# STRONG APPROXIMATION FOR MIXING SEQUENCES WITH INFINITE VARIANCE

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#### Abstract

In this paper we prove a strong approximation result for a mixing sequence with infinite variance and logarithmic decay rate of the mixing coefficient. The result is proved under the assumption that the distribution is symmetric and lies in the domain of attraction of the normal law. Moreover the function  $L(x) = EX^2 1_{\{|X| \le x\}}$  is supposed to be slowly varying with remainder  $(\log x)^{-\alpha}(\log \log)^{-\beta}(x)$  with  $\alpha, \beta > 1$ .

## 1 Introduction

The concept of mixing is a natural generalization of independence and can be viewed as "asymptotic independence": the dependence between two random variables in a mixing sequence becomes weaker as the distance between their indices becomes larger. There is an immense amount of literature dedicated to limit theorems for mixing sequences, most of it assuming that the moments of second order or higher are finite (see e.g. the recent survey article [3]). One of the most important results in this area is Shao's strong invariance principle [14], from which one can easily deduce many other limit theorems.

In this paper we prove a strong approximation result for a mixing sequence of identically distributed random variables with infinite variance, whose distribution is symmetric and lies in the domain of attraction of the normal law (DAN). This suggests that it may be possible to obtain a similar result for the self-normalized sequence. Self-normalized limit theorems have become increasingly popular in the past few years, but so far only the case of independent

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random variables was considered. Therefore, our result may contain the seeds of future research in the promising new area of self-normalized limit theorems for dependent sequences; see e.g. [1], or [12].

Suppose first that  $\{X_n\}_{n\geq 1}$  is a sequence of i.i.d. random variables with  $S_n=\sum_{i=1}^n X_i$  and  $EX=0,\ EX^2=\infty$  (here X denotes a generic random variable with the same distribution as  $X_n$ ). If  $X\in DAN$  (or equivalently, the function  $L(x)=EX^21_{\{|X|\leq x\}}$  is slowly varying), then the "central limit theorem" continues to hold in the form  $S_n/\eta_n\to_d N(0,1)$ , where  $\{\eta_n\}_n$  is a nondecreasing sequence of positive numbers satisfying

$$\eta_n^2 \sim nL(\eta_n).$$
 (1)

(see e.g. [6], IX.8, XVII.5). Moreover, by Theorem 1 of [5], if the distribution of X is symmetric then

$$\limsup_{n \to \infty} \frac{S_n}{(2\eta_n^2 \log \log \eta_n)^{1/2}} = 1 \text{ or } \infty \text{ a.s.}$$

depending on whether the integral

$$I_{\log \log} := \int_{h}^{\infty} \frac{x^2}{L(x) \log \log x} dF(x)$$

converges or diverges (here  $b := \inf\{x \ge 1; L(x) > 0\}$ ). Hence  $I_{\log \log} < \infty$  is a minimum requirement for the "law of the iterated logarithm" in the case of i.i.d. random variables with infinite variance.

In the 1971 Rietz Lecture, Kesten has discussed Feller's result and raised the question of its correctness; see his Remark 9, [8]. Fortunately, he settled this problem, by replacing Feller's normalizing constant  $(\eta_n^2 \log \log \eta_n)^{1/2}$  with a slightly different constant  $\gamma_n$ , which behaves roughly as a root of the equation  $\gamma_n^2 = CnL(\gamma_n) \log \log \gamma_n$  (see Theorem 7). A more general form of the law of the iterated logarithm for the "trimmed" sum  $S_n^{(r)}$  (i.e. the sum obtained by deleting from  $S_n$  the r-th largest terms) has been recently obtained in [9].

Following these lines, Theorem 2.1 of [11] proved that it is possible to obtain (on a larger probability space), the strong approximation

$$S_n - T_n = o(a_n) \quad \text{a.s.} \tag{2}$$

where  $T_n = \sum_{i=1}^n Y_i$  and  $\{Y_n\}_{n\geq 1}$  is a zero-mean Gaussian sequence (with  $EY_n^2 = \tau_n$  for suitable constants  $\tau_n$ ). His rate  $a_n$  is chosen such that

$$a_n^2 \sim nL(a_n)v(a_n),\tag{3}$$

where v is a nondecreasing slowly varying function with  $\lim_{x\to\infty} v(x) = \infty$  and

$$I := I_{v(\cdot)} = \int_b^\infty \frac{x^2}{L(x)v(x)} dF(x) < \infty.$$
 (4)

In this paper we prove that a strong approximation of type (2) continues to hold in the mixing case.

We recall that a sequence  $\{X_n\}_{n\geq 1}$  of random variables is called  $\rho$ -mixing if

$$\rho(n) := \sup_{k \ge 1} \rho(\mathcal{M}_1^k, \mathcal{M}_{k+n}^{\infty}) \to 0 \quad \text{as} \quad n \to \infty$$

where  $\rho(\mathcal{M}_1^k, \mathcal{M}_{k+n}^{\infty}) := \sup\{|\operatorname{Corr}(U, V)|; U \in L^2(\mathcal{M}_k^n), V \in L^2(\mathcal{M}_{k+n}^{\infty})\}$  and  $\mathcal{M}_a^b$  denotes the  $\sigma$ -field generated by  $X_a, X_{a+1}, \dots, X_b$ .

Here is our result.

**Theorem 1** Let  $\{X_n\}_{n\geq 1}$  be a  $\rho$ -mixing sequence of symmetric identically distributed random variables with EX = 0,  $EX^2 = \infty$  and  $X \in DAN$ , where X denotes a random variable with the same distribution as  $X_n$ . Assume that

$$\rho(n) \le C(\log n)^{-r} \text{ for some } r > 1.$$
(5)

Let v be a nondecreasing slowly varying function such that  $v(x) \geq C \log \log x$  for x large; let  $\tau = \min(3, r+1)$ . Suppose that the function  $L(x) = EX^2 1_{\{|X| \leq x\}}$  satisfies (4) and is slowly varying with remainder  $(\log x)^{-\alpha} v^{-\beta}(x)$  for some  $\alpha > \tau/(\tau-2)$ ,  $\beta > \tau/2$ , i.e. for any  $\lambda \in (0,1)$  there exists C > 0 such that

$$(SR) 1 - \frac{L(\lambda x)}{L(x)} \le C(\log x)^{-\alpha} v^{-\beta}(x) \text{ for } x \text{ large.}$$

Then without changing its distribution, we can redefine  $\{X_n\}_{n\geq 1}$  on a larger probability space together with a standard Brownian motion  $W = \{W(t)\}_{t\geq 0}$  such that for some constants  $s_n^2$ 

$$S_n - W(s_n^2) = o(a_n) \quad a.s. \tag{6}$$

where  $\{a_n\}_n$  is a nondecreasing sequence of positive numbers satisfying (3).

Condition (SR) specifies the rate of convergence of  $L(\lambda x)/L(x)$  to 1, for the slowly varying function L (see p.185 of [2]). It was used in *only one* place, namely to ensure the convergence of the sum (28) in the proof of Lemma 11. Unfortunately, we could not avoid it.

We should mention here that a "functional central limit theorem" for  $\rho$ -mixing sequences with infinite variance was obtained by [15] under the condition  $\sum_{n} \rho(2^{n}) < \infty$ . In order to obtain the strong approximation (6) we needed to impose the logarithmic decay rate of  $\rho(n)$ .

The remaining part of the paper is dedicated to the proof of Theorem 1: the description of the general method is given in Section 2, while the technical details are discussed in Sections 3 and 4. Among other ingredients, the proof uses the blocking technique introduced in [13], according to which the original random variables are replaced by their sums over progressively larger blocks of integers (separated by smaller blocks, whose length is also progressively larger).

Throughout this work, C denotes a generic constant that does not depend on n but may be different from place to place. We denote by I(a,b] the measure attributed by the integral I to the interval (a,b]. We let A(x) = L(x)v(x).

## 2 Sketch of Proof

As in [4] we may take

$$\eta_n = \inf\{s \ge b + 1; \frac{L(s)}{s^2} \le \frac{1}{n}\}, \quad a_n = \inf\{s \ge b + 1; \frac{A(s)}{s^2} \le \frac{1}{n}\}.$$

Clearly (1) and (3) hold. We have  $a_n \geq \eta_n$  and

$$a_n^2 \ge C\eta_n^2 v(\eta_n) \ge C\eta_n^2 \log \log \eta_n. \tag{7}$$

Without loss of generality we will assume that  $\eta_n^2 = nL(\eta_n)$  and  $a_n^2 = nA(a_n)$ .

The proof is based on a double truncation technique at levels  $b_n := v^{-p}(a_n)a_n$  and  $a_n$  (which is due to [5]), and a repeated application of the method of [14] on each of the "truncation" intervals  $[0, b_n]$ ,  $(b_n, a_n]$ .

We assume that p > 1/2. Let

$$\hat{X}_n = X_n I_{\{|X_n| \le b_n\}} \quad X'_n = X_n I_{\{b_n < |X_n| \le a_n\}}, \quad \bar{X}_n = X_n I_{\{|X_n| > a_n\}}.$$

By the symmetry assumption  $E\hat{X}_n = EX'_n = 0$ ; since  $EX_n = 0$ , it follows that  $E\bar{X}_n = 0$ . We have  $X_n = \hat{X}_n + X'_n + \bar{X}_n$  and hence

$$S_n = \hat{S}_n + S_n' + \bar{S}_n \tag{8}$$

where  $\hat{S}_n, S'_n, \bar{S}_n$  denote the partial sums of  $\hat{X}_i, X'_i$ , respectively  $\bar{X}_i$ .

By Lemmas 3.2 and 3.3 of [5] (under the *symmetry* assumption), (4) is equivalent to  $\sum_{n\geq 1} P(|X|>\epsilon a_n)<\infty$  for all  $\epsilon>0$ . Hence

$$\bar{S}_n = o(a_n)$$
 a.s. (9)

In Section 3, we show that the central part  $\hat{S}_n$  gives us the approximation

$$\hat{S}_n - W(s_n^2) = o((\eta_n^2 \log \log \eta_n)^{1/2}) \quad \text{a.s.}$$
 (10)

for some constants  $s_n^2$ . In Section 4 we show that between the two truncations we have

$$S_n' = o(a_n) \quad \text{a.s.} \tag{11}$$

The conclusion (6) follows immediately by (7)-(11).

# 3 The Central Part

The goal of this section is to prove relation (10) on a possibly larger probability space on which the sequence  $\{\hat{X}_n\}_n$  is redefined (without changing its distribution). In order to do this, we introduce the blocks  $H_1, I_1, H_2, I_2, \ldots$  of consecutive integers and we decompose the sum  $\hat{S}_n$  into three terms containing the sums over the "small" blocks  $I_i$ , the sums over the "big" blocks  $H_i$ , and the remaining  $\hat{X}_j$ 's (whose sum is shown to be negligible). The idea is to construct the blocks  $I_i$  small enough to make the term depending on these blocks negligible, but large enough to give sufficient space between the blocks  $H_i$ . The sums  $u_i$  over the blocks  $H_i$  will provide us with the desired approximation (10), by applying an almost sure invariance principle (due to [14]) to the martingale differences  $\xi_i = u_i - E(u_i|u_1, \ldots, u_{i-1})$ , after proving that the sum of the terms  $E(u_i|u_1, \ldots, u_{i-1})$  is negligible as well.

We define the blocks  $H_1, I_1, H_2, I_2, \ldots$  of consecutive integers such that

$$\operatorname{card}(H_i) = [ai^{a-1} \exp(i^a)], \quad \operatorname{card}(I_i) = [ai^{a-1} \exp(i^a/2)] \quad \forall i \ge 1$$

with  $a=1/\alpha$ . Note that  $(1-a)\tau>2$ . Let  $N_m:=\sum_{i=1}^m\operatorname{card}(H_i\cup I_i)\sim \exp(m^a)$  and  $N_{m_n}\leq n< N_{m_{n+1}}$ . Clearly  $N_{m_n}\sim n,\, m_n\sim (\log n)^{1/a}$ .

We define

$$u_i = \sum_{j \in H_i} \hat{X}_j, \quad v_i = \sum_{j \in I_i} \hat{X}_j, \quad \xi_i = u_i - E(u_i | \mathcal{G}_{i-1})$$

where  $\mathcal{G}_m = \sigma(\{u_i; i \leq m\})$ , and write

$$\hat{S}_n = \sum_{i=1}^{m_n} v_i + \sum_{j=N_{m_n}+1}^n \hat{X}_j + \sum_{i=1}^{m_n} E(u_i|\mathcal{G}_{i-1}) + \sum_{i=1}^{m_n} \xi_i.$$
 (12)

The first three terms will be of order  $o(\eta_n)$ . The last term will give us the desired approximation with rate  $o((\eta_n^2 \log \log \eta_n)^{1/2})$ .

We begin with two elementary lemmas.

**Lemma 2** There exists C > 0 such that  $b_n \leq C\eta_n$  for n large, and hence

$$nL(b_n) \le C\eta_n^2 \quad for \ n \ large.$$
 (13)

PROOF. The relation  $b_n \leq C\eta_n$  for n large, can be written as  $a_n/\eta_n \leq Cv^p(a_n)$  for n large; using the definitions of  $a_n$  and  $\eta_n$ , this in turn is equivalent to:

$$\frac{L(a_n)}{L(\eta_n)} \le Cv^{2p-1}(a_n) \quad \text{for } n \text{ large.}$$
(14)

Since L is slowly varying, it follows by Potter's Theorem (Theorem 1.5.6.(i) of [2]) that for any  $C > 1, \delta > 0$  we have

$$\frac{L(a_n)}{L(\eta_n)} \le C \left(\frac{a_n}{\eta_n}\right)^{\delta} = C \left(\frac{L(a_n)v(a_n)}{L(\eta_n)}\right)^{\delta/2} \quad \text{for } n \text{ large}$$

and hence

$$\left(\frac{L(a_n)}{L(\eta_n)}\right)^{1-\delta/2} \le Cv^{\delta/2}(a_n)$$
 for  $n$  large.

This is exactly relation (14) with  $\delta = 2 - 1/p$ . Relationship (13) follows using the fact that L is nondecreasing and slowly varying, and the definition of  $\eta_n$ :  $nL(b_n) \leq nL(C\eta_n) \leq CnL(\eta_n) = C\eta_n^2$ .

**Lemma 3** For any integer  $\lambda > 0$  there exists  $C = C_{\lambda} > 0$  such that  $a_{\lambda n} \leq Ca_n$  and  $b_{\lambda n} \leq Cb_n$  for n large, and hence

$$L(a_{\lambda n}) \le CL(a_n) \text{ and } L(b_{\lambda n}) \le CL(b_n) \text{ for } n \text{ large.}$$
 (15)

PROOF. Using the definition of  $a_n$  and Potter's theorem, we get: for any  $C > 1, \delta \in (0,2)$ 

$$\frac{a_{\lambda n}^2}{a_n^2} = \frac{\lambda n A(a_{\lambda n})}{n A(a_n)} \le \lambda C \left(\frac{a_{\lambda n}}{a_n}\right)^{\delta} \quad \text{for } n \text{ large}$$

and hence  $a_{\lambda n}/a_n \leq C\lambda^{1/(2-\delta)}$  for n large. By the definition of  $b_n$  and Potter's theorem, we have: for any  $C > 1, \varepsilon > 0$ 

$$\frac{b_{\lambda n}}{b_n} = \frac{a_{\lambda n}}{a_n} \cdot \left(\frac{v(a_n)}{v(a_{\lambda n})}\right)^p \le C \left(\frac{a_{\lambda n}}{a_n}\right)^{1+p\varepsilon} \le C\lambda^{(1+p\varepsilon)/(2-\delta)} \quad \text{for } n \text{ large.}$$

The last statement in the lemma follows since L is slowly varying.

We are now ready to treat the first three terms in the decomposition (12).

**Lemma 4** We have  $\sum_{i=1}^{m} v_i = o(m^2 \exp(\frac{1}{3}m^a)L^{1/2}(b_{N_m}))$  a.s. and hence  $\sum_{i=1}^{m_n} v_i = o(\eta_n)$  a.s.

PROOF. We have  $Ev_i^2 \leq C \operatorname{card}(I_i) \cdot \max_{j \in I_i} E\hat{X}_j^2 \leq Ci^{a-1} \exp(\frac{1}{2}i^a) L(b_{N_m})$  for all  $i \leq m$ . Hence  $E(\sum_{i=1}^m v_i)^2 \leq m \sum_{i=1}^m Ev_i^2 \leq CmL(b_{N_m}) \exp(\frac{2}{3}m^a)$ . The first statement in the lemma follows by the Chebyshev's inequality and the Borel-Cantelli lemma. The second statement follows using  $m_n \sim (\log n)^{1/a}$  and relation (13).

To simplify the notation, we let  $c_i = \exp(i^a)L(b_{[\exp(i^a)]})$  and  $d_i = \eta_{[\exp(i^a)]}^2$ . By (13),  $c_i \leq Cd_i$  for i large.

**Lemma 5** We have  $\max_{N_m < n \le N_{m+1}} \left| \sum_{j=N_m+1}^n \hat{X}_j \right| = o(c_m^{1/2})$  a.s. and hence  $\max_{N_{m_n} < n \le N_{m_n+1}} \left| \sum_{j=N_{m_n}+1}^n \hat{X}_j \right| = o(\eta_n)$  a.s.

PROOF. The second statement follows by (13). For the first part, it is enough to prove that for any  $\varepsilon > 0$ 

$$\sum_{k\geq 1} P\left(\max_{N_k < n \leq N_{k+1}} \left| \sum_{j=N_k+1}^n \hat{X}_j \right| > \varepsilon c_k^{1/2} \right) < \infty.$$
 (16)

For this we apply Lemma 2.4 of [14] with

$$q = \tau$$
,  $B = k^{-a(\tau+2)/(\tau-2)} c_k^{1/2}$ ,  $x = \varepsilon c_k^{1/2}$ 

$$n = N_{k+1} - N_k$$
,  $m = [k^{-a(\tau+2)/(\tau-2)}e^{k^a}]$ .

For every  $j = N_k + 1, ..., N_{k+1}$  we have  $E\hat{X}_j^2 1_{\{|\hat{X}_j| > B\}} = EX^2 1_{\{B < |X| \le b_j\}} \le L(b_j) \le L(b_{N_{k+1}}) \le CL(b_{[\exp(k^a)]}) = C(xB)/m$ , using (15) for the last inequality. Relation (16) follows as (2.20) of [14] provided we show that:

$$\sum_{k\geq 1} k^{a-1} e^{-(\tau-2)k^a/2} L^{-\tau/2}(b_{[\exp(k^a)]}) E|X|^{\tau} 1_{\{|X|\leq 2b_{[\exp(k^a)]}\}} < \infty.$$
 (17)

Let  $\alpha_j = a_{[\exp(j^a)]}$  and  $\beta_j = b_{[\exp(j^a)]}$ . The sum in (17) becomes

$$\sum_{k\geq 1} k^{a-1} e^{-(\tau-2)k^a/2} L^{-\tau/2}(\beta_k) (E|X|^{\tau} 1_{\{|X|\leq 2\beta_0\}} + \sum_{j=1}^k E|X|^{\tau} 1_{\{2\beta_{j-1}<|X|\leq 2\beta_j\}})$$

$$\leq C \sum_{j\geq 1} E|X|^{\tau} 1_{\{2\beta_{j-1}<|X|\leq 2\beta_j\}} L^{-\tau/2}(\beta_j) e^{-(\tau-2)j^a/2}$$

(19)

$$\leq C \sum_{j>1} I(\beta_{j-1}, \beta_j] \cdot \beta_j^{\tau-2} A(\beta_j) L^{-\tau/2}(\beta_j) e^{-(\tau-2)j^a/2}, \tag{18}$$

where for the last inequality we used:  $E|X|^{\tau}1_{\{a<|X|\leq b\}} \leq I(a,b] \cdot b^{\tau-2}A(b)$ . By Potter's Theorem (Theorem 1.5.6.(i) of [2]):

$$\frac{v(b_n)}{v(a_n)} \le C \left(\frac{b_n}{a_n}\right)^{-\mu} = C v^{p\mu}(a_n),$$

$$\frac{b_n^2}{nL(b_n)} = \frac{L(a_n)}{L(b_n)} v^{-(2p-1)}(a_n) \le C\left(\frac{a_n}{b_n}\right)^{\delta} v^{-(2p-1)}(a_n) = Cv^{-(2p-1-p\delta)}(a_n),$$

for any  $\mu, \delta > 0$  and n large. Hence

$$\beta_j^{\tau-2} A(\beta_j) L^{-\tau/2}(\beta_j) e^{-(\tau-2)j^a/2} = v(\beta_j) \left( \frac{\beta_j^2}{\exp(j^a) L(\beta_j)} \right)^{(\tau-2)/2}$$

$$< C v^{1+p\mu}(\alpha_j) \cdot v^{-(\tau-2)(2p-1-p\delta)/2}(\alpha_j) = C v^{-\gamma}(\alpha_j) < C,$$

where we selected  $\mu, \delta$  such that  $\gamma := -1 - p\mu + (\tau - 2)(2p - 1 - p\delta)/2 > 0$ . From (18), (19) we see that the sum in (17) is smaller than  $C \sum_{j>i} I(\beta_{j-1}, \beta_j] < \infty$ , using (4).

**Lemma 6** We have  $\sum_{i=1}^{m} E(u_i|\mathcal{G}_{i-1}) = o(m^{-(r-1/2)a} \cdot (\log m)^3 \cdot \exp(\frac{1}{2}m^a) \cdot L^{1/2}(b_{N_m}))$  a.s. and hence  $\sum_{i=1}^{m_n} E(u_i|\mathcal{G}_{i-1}) = o(\eta_n)$  a.s.

PROOF. Let  $T_m = \sum_{i=1}^m E(u_i | \mathcal{G}_{i-1})$  and  $\alpha_m = m^{-(r-1/2)a} (\log m)^3 \exp(\frac{1}{2} m^a) \cdot L^{1/2} (b_{N_m})$ . For any  $i \leq m$  we have

$$Eu_i^2 \le C \cdot \operatorname{card}(H_i) \cdot \max_{j \in H_i} E\hat{X}_j^2 \le Ci^{a-1} \exp(i^a) L(b_{N_m}). \tag{20}$$

By (2.26) of [14] and (5), we get:  $E(\max_{l \leq m} T_l^2) \leq C(\log m)^4 m^{-2ar} \cdot \exp(m^a) L(b_{N_m})$ . Let  $m_k = [k^{1/a}]$ . Using Chebyshev's inequality we get

$$\sum_{k>1} P(\max_{l \le m_k} |T_l| > \varepsilon \alpha_{m_k}) \le \sum_{k>1} \frac{E(\max_{l \le m_k} T_l^2)}{\varepsilon^2 \alpha_{m_k}^2} \le C \sum_{k>1} \frac{1}{m_k^a (\log m_k)^2} < \infty.$$

From here we conclude that

$$\sum_{k \geq 1} P(T_{m_k} > \varepsilon \alpha_{m_k}) < \infty \quad \text{and} \quad \sum_{k \geq 1} P(\max_{m_{k-1} < m \leq m_k} |T_m - T_{m_k}| > \varepsilon \alpha_{m_k}) < \infty$$

for all  $\varepsilon > 0$ . By the Borel-Cantelli lemma, it follows that

$$\frac{T_{m_k}}{\alpha_k} \to 0$$
 a.s.,  $\max_{m_{k-1} < m \le m_k} \frac{|T_m - T_{m_k}|}{\alpha_{m_k}} \to 0$  a.s.

and hence  $T_m/\alpha_m \to 0$  a.s.

Our last theorem gives us the desired approximation for the last term in (12). To prove this theorem we need two lemmas. Let  $\sigma_i^{*2} = E\xi_i^2$ ,  $s_m^{*2} = \sum_{i=1}^m \sigma_i^{*2}$ ,  $s_n^2 = s_{m_n}^{*2}$ .

**Lemma 7** We have  $\sum_{i>1} d_i^{-\tau/2} E|\xi_i|^{\tau} < \infty$ .

PROOF. It is enough to prove the lemma with  $c_i$  instead of  $d_i$ , and  $u_i$  instead of  $\xi_i$ . By Lemma 2.3 of [14] we have

$$E|u_i|^{\tau} \le C\{(\operatorname{card}(H_i))^{\tau/2} \cdot \max_{j \in H_i} (E\hat{X}_j^2)^{\tau/2} + \operatorname{card}(H_i) \cdot \max_{j \in H_i} E|\hat{X}_j|^{\tau}\} \le$$

$$C\left\{ (i^{a-1} \exp(i^a))^{\tau/2} L^{\tau/2} (b_{[\exp(i^a)]}) + i^{a-1} \exp(i^a) E|X|^{\tau} 1_{\{|X| \le 2b_{[\exp(i^a)]}\}} \right\} = Cc_i^{\tau/2} \left\{ i^{-(1-a)\tau/2} + i^{a-1} e^{-(\tau-2)i^a/2} L^{-\tau/2} (b_{[\exp(i^a)]}) E|X|^{\tau} 1_{\{|X| \le 2b_{[\exp(i^a)]}\}} \right\}.$$
(21)

The lemma follows by (17).

**Lemma 8** We have  $\sum_{i=1}^{m} (E(\xi_i^2 | \mathcal{G}_{i-1}) - E\xi_i^2) = o(d_m)$  a.s.

PROOF. It is enough to prove the lemma with  $c_m$  instead of  $d_m$ . Let  $u_i^* = u_i^2 1_{\{|u_i| \le c_i^{1/2}\}}$  and  $u_i^{**} = u_i^2 1_{\{|u_i| > c_i^{1/2}\}}$ . The conclusion will follow from:

$$\sum_{i=1}^{m} (E(u_i^{**}|\mathcal{G}_{i-1}) + Eu_i^{**}) = o(c_m) \quad \text{a.s.}$$
(22)

$$U_m := \sum_{i=1}^{m} (E(u_i^* | \mathcal{G}_{i-1}) - Eu_i^*) = o(m^{-(r-1/2)a} (\log m)^3 c_m) \quad \text{a.s.}$$
 (23)

$$\sum_{i=1}^{m} (E^{2}(u_{i}|\mathcal{G}_{i-1}) + EE^{2}(u_{i}|\mathcal{G}_{i-1})) = o(m^{-(2r-1)a}(\log m)^{2}c_{m}) \quad \text{a.s.}$$
 (24)

To prove (22), note that  $E|u_i|^{\tau} \ge E|u_i|^{\tau} 1_{\{|u_i| > c_i^{1/2}\}} \ge c_i^{(\tau-2)/2} Eu_i^{**}$ . Relationship (22) follows by Kronecker's lemma, (21) and (17).

To prove (23), let  $\beta_m = m^{-(r-1/2)a} (\log m)^3 c_m$ . For any  $i \leq m$ 

$$Eu_i^{*2} = Eu_i^4 1_{\{|u_i| < c_i^{1/2}\}} \le c_i Eu_i^2 \le Ci^{a-1} \exp(i^a) L(b_{[\exp(m^a)]}) c_m$$

where we used (20) in the last inequality. By (2.34) of [14] and (5),we get:  $E(\max_{l \leq m} U_l^2) \leq C(\log m)^4 m^{-2ar} \exp(m^a) L(b_{[\exp(m^a)]}) c_m$ . Let  $m_k = [k^{1/a}]$ . Using the same argument based on a subsequence convergence criterion as in the proof of Lemma 6, we get  $U_m = o(\beta_m)$  a.s. It remains to prove (24). By the mixing property, (20) and (5), we have  $EE^2(u_i|\mathcal{G}_{i-1}) \leq Ci^{-(2r-1)a-1} \exp(i^a) L(b_{[\exp_i a]}) = Ci^{-(2r-1)a-1} c_i$ . Relation (24) follows by the Kronecker's lemma

Here is the main result of this section.

**Theorem 9** Without changing its distribution, we can redefine the sequence  $\{\xi_i\}_{i\geq 1}$  on a larger probability space together with a standard Brownian motion  $W = \{W(t)\}_{t\geq 0}$  such that

$$\sum_{i=1}^{m_n} \xi_i - W(s_n^2) = o((\eta_n^2 \log \log \eta_n)^{1/2}) \quad a.s.$$

PROOF. By Theorem 2.1 of [14], Lemma 7 and Lemma 8, we can redefine the sequence  $\{\xi_i\}_{i\geq 1}$  on a larger probability space together with a standard Brownian motion  $W = \{W(t)\}_{t\geq 0}$  such that

$$\sum_{i=1}^{m} \xi_i - W(s_m^{*2}) = o(\left\{ d_m \left( \log \frac{s_m^{*2}}{d_m} + \log \log d_m \right) \right\}^{1/2}) \quad \text{a.s.}$$
 (25)

Using the mixing property, (20) and (13), we obtain:

$$s_m^{*2} = \sum_{i=1}^m Eu_i^2 - \sum_{i=1}^m E(u_i)E(u_i|\mathcal{G}_{i-1}) \le C\sum_{i=1}^m Eu_i^2 \le C\exp(m^a)L(b_{N_m}) \le C\eta_{N_m}^2$$

The result now follows by noting that  $d_{m_n} = \eta_n^2$ .

#### 4 Between the Two Truncations

This section is dedicated to the proof of relation (11):  $S'_n/a_n \to 0$  a.s. For this we consider the same blocks  $H_i$ ,  $I_i$  as in Section 3 and we decompose the sum  $S'_n$  into three components, depending on the sums over the blocks  $I_i$ , the sums over the blocks  $H_i$  and the remaining terms  $X'_j$ . The sums  $u'_i$  over the blocks  $H_i$  are once again approximated by the corresponding martingale differences  $\xi'_i$  and relation (11) follows by a martingale subsequence convergence criterion.

Let  $H_1, I_1, H_2, I_2, \ldots$  be the blocks introduced in Section 3. We define

$$u'_i = \sum_{j \in H_i} X'_j, \quad v'_i = \sum_{j \in I_i} X'_j, \quad \xi'_i = u'_i - E(u'_i | \mathcal{G}'_{i-1})$$

where  $\mathcal{G}'_m = \sigma(\{u'_i; i \leq m\})$ , and write

$$S'_{n} = \sum_{i=1}^{m_{n}} v'_{i} + \sum_{j=N_{m_{n}}+1}^{n} X'_{j} + \sum_{i=1}^{m_{n}} E(u'_{i}|\mathcal{G}'_{i-1}) + \sum_{i=1}^{m_{n}} \xi'_{i}.$$
 (26)

We will prove that all the 4 terms in the above decomposition are of order  $o(a_n)$ .

We begin by treating the first three terms. Note that  $EX_j^{'2} = L(a_j) - L(b_j) \leq L(a_j)$  and  $nL(a_n) \leq Ca_n^2$ .

**Lemma 10** We have  $\sum_{i=1}^{m} v_i' = o(m^2 \exp(\frac{1}{3}m^a)L^{1/2}(a_{N_m}))$  a.s. and hence  $\sum_{i=1}^{m_n} v_i' = o(a_n)$  a.s.

Proof. Same argument as in Lemma 4.

**Lemma 11** We have  $\max_{N_m < n \le N_{m+1}} \left| \sum_{j=N_m+1}^n X_j' \right| = o(\exp(\frac{1}{2}m^a)L^{1/2}(a_{[\exp(m^a)]}))$  a.s. and hence  $\max_{N_{m_n} < n \le N_{m_n+1}} \left| \sum_{j=N_{m_n}+1}^n X_j' \right| = o(a_n)$  a.s.

PROOF. Using the same argument as in Lemma 5, it suffices to show that

$$\sum_{k>1} k^{a-1} e^{-(\tau-2)k^a/2} L^{-\tau/2} (a_{[\exp(k^a)]}) E|X|^{\tau} 1_{\{|X| \le 2a_{[\exp(k^a)]}\}} < \infty.$$
 (27)

Let  $n_j = [\exp(j^a)]$  and  $\alpha_j = a_{n_j}$ . Note that the sum in (27) is smaller than

$$C\sum_{j>1} E|X|^{\tau} 1_{\{2\alpha_{j-1}<|X|\leq 2\alpha_j\}} L^{-\tau/2}(\alpha_j) e^{-(\tau-2)j^a/2} \le$$

$$C\sum_{j\geq 1} (L(2\alpha_j) - L(2\alpha_{j-1})) \cdot \alpha_j^{\tau-2} L^{-\tau/2}(\alpha_j) e^{-(\tau-2)j^a/2},$$

where we used the inequality:  $E|X|^{\tau}1_{\{a<|X|< b\}} \leq (L(b)-L(a))b^{\tau-2}$ . Note that

$$\alpha_j^{\tau-2} L^{-\tau/2}(\alpha_j) e^{-(\tau-2)j^a/2} = L^{-1}(\alpha_j) \left( \frac{\alpha_j^2}{\exp(j^a) L(\alpha_j)} \right)^{(\tau-2)/2}$$

$$\leq CL^{-1}(2\alpha_j) \cdot v^{(\tau-2)/2}(\alpha_j).$$

Since  $\alpha_j \sim \alpha_{j-1}$ , we have  $2\alpha_{j-1} \geq \alpha_j$  for j large. We conclude that the sum in (27) is smaller than

$$C\sum_{j>1} \left[ 1 - \frac{L(\alpha_j)}{L(2\alpha_j)} \right] v^{(\tau-2)/2}(\alpha_j). \tag{28}$$

Using (SR) and the fact that  $\alpha_j \geq C n_i^{1/2}$  and  $v(x) \geq C \log \log x$ , we get

$$\sum_{j\geq 1} \left(1 - \frac{L(\alpha_j)}{L(2\alpha_j)}\right) v^{(\tau-2)/2}(\alpha_j) \leq C \sum_{j\geq 1} (\log \alpha_j)^{-1/a} v^{-d}(\alpha_j) \leq$$

$$C\sum_{j>1} (\log n_j)^{-1/a} (\log \log n_j)^{-d} \le C\sum_{j>1} j^{-1} (\log j)^{-d} < \infty,$$

where  $d := \beta - (\tau - 2)/2 > 1$  (and we recall that  $a = 1/\alpha$ ). This concludes the proof of (27).

**Lemma 12** We have  $\sum_{i=1}^{m} E(u'_{i}|\mathcal{G}'_{i-1}) = o(m^{-(r-1/2)a} \cdot (\log m)^{3} \cdot \exp(\frac{1}{2}m^{a}) \cdot L^{1/2}(a_{N_{m}}))$  a.s. and hence  $\sum_{i=1}^{m_{n}} E(u'_{i}|\mathcal{G}'_{i-1}) = o(a_{n})$  a.s.

Proof. Same argument as in Lemma 6.

For our last result, we will need the following martingale subsequence convergence criterion (which is probably well-known).

**Lemma 13** Let  $\{S_n, \mathcal{F}_n\}_{n\geq 1}$  be a zero-mean martingale and  $\{a_n\}_{n\geq 1}$  a nondecreasing sequence of positive numbers with  $\lim_n a_n = \infty$ . If there exists a subsequence  $\{n_k\}_k$  such that  $a_{n_k}/a_{n_{k-1}} \leq C$  for all k, and

$$\sum_{k>1} \frac{E|S_{n_k} - S_{n_{k-1}}|^p}{a_{n_k}^p} < \infty \quad \text{for some } p \in [1, 2],$$
(29)

then  $S_n = o(a_n)$  a.s.

PROOF. Note that  $\{S_{n_k}, \mathcal{F}_{n_k}\}_{k\geq 1}$  is a martingale. From (29) it follows that  $S_{n_k}/a_{n_k} \to 0$  a.s. (see Theorem 2.18 of [7]). By the extended Kolmogorov inequality (see p. 65 of [10]), we have

$$\sum_{k\geq 1} P(\max_{n_{k-1} < n \leq n_k} |S_n - S_{n_k}| > \varepsilon a_{n_k}) \leq \sum_{k\geq 1} \frac{E|S_{n_k} - S_{n_{k-1}}|^p}{\varepsilon^p a_{n_k}^p} < \infty$$

for every  $\varepsilon > 0$ , and hence  $T_k := a_{n_k}^{-1} \max_{n_{k-1} < n \le n_k} |S_n - S_{n_k}| \to 0$  a.s. Finally for  $n_{k-1} < n \le n_k$  we have:

$$\frac{|S_n|}{a_n} \leq \frac{|S_{n_{k-1}}|}{a_{n_{k-1}}} + \frac{|S_n - S_{n_{k-1}}|}{a_{n_{k-1}}} \leq \frac{|S_{n_{k-1}}|}{a_{n_{k-1}}} + \frac{a_{n_k}}{a_{n_{k-1}}} \cdot T_k \leq \frac{|S_{n_{k-1}}|}{a_{n_{k-1}}} + C \cdot T_k \to 0 \quad \text{a.s.}$$

Finally, we treat the last term in the decomposition (26).

Theorem 14 We have

$$\sum_{i=1}^{m_n} \xi_i' = o(a_n) \quad a.s.$$

PROOF. Let  $U_n := \sum_{i=1}^{m_n} \xi_i'$  and note that  $\{U_n, \mathcal{G}_{m_n}'\}_{n \geq 1}$  is a zero-mean martingale. By Lemma 13, it is enough to prove that for a suitable subsequence  $\{n_k\}_k$  we have

$$\sum_{k\geq 1} \frac{E|U_{n_k} - U_{n_{k-1}}|^2}{a_{n_k}^2} < \infty. \tag{30}$$

Similarly to the proof of Lemma 2.3 of [11], we take a subsequence  $\{n_k\}_k$  satisfying  $n_k \sim n_{k-1}(1+\phi^{-1}(k))$ , where the function  $\phi$  is chosen such that  $\lim_{k\to\infty} \phi(k) = \infty$  and

$$\frac{1}{\phi(k)+1} \cdot I(b_{n_k}, a_{n_k}] \le CI(a_{n_{k-1}}, a_{n_k}]. \tag{31}$$

Clearly  $n_k \sim n_{k+1}$  and hence  $a_{n_k} \sim a_{n_{k+1}}$  and  $b_{n_k} \sim b_{n_{k+1}}$ . We proceed now with the proof of (30). Let

$$Z_k := U_{n_k} - U_{n_{k-1}} = \sum_{m_{n_{k-1}} < i \le m_{n_k}} \xi_i'.$$

By the martingale property

$$EZ_k^2 = \sum_{m_{n_{k-1}} < i \le m_{n_k}} E\xi_i^{'2} \le (m_{n_k} - m_{n_{k-1}}) \max_{m_{n_{k-1}} < i \le m_{n_k}} E\xi_i^{'2}.$$
(32)

Using Lemma 2.3 of [14] we have: for every  $m_{n_{k-1}} < i \le m_{n_k}$ ,

$$E\xi_i^{'2} \le Eu_i^{'2} \le Ci^{a-1}e^{i^a} \cdot \max_{j \in H_i} EX_j^{'2} \le C(\log n_k)^{(a-1)/a}n_k \cdot \max_{j \in H_i} EX_j^{'2}. \tag{33}$$

Now for any  $j \in H_i$  and  $m_{n_{k-1}} < i \le m_{n_k}$  we have

$$EX_{i}^{'2} \le A(a_{i})I(b_{i}, a_{i}) \le A(a_{N_{i}})I(b_{N_{i-1}}, a_{N_{i}}) \le CA(a_{n_{k}})I(b_{n_{k}}, a_{n_{k}}). \tag{34}$$

Using (34) and (33) we get: for every  $m_{n_{k-1}} < i \le m_{n_k}$ ,

$$E\xi_i^{'2} \le C(\log n_k)^{(a-1)/a} n_k \cdot A(a_{n_k}) I(b_{n_k}, a_{n_k}] = C(\log n_k)^{(a-1)/a} a_{n_k}^2 I(b_{n_k}, a_{n_k}].$$
(35)

From (32) and (35) and recalling that  $m_n \sim (\log n)^{1/a}$ , we get

$$\frac{EZ_k^2}{a_{n_k}^2} \leq C[(\log n_k)^{1/a} - (\log n_{k-1})^{1/a}] \cdot (\log n_k)^{(a-1)/a} I(b_{n_k}, a_{n_k}] 
\leq C(\log n_{k-1})^{(1-a)/a} \frac{1}{n_{k-1}} (n_k - n_{k-1}) \cdot (\log n_k)^{(a-1)/a} I(b_{n_k}, a_{n_k}] 
= C \frac{n_k - n_{k-1}}{n_{k-1}} I(b_{n_k}, a_{n_k}] \leq C \frac{1}{\phi(k) + 1} I(b_{n_k}, a_{n_k}] \leq C I(a_{n_{k-1}}, a_{n_k}],$$

where we used the inequality  $f(y) - f(x) \le f'(x)(y-x)$  for the concave function  $f(x) = (\log x)^{1/a}$  for the second inequality, and the choice (31) of the function  $\phi$  for the last inequality. Relationship (30) follows by (4).

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