{0\leq s\leq 1} |Y(t_k+s) - Y(t_k)| \leq (1+\delta)z; d(t_k) < t_k + 1 - \delta\right).$ 

Let  $\widehat{W}(s) = W(d(t_k) + s)$  for  $s \ge 0$  and  $\widehat{Y}(s)$  be the associated principal values. Observe that on  $\{\sup_{0\le s\le 1} |Y(t_k + s) - Y(t_k)| \le (1+\delta)z; d(t_k) < t_k + 1 - \delta\}$ , we have  $\sup_{0\le u\le \delta} |\widehat{Y}(u) + (Y(d(t_k)) - Y(t_k))| < (1+\delta)z$ , and  $|Y(d(t_k)) - Y(t_k)| \le (1+\delta)z$  which imply that

$$\sup_{0 \le u \le \delta} |\widehat{Y}(u)| < 2(1+\delta)z.$$

By scaling and Fact 2.3 we have

$$\mathbb{P}\left(\sup_{0\leq u\leq \delta}|\widehat{Y}(u)|<2(1+\delta)z\right)\leq c_4\exp\left(-\frac{c_5\delta}{4(1+\delta)^2z^2}\right).$$

Therefore, we obtain:

$$P_1 \le 4 \exp\left(-\frac{(1-\delta)^2}{2(1+\delta)^2 z^2 T}\right) + c_4(N+1) \exp\left(-\frac{c_5\delta}{4(1+\delta)^2 z^2}\right)$$

Hence

$$\begin{split} P_1 + P_2 &\leq 4 \exp\left(-\frac{(1-\delta)^2}{2(1+\delta)^2 z^2 T}\right) + c_4(N+1) \exp\left(-\frac{c_5 \delta}{4(1+\delta)^2 z^2}\right) \\ &+ 2c_6(N+1) \exp\left(-\frac{\delta^2 z^2}{8h(1+\delta)}\right). \end{split}$$

By taking  $N = [e^{c_5\delta/(4(1+\delta)^2z^2)}] + 1$ , we get

$$P_1 + P_2 \le c_{16} \left( \exp\left(-\frac{(1-\delta)^2}{2(1+\delta)^2 z^2 T}\right) + \exp\left(-\frac{c_5\delta}{4(1+\delta)^2 z^2}\right) + \exp\left(\frac{c_{17}}{z^2} - \frac{c_{18}z^2}{T}e^{c_{19}/z^2}\right) \right)$$

with relevant constants  $c_{16}$ ,  $c_{17}$ ,  $c_{18}$ ,  $c_{19}$ , proving (3.13).

The upper estimation, i.e.

(4.1) 
$$\limsup_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{8a_T \left(\log \sqrt{T/a_T} + \log \log T\right)}} \le 1, \quad \text{a.s.}$$

follows easily from Wen's Theorem E.

Now we prove the lower bound, i.e.

(4.2) 
$$\limsup_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{8a_T \left(\log \sqrt{T/a_T} + \log \log T\right)}} \ge 1, \quad \text{a.s.}$$

In the case when  $a_T = T$ , (4.2) follows from the law of the iterated logarithm (1.3) of Theorem A. Now we assume that  $a_T/T \le \rho < 1$ , with some constant  $\rho$  for all T > 0.

By scaling, (3.2) of Lemma 3.2 is equivalent to

(4.3)  
$$\mathbb{P}\left(\sup_{0\leq t\leq T-a} (Y(t+a) - Y(t)) \geq z\sqrt{a}\right)$$
$$\geq \min\left(\frac{1}{2}, \frac{c_7\sqrt{T/a-1}}{z} \exp\left(-\frac{z^2}{8(1-\delta)}\right)\right) - \exp\left(-z^2\right)$$

for 0 < a < T,  $0 < \delta < 1/2$ , z > 1.

Define the sequences

(4.4) 
$$t_k := e^{7k \log k}, \qquad k = 1, 2, \dots$$

and  $\theta_0 := 0$ ,

(4.5) 
$$\theta_k := \inf\{t > T_k : W(t) = 0\}, \qquad k = 1, 2, \dots,$$

where  $T_k := \theta_{k-1} + t_k$ . For  $0 < \delta < \min(1/2, 1 - \rho)$  define the events

$$A_k := \left\{ \sup_{0 \le t \le t_k (1-\delta) - a_{t_k}} (Y(\theta_{k-1} + t + a_{t_k}) - Y(\theta_{k-1} + t)) \ge (1-\delta)\beta_k \right\}, \quad k = 1, 2, \dots$$

with

$$\beta_k := \sqrt{8a_{t_k} \left(\log \sqrt{\frac{t_k}{a_{t_k}}} + \log \log t_k\right)}.$$

Applying (4.3) with  $T = t_k(1-\delta)$ ,  $a = a_{t_k}$ ,  $z = (1-\delta)\sqrt{8(\log\sqrt{t_k/a_{t_k}} + \log\log t_k)}$ , we have for k large

$$\mathbb{P}(A_k) = \mathbb{P}\left(\sup_{0 \le t \le t_k(1-\delta) - a_{t_k}} (Y(t+a_{t_k}) - Y(t)) \ge (1-\delta)\beta_k\right)$$
$$\ge \min\left(\frac{1}{2}, \frac{b_k}{(\log t_k)^{1-\delta}}\right) - \frac{1}{(\log t_k)^{8(1-\delta)^2}}$$

with

$$b_k = \frac{c_7 \sqrt{t_k (1-\delta)/a_{t_k} - 1}}{(t_k/a_{t_k})^{(1-\delta)/2} \sqrt{\log \sqrt{t_k/a_{t_k}} + \log \log t_k}} \ge \frac{c_{20}}{\sqrt{\log k}}.$$

Hence  $\sum_k \mathbb{P}(A_k) = \infty$  and since  $A_k$  are independent, Borel-Cantelli lemma yields

$$\mathbb{P}(A_k \text{ i.o.}) = 1.$$

It follows that

(4.6) 
$$\limsup_{k \to \infty} \frac{\sup_{0 \le t \le t_k (1-\delta) - a_{t_k}} (Y(\theta_{k-1} + t + a_{t_k}) - Y(\theta_{k-1} + t))}{\sqrt{8a_{t_k} \left(\log \sqrt{\frac{t_k}{a_{t_k}}} + \log \log t_k\right)}} \ge 1 - \delta, \quad \text{a.s.}$$

It can be seen (cf. [9]) that we have almost surely for large enough k

$$t_k \le T_k \le t_k \left(1 + \frac{1}{k}\right),$$
  
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consequently

(4.7) 
$$\lim_{k \to \infty} \frac{t_k}{T_k} = 1, \qquad \text{a.s.}$$

Since by our assumptions

$$\frac{t_k}{T_k} \le \frac{a_{t_k}}{a_{T_k}} \le 1,$$

we have also

(4.8) 
$$\lim_{k \to \infty} \frac{a_{t_k}}{a_{T_k}} = 1, \qquad \text{a.s.}$$

On the other hand, for any  $\delta > 0$  small enough we have almost surely for large k

$$a_{T_k} \le (1+\delta)a_{t_k} \le t_k\delta + a_{t_k},$$

thus

$$T_k - a_{T_k} \ge T_k - t_k \delta - a_{t_k},$$

consequently

(4.9) 
$$\sup_{\substack{0 \le t \le T_k - a_{T_k} \ 0 \le s \le a_{T_k}}} \sup_{\substack{V(t+s) - Y(t) \\ \ge \sup_{0 \le t \le t_k (1-\delta) - a_{t_k}}} (Y(\theta_{k-1} + t + a_{t_k}) - Y(\theta_{k-1} + t)),$$

hence we have also

(4.10) 
$$\limsup_{k \to \infty} \frac{\sup_{0 \le t \le T_k - a_{T_k}} \sup_{0 \le s \le a_{T_k}} |Y(t+s) - Y(t)|}{\sqrt{8a_{t_k} \left(\log \sqrt{\frac{t_k}{a_{t_k}}} + \log \log t_k\right)}} \ge 1 - \delta, \quad \text{a.s.}$$

and since  $\delta > 0$  can be arbitrary small, (4.2) follows by combining (4.7), (4.8), (4.9) and (4.10).

# 5. Proof of Theorem 1.1(ii)

First assume that

(5.1) 
$$a_T > \frac{T}{(\log T)^{\alpha}}$$
 for some  $\alpha < 2$ .

By Theorem C,

(5.2)  
$$\lim_{T \to \infty} \sqrt{\frac{\log \log T}{a_T}} \sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|$$
$$\ge \liminf_{T \to \infty} \sqrt{\frac{\log \log a_T}{a_T}} \sup_{0 \le s \le a_T} |Y(s)| \ge K_1, \quad \text{a.s.},$$

proving the lower bound in (1.12).

To get an upper bound, note that by scaling, (3.7) of Lemma 3.4 is equivalent to

(5.3) 
$$\mathbb{P}\left(\sup_{0 \le t \le T-a} \sup_{0 \le s \le a} |Y(s+t) - Y(t)| < z\sqrt{a}\right) \ge c_{10}\sqrt{\frac{a}{T}} \exp\left(-\frac{c_{11}}{z^2}\right)$$

for  $T \ge a, \ 0 < z \le 1/2$ .

Let  $t_k$  and  $\theta_k$  be defined by (4.4) and (4.5), resp.,  $T_k = \theta_{k-1} + t_k$  as in the proof of Theorem 1.1(i). Let  $c_{11}$  be the constant as in (5.3) and choose  $\delta > 0$  such that  $\alpha/2 + c_{11}/\delta^2 < 1$ . For  $\varepsilon > 0$  define the events

$$E_k := \left\{ \sup_{0 \le t \le (1+\varepsilon)t_k - a_{t_k}(1+\varepsilon)} \sup_{0 \le s \le a_{t_k}(1+\varepsilon)} |Y(\theta_{k-1} + t + s) - Y(\theta_{k-1} + t)| \le \delta \sqrt{\frac{a_{t_k}}{\log\log t_k}} \right\}.$$

Then putting  $T = (1 + \varepsilon)t_k$ ,  $a = a_{(1+\varepsilon)t_k}$ ,  $z = \delta/\sqrt{\log \log t_k}$  into (5.3), we get

$$\mathbb{P}(E_k) = \mathbb{P}\left(\sup_{0 \le t \le (1+\varepsilon)t_k - a_{t_k}(1+\varepsilon)} \sup_{0 \le s \le a_{t_k}(1+\varepsilon)} |Y(t+s) - Y(t)| \le \delta \sqrt{\frac{a_{t_k}}{\log\log t_k}}\right)$$
  
$$\ge c_{10}\sqrt{\frac{a_{t_k}}{t_k}} \exp(-(c_{11}/\delta^2)\log\log(t_k) \ge \frac{c_{10}}{(\log t_k)^{\alpha/2 + c_{11}/\delta^2}} = \frac{c_{10}}{(7k\log k)^{\alpha/2 + c_{11}/\delta^2}},$$

hence  $\sum_k \mathbb{P}(E_k) = \infty$ , and since  $E_k$  are independent, we have  $\mathbb{P}(E_k \text{ i.o.}) = 1$ , i.e.

(5.4) 
$$\liminf_{k \to \infty} \sqrt{\frac{\log \log t_k}{a_{t_k}}} \sup_{0 \le t \le (1+\varepsilon)t_k - a_{t_k}(1+\varepsilon)} \sup_{0 \le s \le a_{t_k}(1+\varepsilon)} |Y(\theta_{k-1} + t + s) - Y(\theta_{k-1} + t)| \le \delta, \quad \text{a.s.}$$

for any  $\varepsilon > 0$ . For large enough k by (4.7) and (4.8) we have  $a_{T_k} \leq (1+\varepsilon)a_{t_k}$ , a.s. and  $T_k - a_{T_k} \leq \theta_{k-1} + (1+\varepsilon)t_k - (1+\varepsilon)a_{t_k}$ , a.s. Thus given any  $\varepsilon > 0$ , we have for large k

(5.5) 
$$\sup_{\substack{0 \le t \le T_k - a_{T_k} \ 0 \le s \le a_{T_k}}} \sup_{\substack{0 \le t \le \theta_{k-1} \ 0 \le s \le a_{T_k}}} |Y(t+s) - Y(t)| \\ \le 2 \sup_{\substack{0 \le t \le \theta_{k-1}}} |Y(t)| + \sup_{\substack{0 \le t \le (1+\varepsilon)t_k - a_{t_k}(1+\varepsilon)}} \sup_{\substack{0 \le s \le a_{t_k}(1+\varepsilon)}} |Y(\theta_{k-1} + t+s) - Y(\theta_{k-1} + t)|.$$

By Theorem A, Fact 2.8, (4.7), (5.1) and simple calculation,

(5.6)  
$$\sup_{0 \le t \le \theta_{k-1}} |Y(t)| = O(\theta_{k-1} \log \log \theta_{k-1})^{1/2}$$
$$= O(t_{k-1} (\log t_{k-1})^3 \log \log t_{k-1})^{1/2} = O\left(\frac{a_{t_k}}{\log \log t_k}\right)^{1/2}, \quad \text{a.s.}$$

as  $k \to \infty$ . Assembling (5.4), (5.5) and (5.6), we get

$$\liminf_{k \to \infty} \sqrt{\frac{\log \log t_k}{a_{t_k}}} \sup_{0 \le t \le T_k - a_{T_k}} \sup_{0 \le s \le a_{T_k}} |Y(t+s) - Y(t)|$$

$$= \liminf_{k \to \infty} \sqrt{\frac{\log \log T_k}{a_{T_k}}} \sup_{0 \le t \le T_k - a_{T_k}} \sup_{0 \le s \le a_{T_k}} |Y(t+s) - Y(t)| \le \delta, \qquad \text{a.s}$$

which together with (5.2) yields (1.12).

Now assume that

(5.7) 
$$a_T \le \frac{T}{(\log T)^{\alpha}}$$
 for some  $\alpha > 2$ .

By Theorem 1.1(i),

(5.8)  
$$\lim_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{a_T \log(T/a_T)}}$$
$$\leq \limsup_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{a_T \log(T/a_T)}}$$
$$\leq \limsup_{T \to \infty} \frac{\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)|}{\sqrt{\frac{2\alpha a_T}{\alpha + 2} \left(\log \sqrt{T/a_T} + \log \log T\right)}} \le 2\sqrt{\frac{\alpha + 2}{\alpha}},$$

i.e., an upper bound in (1.13) follows.

To get a lower bound under (5.7), observe that by scaling, (3.6) of Lemma 3.3 is equivalent to

$$\mathbb{P}\left(\sup_{0\leq t\leq T-a} (Y(t+a)-Y(t)) < z\sqrt{a}\right) \leq 5\left(\frac{a}{T}\right)^{\kappa/2} + \exp\left(-c_9\left(\frac{T}{a}\right)^{(1-\kappa)/2} e^{-(1+\delta)z^2/8}\right)$$

for  $a \leq T, 0 \leq \kappa < 1, 0 < \delta, 0 < z$ . Using (5.7) we get further

(5.9) 
$$\mathbb{P}\left(\sup_{0 \le t \le T-a} (Y(t+a) - Y(t)) < z\sqrt{a}\right)$$
$$\le \frac{5}{(\log T)^{\alpha \kappa/2}} + \exp\left(-c_9 (\log T)^{\alpha(1-\kappa)/2} e^{-(1+\delta)z^2/8}\right).$$

In the case when (1.7) holds, (1.13) was proved in [7]. In other cases the proof is similar. Let  $T_k = e^k$  and define the events

$$F_{k} = \left\{ \sup_{0 \le t \le T_{k} - a_{T_{k}}} (Y(t + a_{T_{k}}) - Y(t)) \le C_{1} \sqrt{a_{T_{k}} \log \frac{T_{k}}{a_{T_{k}}}} \right\}$$

with

$$C_1 = 2\sqrt{\frac{\alpha - 2 - 2\varepsilon\alpha}{(1+\delta)\alpha}}.$$

By (5.9) with  $\kappa = 2/\alpha + \varepsilon$ ,

$$\mathbb{P}(F_k) \le \frac{5}{k^{\alpha\kappa/2}} + \exp\left(-c_9 k^{\alpha((1-\kappa)/2 - (1+\delta)C_1^2/8)}\right) \le \frac{5}{k^{1+\alpha\varepsilon/2}} + \exp\left(-c_9 k^{\alpha\varepsilon/2}\right)$$

One can easily see that with these choices  $\sum_k \mathbb{P}(F_k) < \infty$ , consequently

$$\liminf_{k \to \infty} \frac{\sup_{0 \le t \le T_k - a_{T_k}} (Y(t + a_{T_k}) - Y(t))}{\sqrt{a_{T_k} \log \frac{T_k}{a_{T_k}}}} \ge C_1, \qquad \text{a.s.},$$

implying also

$$\liminf_{k \to \infty} \frac{\sup_{0 \le t \le T_k - a_{T_k}} \sup_{0 \le s \le a_{T_k}} |Y(t+s) - Y(t)|}{\sqrt{a_{T_k} \log \frac{T_k}{a_{T_k}}}} \ge 2\sqrt{\frac{\alpha - 2}{\alpha}}, \qquad \text{a.s.},$$

for  $\varepsilon$  can be choosen arbitrary small.

Since  $\sup_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t + s) - Y(t)|$  is increasing in T, we obtain a lower bound in (1.13). This together with the 0-1 law for Brownian motion complete the proof of Theorem 1.1(ii).

## 6. Proof of Theorem 1.2(i)

If  $a_T = T$ , then (1.14) is equivalent to Theorem C. Now assume that  $\rho := \lim_{T \to \infty} a_T/T < 1$ .

First we prove the lower bound, i.e.

(6.1) 
$$\liminf_{T \to \infty} \frac{\sqrt{T \log \log T}}{a_T} \inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| \ge c, \quad \text{a.s.}$$

By scaling, (3.13) of Lemma 3.5 is equivalent to

(6.2) 
$$\mathbb{P}\left(\inf_{0 \le t \le T-a} \sup_{0 \le s \le a} |Y(t+s) - Y(t)| < z\sqrt{a}\right) \\ \le c_{16} \left(\exp\left(-\frac{a(1-\delta)^2}{2(1+\delta)^2 z^2 T}\right) + \exp\left(-\frac{c_5\delta}{4(1+\delta)^2 z^2}\right) + \exp\left(\frac{c_{17}}{z^2} - \frac{c_{18}az^2}{T}e^{c_{19}/z^2}\right)\right)$$

for  $a < T, \, 0 < z \le 1/2, \, 0 < \delta \le 1/2.$ 

Define the events

$$G_k = \left\{ \inf_{0 \le t \le T_{k+1} - a_{T_k}} \sup_{0 \le s \le a_{T_k}} |Y(t+s) - Y(t)| < z_k \sqrt{a_{T_k}} \right\} \quad k = 1, 2, \dots$$

Let  $T_k = e^k$  and put  $T = T_{k+1}$ ,  $a = a_{T_k}$ ,

$$z = z_k = C_2 \sqrt{\frac{a_{T_k}}{T_{k+1} \log \log T_{k+1}}}$$

into (6.2). The constant  $C_2$  will be choosen later. Denoting the terms on the right-hand side of (6.2) by  $I_1$ ,  $I_2$ ,  $I_3$ , resp., we have

$$\mathbb{P}(G_k) \le c_{16}(I_1^{(k)} + I_2^{(k)} + I_3^{(k)}),$$
  
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where

$$I_1^{(k)} = \exp\left(-\frac{c_{21}}{C_2^2}\log\log T_{k+1}\right),$$

$$I_2^{(k)} = \exp\left(-\frac{c_{22}T_k}{C_2^2a_{T_k}}\log\log T_{k+1}\right),$$

$$I_3^{(k)} = \exp\left(\frac{c_{23}T_k\log\log T_{k+1}}{C_2^2a_{T_k}} - \frac{c_{24}C_2^2a_{T_k}^2}{T_k^2\log\log T_{k+1}}\left(\log T_{k+1}\right)^{\frac{c_{25}T_k}{C_2^2a_{T_k}}}\right)$$

with some constants  $c_{21} = c_{21}(\delta)$ ,  $c_{22} = c_{22}(\delta)$ ,  $c_{23}$ ,  $c_{24}$ ,  $c_{25}$ .

One can see easily that for any choice of positive  $C_2$  and for all possible  $a_T$  (satisfying our conditions) we have  $\sum_k I_3^{(k)} < \infty$ . So we show that for appropriate choice of  $C_2$  we have also  $\sum_k I_j^{(k)} < \infty, j = 1, 2$ .

First consider the case  $0 < \rho$ . Choosing a positive  $\delta$ , one can select  $C_2 < \min\left(\sqrt{c_{21}}, \sqrt{\frac{c_{22}}{\rho}}\right)$ and it is easy to verify that  $\sum_k I_j^{(k)} < \infty$ , j = 1, 2, hence also  $\sum_k \mathbb{P}(G_k) < \infty$ .

In the case  $\rho = 0$  choose  $C_2 < (1 - \delta)/((1 + \delta)\sqrt{2})$ . With this choice we have  $\sum_k I_1^{(k)} < \infty$ for arbitrary  $\delta > 0$ . Since  $\lim_{k\to\infty} (T_k/a_{T_k}) = \infty$ , we have also  $\sum_k I_2^{(k)} < \infty$  and  $\sum_k \mathbb{P}(G_k) < \infty$ . The Borel-Cantelli lemma and interpolation between  $T_k$ 's finish the proof of (6.1). We have also verified that in the case  $\rho = 0$  one can choose  $c = 1/\sqrt{2}$  in (6.1), since  $\delta > 0$  can be choosen arbitrary small.

Now we turn to the proof of the upper bound, i.e.

(6.3) 
$$\lim_{T \to \infty} \inf_{a_T} \frac{\sqrt{T \log \log T}}{a_T} \inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| \le C_3, \quad \text{a.s.}$$

with some constant  $C_3$ .

If  $\rho > 0$ , then

$$\inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| \le \sup_{0 \le s \le a_T} |Y(s)| \le \sup_{0 \le s \le T} |Y(s)|$$

and hence (6.3) with some positive constant  $C_3$  follows from Theorem C.

If  $\rho = 0$ , then let for any  $\varepsilon > 0$ 

(6.4) 
$$\lambda_T := \inf\{t : |W(t)| = \sup_{0 \le s \le T(1-\varepsilon)} |W(s)|\}.$$

According to the law of the iterated logarithm, with probability one there exists a sequence  $\{T_i, i \ge 1\}$  such that  $\lim_{i\to\infty} T_i = \infty$  and

(6.5) 
$$|W(\lambda_{T_i})| \ge \sqrt{2T_i(1-\varepsilon)\log\log T_i}.$$

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But Fact 2.4 implies that for  $\varepsilon>0$ 

(6.6) 
$$|W(\lambda_{T_i}) - W(s)| \le \sqrt{2(1+\varepsilon)\varepsilon T_i \log \log T_i}, \quad \lambda_{T_i} \le s \le \lambda_{T_i} + \varepsilon T_i, \quad i \ge 1.$$

Now assume that  $W(\lambda_{T_i}) > 0$ . The case when  $W(\lambda_{T_i}) < 0$  is similar. Then (6.5) and (6.6) imply

(6.7) 
$$W(s) \ge \left(\sqrt{1-\varepsilon} - \sqrt{\varepsilon(1+\varepsilon)}\right) \sqrt{2T_i \log \log T_i}, \quad \lambda_{T_i} \le s \le \lambda_{T_i} + \varepsilon T_i$$

 $\rho = 0$  implies that  $a_T \leq \varepsilon T$  for any  $\varepsilon > 0$  and large enough T, hence we have from (6.7) for large i

$$\sup_{0 \le s \le a_{T_i}} \left( Y(\lambda_{T_i} + s) - Y(\lambda_{T_i}) \right) = Y(\lambda_{T_i} + a_{T_i}) - Y(\lambda_{T_i}) = \int_{\lambda_{T_i}}^{\lambda_{T_i} + a_{T_i}} \frac{\mathrm{d}s}{W(s)}$$
$$\le \frac{a_{T_i}}{\left(\sqrt{1 - \varepsilon} - \sqrt{\varepsilon(1 + \varepsilon)}\right)\sqrt{2T_i \log \log T_i}}.$$

Since  $\varepsilon > 0$  is arbitrary, (6.3) follows with  $C_3 = 1/\sqrt{2}$ . This completes the proof of Theorem 1.2(i).

# 7. Proof of Theorem 1.2(ii)

If  $\rho = 1$ , then (1.15) is equivalent to (1.3) of Theorem A. So we may assume that  $0 < \rho < 1$ . It suffices to show (1.15) when  $a_T = \rho T$ .

First we prove the upper bound

(7.1) 
$$\limsup_{T \to \infty} \frac{\inf_{0 \le t \le T - \rho T} \sup_{0 \le s \le \rho T} |Y(t+s) - Y(t)|}{\sqrt{8T \log \log T}} \le \rho, \quad \text{a.s.}$$

Let k be the largest integer for which  $k\rho < 1$  and put  $x_i = i\rho$ , i = 0, 1, ..., k,  $x_{k+1} = 1$ . It suffices to show that if  $f \in S$  defined by (1.5), then

$$\min_{1 \le i \le k+1} |f(x_i) - f(x_{i-1})| \le \rho.$$

Assume on the contrary that

$$|f(x_i) - f(x_{i-1})| > \rho, \quad \forall i = 1, 2, \dots, k+1.$$

Then

$$\sum_{i=1}^{k+1} \frac{(f(x_i) - f(x_{i-1}))^2}{x_i - x_{i-1}} > \sum_{i=1}^k \frac{\rho^2}{\rho} + \frac{\rho^2}{1 - k\rho} = k\rho + \frac{\rho^2}{1 - k\rho} \ge 1,$$

contradicting (2.12) of Fact 2.5. This proves (7.1).

The lower bound

(7.2) 
$$\limsup_{T \to \infty} \frac{\inf_{0 \le t \le T - \rho T} \sup_{0 \le s \le \rho T} |Y(t+s) - Y(t)|}{\sqrt{8T \log \log T}} \ge \rho, \quad \text{a.s}$$

follows from the fact that by Theorem B the function  $f(x) = x, 0 \le x \le 1$  is a limit point of

$$\frac{Y(xt)}{\sqrt{8T\log\log T}}$$

and for this function

$$\min_{0 \le x \le 1-\rho} |f(x+\rho) - f(x)| = \rho.$$

This completes the proof of Theorem 1.2(iia).

Now assume that

(7.3) 
$$\lim_{T \to \infty} \frac{a_T (\log \log T)^2}{T} = 0.$$

Define  $\lambda_T$  as in (6.4). Then according to Chung's LIL (cf. Fact 2.6)

(7.4) 
$$|W(\lambda_T)| \ge \frac{\pi}{\sqrt{8}} (1-\varepsilon) \sqrt{\frac{T}{\log \log T}}$$

for  $\varepsilon > 0$  and all T sufficiently large. But according to Fact 2.4,

$$\sup_{0 \le s \le a_T} |W(\lambda_T + s) - W(\lambda_T)|$$
  
$$\le \sqrt{(2 + \varepsilon)a_T(\log(T/a_T) + \log\log T)} \le \sqrt{\frac{(2 + \varepsilon)\varepsilon T}{\log\log T}}.$$

Assuming  $W(\lambda_T) > 0$ , on choosing suitable  $\varepsilon > 0$  we get for some  $c_{26} > 0$ 

$$W(\lambda_T + s) \ge W(\lambda_T) - \sqrt{\frac{(2+\varepsilon)\varepsilon T}{\log\log T}} \ge c_{26}\sqrt{\frac{T}{\log\log T}}.$$

Hence

$$\inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| \le Y(\lambda_T + a_T) - Y(\lambda_T)$$
$$= \int_0^{a_T} \frac{\mathrm{d}s}{W(\lambda_T + s)} \le \frac{a_T}{c_{26}} \sqrt{\frac{\log \log T}{T}}$$

for all large T.

The case when  $W(\lambda_T) < 0$  is similar. This shows the upper bound in (1.16).

For the lower bound we use Fact 2.7: with probability one

(7.5) 
$$g_T \le \frac{T}{(\log \log T)^2}, \quad \max_{0 \le u \le T} |W(u)| \le \frac{\pi}{\sqrt{2}} \sqrt{\frac{T}{\log \log T}}, \quad \text{i.o.}$$

According to Theorem 1.2(i) we have for any  $\varepsilon > 0$  and all large T

(7.6)  

$$\inf_{0 \le t \le T (\log \log T)^{-2}} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| \\
\ge \frac{(K_4 - \varepsilon)a_T}{\sqrt{\left(\frac{T}{(\log \log T)^2} + a_T\right) \log \log T}} \ge \frac{(K_4 - \varepsilon)a_T \sqrt{\log \log T}}{\sqrt{(1 + \varepsilon)T}}.$$

On the other hand, if  $T(\log \log T)^{-2} \le t \le T - a_T$ , and (7.5) is satisfied, then

(7.7) 
$$|Y(t+a_T) - Y(t)| = \int_t^{t+a_T} \frac{\mathrm{d}s}{|W(s)|} \ge \frac{a_T \sqrt{2\log\log T}}{\pi\sqrt{T}}, \quad \text{i.o.}$$

Combining (7.6) and (7.7) for  $\varepsilon > 0$  with probability one

$$\inf_{0 \le t \le T - a_T} \sup_{0 \le s \le a_T} |Y(t+s) - Y(t)| \ge \min\left(\frac{K_4 - \varepsilon}{\sqrt{1 + \varepsilon}}, \frac{\sqrt{2}}{\pi}\right) \frac{a_T \sqrt{\log \log T}}{T}, \quad \text{i.o.}$$

This shows the lower bound in (1.16). The proof of Theorem 1.2(iib) is complete by applying the 0-1 law for Brownian motion.  $\Box$ 

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