



## AN ENTROPY POWER INEQUALITY FOR THE BINOMIAL FAMILY

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ABSTRACT. In this paper, we prove that the classical Entropy Power Inequality, as derived in the continuous case, can be extended to the discrete family of binomial random variables with parameter  $1/2$ .

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### 1. INTRODUCTION

The continuous Entropy Power Inequality

$$(1.1) \quad e^{2h(X)} + e^{2h(Y)} \leq e^{2h(X+Y)}$$

was first stated by Shannon [1] and later proved by Stam [2] and Blachman [3]. Later, several related inequalities for continuous variables were proved in [4], [5] and [6]. There have been several attempts to provide discrete versions of the Entropy Power Inequality: in the case of Bernoulli sources with addition modulo 2, results have been obtained in a series of papers [7], [8], [9] and [11].

In general, inequality (1.1) does not hold when  $X$  and  $Y$  are discrete random variables and the differential entropy is replaced by the discrete entropy: a simple counterexample is provided when  $X$  and  $Y$  are deterministic.

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In what follows,  $X_n \sim B\left(n, \frac{1}{2}\right)$  denotes a binomial random variable with parameters  $n$  and  $\frac{1}{2}$ , and we prove our main theorem:

**Theorem 1.1.** *The sequence  $X_n$  satisfies the following Entropy Power Inequality*

$$\forall m, n \geq 1, \quad e^{2H(X_n)} + e^{2H(X_m)} \leq e^{2H(X_n+X_m)}.$$

With this aim in mind, we use a characterization of the superadditivity of a function, together with an entropic inequality.

## 2. SUPERADDITIVITY

**Definition 2.1.** A function  $n \mapsto Y_n$  is superadditive if

$$\forall m, n \quad Y_{m+n} \geq Y_m + Y_n.$$

A sufficient condition for superadditivity is given by the following result.

**Proposition 2.1.** *If  $\frac{Y_n}{n}$  is increasing, then  $Y_n$  is superadditive.*

*Proof.* Take  $m$  and  $n$  and suppose  $m \geq n$ . Then by assumption

$$\frac{Y_{m+n}}{m+n} \geq \frac{Y_m}{m}$$

or

$$Y_{m+n} \geq Y_m + \frac{n}{m}Y_m.$$

However, by the hypothesis  $m \geq n$

$$\frac{Y_m}{m} \geq \frac{Y_n}{n}$$

so that

$$Y_{m+n} \geq Y_m + Y_n.$$

□

In order to prove that the function

$$(2.1) \quad Y_n = e^{2H(X_n)}$$

is superadditive, it suffices then to show that function  $n \mapsto \frac{Y_n}{n}$  is increasing.

## 3. AN INFORMATION THEORETIC INEQUALITY

Denote as  $B \sim \text{Ber}(1/2)$  a Bernoulli random variable so that

$$(3.1) \quad X_{n+1} = X_n + B$$

and

$$(3.2) \quad P_{X_{n+1}} = P_{X_n} * P_B = \frac{1}{2}(P_{X_n} + P_{X_{n+1}}),$$

where  $P_{X_n} = \{p_k^n\}$  denotes the probability law of  $X_n$  with

$$(3.3) \quad p_k^n = 2^{-n} \binom{n}{k}.$$

A direct application of an equality by Topsøe [12] yields

$$(3.4) \quad H(P_{X_{n+1}}) = \frac{1}{2}H(P_{X_{n+1}}) + \frac{1}{2}H(P_{X_n}) + \frac{1}{2}D(P_{X_{n+1}}||P_{X_{n+1}}) + \frac{1}{2}D(P_{X_n}||P_{X_{n+1}}).$$

Introduce the Jensen-Shannon divergence

$$(3.5) \quad \mathcal{JSD}(P, Q) = \frac{1}{2}D\left(P \left\| \frac{P+Q}{2}\right.\right) + \frac{1}{2}D\left(Q \left\| \frac{P+Q}{2}\right.\right)$$

and remark that

$$(3.6) \quad H(P_{X_n}) = H(P_{X_{n+1}}),$$

since each distribution is a shifted version of the other. We conclude thus that

$$(3.7) \quad H(P_{X_{n+1}}) = H(P_{X_n}) + \mathcal{JSD}(P_{X_{n+1}}, P_{X_n}),$$

showing that the entropy of a binomial law is an increasing function of  $n$ . Now we need the stronger result that  $\frac{Y_n}{n}$  is an increasing sequence, or equivalently that

$$(3.8) \quad \log \frac{Y_{n+1}}{n+1} \geq \log \frac{Y_n}{n}$$

or

$$(3.9) \quad \mathcal{JSD}(P_{X_{n+1}}, P_{X_n}) \geq \frac{1}{2} \log \frac{n+1}{n}.$$

We use the following expansion of the Jensen-Shannon divergence, due to B.Y. Ryabko and reported in [13].

**Lemma 3.1.** *The Jensen-Shannon divergence can be expanded as follows*

$$\mathcal{JSD}(P, Q) = \frac{1}{2} \sum_{\nu=1}^{\infty} \frac{1}{2\nu(2\nu-1)} \Delta_{\nu}(P, Q)$$

with

$$\Delta_{\nu}(P, Q) = \sum_{i=1}^n \frac{|p_i - q_i|^{2\nu}}{(p_i + q_i)^{2\nu-1}}.$$

This lemma, applied in the particular case where  $P = P_{X_n}$  and  $Q = P_{X_{n+1}}$  yields the following result.

**Lemma 3.2.** *The Jensen-Shannon divergence between  $P_{X_{n+1}}$  and  $P_{X_n}$  can be expressed as*

$$\mathcal{JSD}(P_{X_{n+1}}, P_{X_n}) = \sum_{\nu=1}^{\infty} \frac{1}{\nu(2\nu-1)} \cdot \frac{2^{2\nu-1}}{(n+1)^{2\nu}} m_{2\nu} \left( B \left( n+1, \frac{1}{2} \right) \right),$$

where  $m_{2\nu} \left( B \left( n+1, \frac{1}{2} \right) \right)$  denotes the order  $2\nu$  central moment of a binomial random variable  $B \left( n+1, \frac{1}{2} \right)$ .

*Proof.* Denote  $P = p_i$ ,  $Q = p_i^+$  and  $\bar{p}_i = (p_i + p_i^+)/2$ . For the term  $\Delta_{\nu}(P_{X_{n+1}}, P_{X_n})$  we have

$$\begin{aligned} \Delta_{\nu}(P_{X_{n+1}}, P_{X_n}) &= \sum_{i=1}^n \frac{|p_i^+ - p_i|^{2\nu}}{(p_i^+ + p_i)^{2\nu-1}} \\ &= 2 \sum_{i=1}^n \left( \frac{p_i^+ - p_i}{p_i^+ + p_i} \right)^{2\nu} \bar{p}_i \end{aligned}$$

and

$$\begin{aligned} \frac{p_i^+ - p_i}{p_i^+ + p_i} &= \frac{2^{-n} \binom{n}{i-1} - 2^{-n} \binom{n}{i}}{2^{-n} \binom{n}{i-1} + 2^{-n} \binom{n}{i}} \\ &= \frac{2i - n - 1}{n + 1} \end{aligned}$$

so that

$$\begin{aligned}\Delta_\nu(P_{X_{n+1}}, P_{X_n}) &= 2 \sum_{i=1}^n \left( \frac{2i - n - 1}{n + 1} \right)^{2\nu} \bar{p}_i \\ &= 2 \left( \frac{2}{n + 1} \right)^{2\nu} \sum_{i=1}^n \left( i - \frac{n + 1}{2} \right)^{2\nu} \bar{p}_i \\ &= \frac{2^{2\nu+1}}{(n + 1)^{2\nu}} m_{2\nu} \left( B \left( n + 1, \frac{1}{2} \right) \right).\end{aligned}$$

Finally, the Jensen-Shannon divergence becomes

$$\begin{aligned}\mathcal{JSD}(P_{X_{n+1}}, P_{X_n}) &= \frac{1}{4} \sum_{\nu=1}^{+\infty} \frac{1}{\nu(2\nu - 1)} \Delta_\nu(P_{X_{n+1}}, P_{X_n}) \\ &= \sum_{\nu=1}^{+\infty} \frac{1}{\nu(2\nu - 1)} \cdot \frac{2^{2\nu-1}}{(n + 1)^{2\nu}} m_{2\nu} \left( B \left( n + 1, \frac{1}{2} \right) \right).\end{aligned}$$

□

#### 4. PROOF OF THE MAIN THEOREM

We are now in a position to show that the function  $n \curvearrowright \frac{Y_n}{n}$  is increasing, or equivalently that inequality (3.9) holds.

*Proof.* We remark that it suffices to prove the following inequality

$$(4.1) \quad \sum_{\nu=1}^3 \frac{1}{\nu(2\nu - 1)} \cdot \frac{2^{2\nu-1}}{(n + 1)^{2\nu}} m_{2\nu} \left( B \left( n + 1, \frac{1}{2} \right) \right) \geq \frac{1}{2} \log \left( 1 + \frac{1}{n} \right)$$

since the terms  $\nu > 3$  in the expansion of the Jensen-Shannon divergence are all non-negative. Now an explicit computation of the three first even central moments of a binomial random variable with parameters  $n + 1$  and  $\frac{1}{2}$  yields

$$m_2 = \frac{n + 1}{4}, \quad m_4 = \frac{(n + 1)(3n + 1)}{16} \quad \text{and} \quad m_6 = \frac{(n + 1)(15n^2 + 1)}{64},$$

so that inequality (4.1) becomes

$$\frac{1}{60} \frac{30n^4 + 135n^3 + 245n^2 + 145n + 37}{(n + 1)^5} \geq \frac{1}{2} \log \left( 1 + \frac{1}{n} \right).$$

Let us now upper-bound the right hand side as follows

$$\log \left( 1 + \frac{1}{n} \right) \leq \frac{1}{n} - \frac{1}{2n^2} + \frac{1}{3n^3}$$

so that it suffices to prove that

$$\frac{1}{60} \cdot \frac{30n^4 + 135n^3 + 245n^2 + 145n + 37}{(n + 1)^5} - \frac{1}{2} \left( \frac{1}{n} - \frac{1}{2n^2} + \frac{1}{3n^3} \right) \geq 0.$$

Rearranging the terms yields the equivalent inequality

$$\frac{1}{60} \cdot \frac{10n^5 - 55n^4 - 63n^3 - 55n^2 - 35n - 10}{(n + 1)^5 n^3} \geq 0$$

which is equivalent to the positivity of polynomial

$$P(n) = 10n^5 - 55n^4 - 63n^3 - 55n^2 - 35n - 10.$$

Assuming first that  $n \geq 7$ , we remark that

$$P(n) \geq 10n^5 - n^4 \left( 55 + \frac{63}{6} + \frac{55}{6^2} + \frac{35}{6^3} + \frac{10}{6^4} \right) = \left( 10n - \frac{5443}{81} \right) n^4$$

whose positivity is ensured as soon as  $n \geq 7$ .

This result can be extended to the values  $1 \leq n \leq 6$  by a direct inspection at the values of function  $n \curvearrowright \frac{Y_n}{n}$  as given in the following table.

$n$	1	2	3	4	5	6
$\frac{e^{2H(X_n)}}{n}$	4	4	4.105	4.173	4.212	4.233

Table 4.1: Values of the function  $n \curvearrowright \frac{Y_n}{n}$  for  $1 \leq n \leq 6$ .

□

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