

Journal URL

<http://www.math.washington.edu/~ejpecp/>

Paper URL

<http://www.math.washington.edu/~ejpecp/EjpVol7/paper21.abs.html>

## THE NOISE MADE BY A POISSON SNAKE

**Jon Warren**

Department of Statistics, University of Warwick

Coventry, CV4 7AL, United Kingdom

[warren@stats.warwick.ac.uk](mailto:warren@stats.warwick.ac.uk)

**Abstract:** The purpose of this article is to study a coalescing flow of sticky Brownian motions on  $[0, \infty)$ . Sticky Brownian motion arises as the weak solution of a stochastic differential equation, and the study of flow reveals the nature of the extra randomness that must be added to the driving Brownian motion. This can be represented in terms of Poissonian marking of the trees associated with the excursions of Brownian motion. We also study the noise, in the sense of Tsirelson, generated by the flow. It is shown that this noise is not generated by any Brownian motion, even though it is predictable.

**Keywords and phrases:** stochastic flow, sticky Brownian motion, coalescence, stochastic differential equation, noise.

**AMS subject classification (2000):** 60H10, 60J60.

Submitted to EJP on September 2, 1999. Final version accepted on June 24, 2002.

# 1 INTRODUCTION

Suppose that  $(\Omega, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$  is a filtered probability space satisfying the usual conditions, and that  $(Z_t; t \geq 0)$  is a continuous, adapted process taking values in  $[0, \infty)$  which satisfies the stochastic differential equation

$$Z_t = x + \int_0^t 1_{(Z_s > 0)} dW_s + \theta \int_0^t 1_{(Z_s = 0)} ds, \quad (1.1)$$

where  $(W_t; t \geq 0)$  is a real-valued  $\mathcal{F}_t$ -Brownian motion and  $x \geq 0$  and  $\theta \in (0, \infty)$  are constants. We say that  $Z$  is sticky Brownian motion with parameter  $\theta$  started from  $x$ . Sticky Brownian motion arose in the work of Feller [6] on strong Markov processes taking values in  $[0, \infty)$  that behave like Brownian motion away from 0. The parameter  $\theta$  determines the stickiness of zero: the cases (which we usually exclude)  $\theta = 0$  and  $\theta = \infty$  correspond respectively to Brownian motion absorbed or instantaneously reflected on hitting zero. For  $\theta \in (0, \infty)$  sticky Brownian motion can be constructed quite simply as a time change of reflected Brownian motion so that the resulting process is slowed down at zero, and so spends a non-zero amount of (real) time there. However here our interest will be focused on it arising as a solution of the stochastic differential equation (1.1). This equation does not admit a strong solution, in order to construct  $Z$  it is necessary to add to the driving Brownian motion  $W$  some extra ‘randomness’. The nature of this randomness was first investigated in [15]. More recently it has been shown (Warren, [16], and Watanabe, [17], following Tsirelson, [13]) that the filtration  $(\mathcal{F}_t)_{t \geq 0}$  cannot be generated by *any* Brownian motion.

The main object of study in this paper is a coalescing flow on  $\mathbb{R}_+$  which we denote by  $(Z_{s,t}; 0 \leq s \leq t < \infty)$ , where each  $Z_{s,t}$  is an increasing function from  $\mathbb{R}_+$  to itself. We may describe the flow rather informally in terms of the motion of particles. For any  $t$  and  $x$  the trajectory  $h \mapsto Z_{t+h}(x)$  describes the motion of a particle which starts at time  $t$  from a position  $x$ . We shall be concerned with a flow in which this motion is determined by a sticky Brownian motion. Away from zero, particles move parallel to each other- which means that the SDEs (similar to equation (1.1)) which describe their trajectories are all driven by the same Brownian motion  $W$ . Particles collide while visiting zero and thereafter they move together- and the map  $Z_{s,t}$  is typically neither injective nor surjective.

The stochastic differential equation (1.1) was studied in [15]. There it was shown that the solution is weak, and that for each  $t$ , assuming  $Z_0 = 0$ , the conditional distribution of  $Z_t$  given the entire path of the driving Brownian motion  $W$  is given by:

$$\mathbb{P}(Z_t \leq z | \sigma(W_s; s \geq 0)) = \exp(-2\theta(W_t + L_t - z)) \quad z \in [0, W_t + L_t], \quad (1.2)$$

where  $L_t = \sup_{s \leq t} (-W_s)$ .

An interpretation of (1.2) was given in [15] which we can summarise loosely as follows. Define a process  $\xi$  via

$$\xi_t = W_t + \sup_{s \leq t} (-W_s) \quad t \geq 0. \quad (1.3)$$

$\xi$  is then a Brownian motion reflected from zero. Each excursion of  $\xi$  from zero gives rise to a rooted tree. Each time  $t$  determines a path in the tree corresponding to the excursion of  $\xi$  which straddles  $t$ . This path starts from the root of the tree and is of length  $\xi_t$ . If times  $s$  and  $t$  are both straddled by the same excursion then the paths they determine co-incide from the root for a length  $I_{s,t} = \inf_{h \in [s,t]} \xi_h$ , and thereafter differ. The reader not familiar with the correspondence between trees and excursions may consult Aldous [1] or [2]. Next

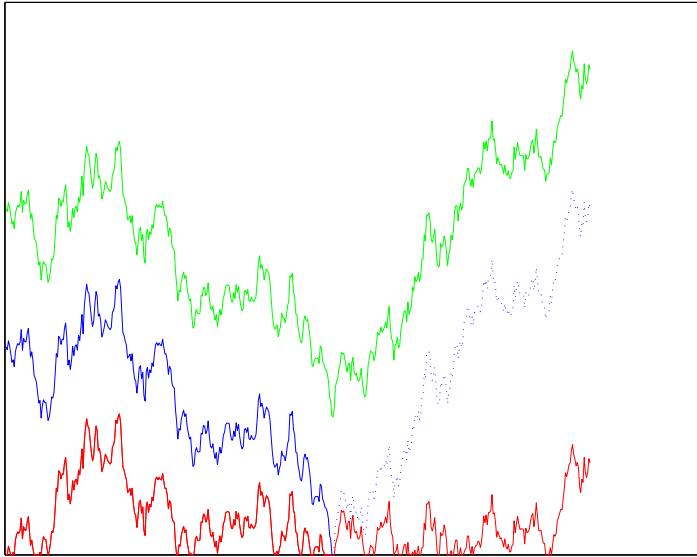


Figure 1: The trajectories of three particles in a typical realization of the flow. The upper particle never hits zero and its trajectory is a translate of the driving Brownian motion. The middle particle has started from a critical height- it just hits zero and from then on follows the same trajectory as the lowest particle which was started from zero. The dotted line shows the path that would be taken by the middle particle if the boundary were not sticky.

we add marks to these trees according to a Poisson point process of intensity  $2\theta$ . We can then construct the sticky Brownian motion  $Z$  from these marked trees via, for each  $t \geq 0$ ,

$$Z_t = \begin{cases} 0 & \text{if the path corresponding to } t \text{ is not marked,} \\ \xi_t - h_t & \text{otherwise,} \end{cases} \quad (1.4)$$

where  $h_t$  is the distance along the path corresponding to  $t$  from the root to the first mark. Notice that not all the marks carried by the trees are used in constructing the process  $Z$  in this way.

In this paper the description of the previous paragraph is extended to the flow  $(Z_{s,t}; s \leq t)$  of sticky Brownian motions. By replacing the single sticky Brownian motion by an entire flow we make natural use of all the marks on the trees. A discrete version of the correspondence between a flow and a family of marked trees is illustrated in Figures 2 and 3.

Section 2 of the paper details the construction of the flow. We begin with a process  $(X_t; t \geq 0)$ , taking values in  $\mathbb{R}^\infty$  which solves an infinite family of stochastic differential equations. From  $X$  we are able to construct the flow in a pathwise manner.

In Section 3 we show that a generalization of (1.4) holds for the process  $X$ . There exists a family of marked trees such that the components  $X^{(1)}, X^{(2)}, \dots, X^{(k)}, \dots$  of  $X$  satisfy

$$X_t^{(k)} = \begin{cases} 0 & \text{if the path corresponding to } t \text{ carries fewer than } k \text{ marks,} \\ \xi_t - h_t^{(k)} & \text{otherwise,} \end{cases} \quad (1.5)$$

where  $h_t^{(k)}$  is the distance along the path corresponding to  $t$  from the root to the  $k$ th mark. A consequence of this description of the process  $X$  is that it may be identified with a simple path-valued Markov process similar in nature to the Brownian snake of Le Gall [8]; the difference being that the white noise along the branches of the tree is replaced with Poisson noise. For an interpretation in terms of killing in superprocesses, see Watanabe [18]. A

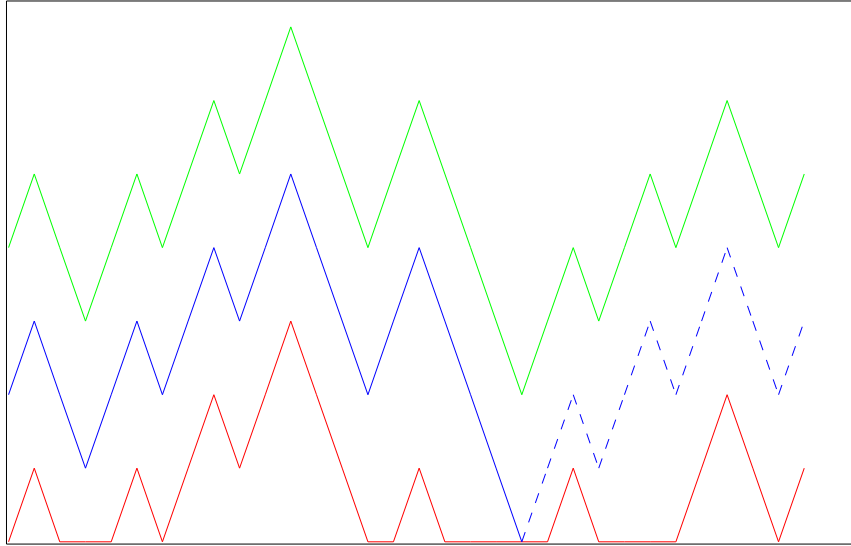


Figure 2: Three typical trajectories for a discrete version of the flow. As in Figure 1 the dashed line is to illustrate the effect of the sticky boundary.

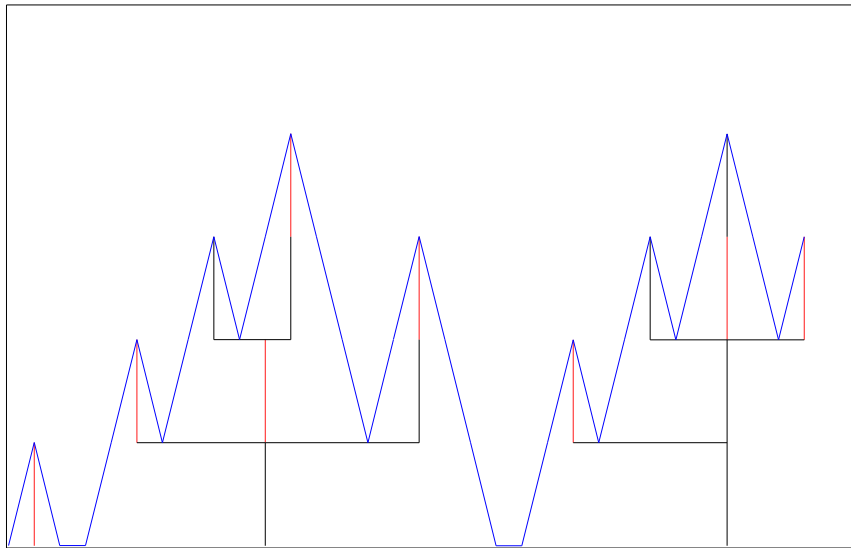


Figure 3: Showing some marked trees consistent with Figure 2. The driving random walk determines a family of trees. The vertical edges of the trees are marked by independently tossing a coin for each edge. This carries the information about the stickiness of the boundary. A particle at the origin at time  $n$  will move to level 1 at time  $n + 1$  if the driving walk makes an upward step between times  $n$  and  $n + 1$  and the corresponding edge carries a mark.

closely related construction involving Poisson marking of trees is described by Aldous and Pitman [3].

In the final section we consider the flow as a noise, in the sense of Tsirelson [11] and [12]. The flow  $Z$  has independent stationary increments, in that, for any  $0 \leq t_0 \leq t_1 \leq \dots \leq t_n$  the random maps  $Z_{t_0, t_1}, Z_{t_1, t_2}, \dots, Z_{t_{n-1}, t_n}$  are independent and for  $s \leq t$  and  $h > 0$  the maps  $Z_{s, t}$  and  $Z_{s+h, t+h}$  have the same distribution. We will write  $\mathcal{F}_{s, t}$  for the  $\sigma$ -algebra generated by  $(Z_{u, v}; s \leq u \leq v \leq t)$ . Following Tsirelson the doubly indexed family  $(\mathcal{F}_{s, t}; s \leq t)$  is called a noise. It is a non-white noise because, as we shall see, there is no Brownian motion whose increments generate it.

## 2 CONSTRUCTING THE FLOW

Let  $\mathbb{R}^\infty$  denote the space of real-valued sequences with the product topology. We will denote (for reasons shortly to become clear) the canonical process on the space  $\mathcal{C}(\mathbb{R}_+, \mathbb{R}^\infty)$  of continuous  $\mathbb{R}^\infty$ -valued paths by  $X_t = (W_t, X_t^{(1)}, X_t^{(2)}, \dots)$ . Often we use the notation  $X_t^{(0)} = W_t + \sup_{s \leq t} (-W_s)$ .

**Theorem 1.** *Let  $\theta \in (0, \infty)$ . There exists a unique probability measure  $\mathbb{P}$  on  $\mathcal{C}(\mathbb{R}_+, \mathbb{R}^\infty)$  such that the canonical process  $(X_t; t \geq 0)$  satisfies the following.*

- The first component,  $W$ , is a Brownian motion starting from 0 with respect to the filtration generated by  $X$ .
- For  $k \geq 1$ , the processes  $X^{(k)}$  are non-negative and the following SDEs are satisfied:

$$\begin{aligned} X_t^{(1)} &= \int_0^t 1_{\{X_s^{(1)} > 0\}} dW_s + \theta \int_0^t 1_{\{X_s^{(1)} = 0\}} ds, \\ X_t^{(k)} &= \int_0^t 1_{\{X_s^{(k)} > 0, \dots, X_s^{(1)} > 0\}} dW_s + \theta \int_0^t 1_{\{X_s^{(k)} = 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} ds, \quad k \geq 2. \end{aligned}$$

*Proof.* We begin by proving that  $\mathbb{P}$  is unique. Let  $\mathbb{P}^{(n)}$  denote the restriction of  $\mathbb{P}$  to the  $\sigma$ -algebra  $\mathcal{F}^{(n)}$  generated by  $W, X^{(1)}, \dots, X^{(n)}$ . Since the union of such  $\sigma$ -algebras is an algebra generating the entire Borel  $\sigma$ -algebra, it suffices to prove that each  $\mathbb{P}^{(n)}$  is unique.

Let, for any  $k \geq 1$ ,

$$A_t^{(k)} = \int_0^t 1_{\{X_s^{(k)} = 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} ds, \quad A_t^{(k+)} = \int_0^t 1_{\{X_s^{(k)} > 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} ds,$$

and denote their right-continuous inverses by  $\alpha^{(k)}$  and  $\alpha^{(k+)}$  respectively. We shall see shortly that  $A_\infty^{(k)} = A_\infty^{(k+)} = \infty$  with probability one, and so we may define, for any  $k \geq 1$  and for  $0 \leq u < \infty$ ,

$$\begin{aligned} B_u^{(k)} &= \int_0^{\alpha_u^{(k)}} 1_{\{X_s^{(k)} = 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} dW_s, \\ B_u^{(k+)} &= \int_0^{\alpha_u^{(k+)}} 1_{\{X_s^{(k)} > 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} dW_s. \end{aligned}$$

For each  $n$ , under  $\mathbb{P}^{(n)}$ , the processes  $B^{(1)}, \dots, B^{(n)}, B^{(n+)}$  must be independent Brownian motions, as a consequence of Knight's theorem on orthogonal martingales (see [9], Chapter

V, Theorem 1.9). Hence if we can recover  $W, X^{(1)}, \dots, X^{(n)}$  from them, the measure  $\mathbb{P}^{(n)}$  will be determined uniquely.

Now, we may write,

$$X_t^{(k)} = B_{A_t^{(k+)}}^{(k+)} + \theta A_t^{(k)}.$$

Since the lefthandside is non-negative, and  $A^{(k)}$  grows only when it is zero, we recall that the Lemma of Skorokhod (see [9], Chapter VI, Lemma 2.1) tells us that,

$$\theta A_t^{(k)} = S_{A_t^{(k+)}}^{(k)},$$

where we denote,

$$S_u^{(k)} = \sup_{h \leq u} (-B_h^{(k+)}).$$

Note that,

$$A_t^{((k-1)+)} = A_t^{(k)} + A_t^{(k+)} = \frac{1}{\theta} S_{A_t^{(k+)}}^{(k)} + A_t^{(k+)}.$$

When  $k = 1$  we may take the lefthandside to be given by  $A_t^{(0+)} = t$ . We can now check the earlier claim that  $A_\infty^{(k)} = A_\infty^{(k+)} = \infty$ . If, for some  $k$ , we had  $A_\infty^{(k+)} < \infty$  then repeated application of the preceding equation would give  $A_\infty^{(0+)} < \infty$  - a falsehood. So we deduce that  $A_\infty^{(k+)} = \infty$ , but then we have  $A_\infty^{(k)} = S_\infty^{(k)}/\theta = \infty$  with probability one.

Denote by  $\sigma^{(k)}$  the inverse of the continuous, strictly increasing function:

$$u \mapsto \frac{1}{\theta} S_u^{(k)} + u.$$

Then, we find that,

$$(*) \quad B_t^{((k-1)+)} = B_{\sigma_t^{(k)}}^{(k+)} + B_{t-\sigma_t^{(k)}}^{(k)},$$

and

$$(**) \quad A_t^{(k+)} = \sigma_{A_t^{((k-1)+)}}^{(k)}.$$

These two relations hold even for  $k = 1$  provided that we take  $B_t^{(0+)} = W_t$  and  $A_t^{(0+)} = t$ . Now for any fixed  $n$  we can use (\*) to recover  $B^{(0+)}, \dots, B^{(n+)}$  from  $B^{(1)}, \dots, B^{(n)}, B^{(n+)}$ . Moreover we can then use (\*\*) to obtain  $A^{(1+)}, \dots, A^{(n+)}$ , and then, finally, we obtain  $X^{(1)}, \dots, X^{(n)}$ , since

$$(***) \quad X_t^{(k)} = B_{A_t^{(k+)}}^{(k+)} + S_{A_t^{(k+)}}^{(k)}.$$

More precisely we are able to write, for any finite set of times  $t_1, t_2, \dots, t_m$ , the  $m(n+1)$ -tuple  $W_{t_1}, X_{t_1}^{(1)}, \dots, X_{t_1}^{(n)}, \dots, \dots, W_{t_m}, X_{t_m}^{(1)}, \dots, X_{t_m}^{(n)}$  as a jointly measurable function of  $B^{(1)}, \dots, B^{(n)}, B^{(n+)}$ , and this determines  $\mathbb{P}^{(n)}$  on a generating  $\pi$ -system. This completes the proof of uniqueness.

To prove existence we construct, for each  $n$  the probability measure  $\mathbb{P}^{(n)}$  on  $(\mathcal{C}(\mathbb{R}_+, \mathbb{R}^\infty), \mathcal{F}^{(n)})$  from the law of an  $(n+1)$ -tuple of processes  $W, X^{(1)}, \dots, X^{(n)}$ . This

sequence of measures is consistent (by virtue of the uniqueness property just obtained), and we may take  $\mathbb{P}$  to be its projective limit.

Fix  $n \geq 1$ . To construct  $W, X^{(1)}, \dots, X^{(n)}$  we start with  $n + 1$  independent Brownian motions denoted by  $B^{(1)}, \dots, B^{(n)}, B^{(n+)}$  and, as in the proof of uniqueness, define, in turn,  $B^{((n-1)+)}, B^{((n-2)+)}, \dots, B^{(0+)}$ , by means of (\*). Then, starting from  $A_t^{(0+)} = t$ , we define via (\*\*) increasing processes  $A^{(1+)}, \dots, A^{(n+)}$ , and finally we use (\*\*\*) to obtain the processes  $X^{(1)}, \dots, X^{(n)}$  and take  $W = B^{(0+)}$ . We must show that the processes so constructed satisfy the appropriate SDEs, and that  $W$  is a Brownian motion. We do so via an inductive argument.

Suppose for some  $k$  we have established that for  $r = 1, 2, \dots, (k-1)$  the process  $X^{(r)}$  satisfies the appropriate SDE and that  $A_t^{(r+)} = \int_0^t \mathbf{1}_{\{X_s^{(r)} > 0, X_s^{(r-1)} > 0, \dots, X_s^{(1)} > 0\}} ds$ . Then, using (\*\*) and (\*\*\*),

$$\begin{aligned} \int_0^t \mathbf{1}_{\{X_s^{(k)} > 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} ds &= \int_0^t \mathbf{1}_{\{X_s^{(k)} > 0\}} dA_s^{((k-1)+)} = \\ &= \int_0^{A_t^{((k-1)+)}} \mathbf{1}_{\{B_{\sigma_v^{(k)}}^{(k+)} + S_{\sigma_v^{(k)}}^{(k)} > 0\}} dv = \int_0^{A_t^{(k+)}} \mathbf{1}_{\{B_u^{(k+)} + S_u^{(k)} > 0\}} \left( \frac{1}{\theta} dS_u^{(k)} + du \right) = A_t^{(k+)}, \end{aligned}$$

because the measure  $dS_u^{(k)}$  is carried by the set  $\{u : B_u^{(k+)} + S_u^{(k)} = 0\}$ . A similar calculation shows

$$\begin{aligned} \int_0^t \mathbf{1}_{\{X_s^{(k)} = 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} ds &= \int_0^t \mathbf{1}_{\{X_s^{(k)} = 0\}} dA_s^{((k-1)+)} = \\ &= \int_0^{A_t^{(k+)}} \mathbf{1}_{\{B_u^{(k+)} + S_u^{(k)} = 0\}} \left( \frac{1}{\theta} dS_u^{(k)} + du \right) = \frac{1}{\theta} S_{A_t^{(k+)}}^{(k)}. \end{aligned}$$

Call the common value  $A_t^{(k)}$  and note that this identifies the second term on the righthand-side of (\*\*\*). It remains to identify the first term as a stochastic integral against  $W$ . Write  $W_t = M_t + N_t$  where  $N_t = B_{A_t^{(k+)}}^{(k+)}$  and  $M_t = B_{A_t^{(1)}}^{(1)} + \dots + B_{A_t^{(k)}}^{(k)}$  and observe that  $dM_t dM_t = dA_t^{(1)} + \dots + dA_t^{(k)}$  while  $dN_t dN_t = dA_t^{(k+)}$ . Next we can write:

$$\begin{aligned} \int_0^t \mathbf{1}_{\{X_s^{(k)} > 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} dW_s &= \int_0^t \mathbf{1}_{\{X_s^{(k)} > 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} dM_s + \\ &= \int_0^t \mathbf{1}_{\{X_s^{(k)} > 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} dN_s. \end{aligned}$$

The first integral must be identically zero because the zero set of the integrand supports the measure  $dM_t dM_t$  while the second integral is simply  $B_{A_t^{(k+)}}^{(k+)}$ . Applying the above arguments for  $k = 1, 2, \dots, n$  we deduce that all the necessary SDEs are satisfied.

The remaining issue is to verify that  $W$  is a Brownian motion with respect to the filtration generated by  $W, X^{(1)}, \dots, X^{(n)}$ . Suppose that  $B^{(k+)}$  is a Brownian motion with respect to some filtration  $\mathcal{G}^{(k)}$  and that  $\mathcal{G}_\infty^{(k)}$  is independent of the Brownian motion  $B^{(k)}$ . Let  $\mathcal{G}^{(k-1)}$  be the filtration generated by

$$\sigma \left( B_{s-\sigma_s^{(k)}}^{(k)}; s \leq t \right) \text{ and } \mathcal{G}_{\sigma_t^{(k)}}^{(k)}.$$

Then it is easy to check that both the time changed Brownian motions  $\left( B_{t-\sigma_t^{(k)}}^{(k)}; t \geq 0 \right)$  and  $\left( B_{\sigma_t^{(k)}}^{(k+)}; t \geq 0 \right)$  are  $\mathcal{G}_t^{(k-1)}$ -martingales and hence that  $B^{((k-1)+)}$ , defined using (\*) is also.

But  $B^{((k-1)+)}$ , having quadratic variation process  $t$ , is thus a  $\mathcal{G}_t^{(k-1)}$ -Brownian motion by Lévy's characterization. If we apply this argument successively, firstly for  $k = n$  and  $\mathcal{G}^{(n)}$  the natural filtration of  $B^{(n)}$  then for  $k = n - 1, \dots, 2, 1$  we find that  $W = B^{(0)}$  is a  $\mathcal{G}^{(0)}$ -Brownian motion. But the processes  $W, X^{(1)}, \dots, X^{(n)}$  are all  $\mathcal{G}^{(0)}$ -adapted and we are finished.  $\square$

In the following we will work with the probability space given by this theorem, and with respect to the filtration denoted by  $(\mathcal{F}_t)_{t \geq 0}$ , the smallest right-continuous and  $\mathbb{P}$ -complete filtration to which  $X$  is adapted. Since we want to construct the flow  $Z$  from  $X$  in a pathwise manner we need to restrict ourselves to a subset of  $\mathcal{C}(\mathbb{R}_+, \mathbb{R}^\infty)$  on which  $X$  is sufficiently well behaved. Our first task is to identify such a subset.

For each  $t$  let  $\hat{N}_t$  denote the smallest  $k \geq 0$  such that  $X_t^{(k+1)}$  is zero, or be infinity if no such  $k$  exists. As we shall see the process  $(\hat{N}_t; t \geq 0)$  must be treated with respect; it is very singular. For any  $s \leq t$  let  $\hat{N}_{s,t} = \inf\{\hat{N}_h : s \leq h \leq t\}$ .

**Lemma 2.** *The following hold with probability one.*

- For every  $t \geq 0$  the sequence  $(X_t^{(k)})_{k \geq 0}$  is decreasing and if  $X_t^{(k)} = X_t^{(k+1)}$  then  $X_t^{(k)} = X_t^{(k+1)} = 0$ .
- $\hat{N}_t < \infty$  except for a set of  $t$  having Lebesgue measure zero, and  $\hat{N}_{s,t} < \infty$  for all  $s < t$ .

*Proof.* We continue to use notation as introduced in the proof of the preceding theorem. We will show that, for all  $u$  and  $k$ ,

$$B_u^{((k-1)+)} + S_u^{(k-1)} \geq B_{\sigma_u^{(k)}}^{(k+)} + S_{\sigma_u^{(k)}}^{(k)},$$

with equality only if both sides are 0. Then by taking  $u = A_t^{((k-1)+)}$  we deduce

$$X_t^{(k-1)} \geq X_t^{(k)}$$

for all  $t$ , and with equality only if both sides are 0.

Recall that

$$\begin{aligned} B_t^{((k-1)+)} &= B_{\sigma_t^{(k)}}^{(k+)} + B_{t-\sigma_t^{(k)}}^{(k)} \\ &= B_{\sigma_t^{(k)}}^{(k+)} + S_{\sigma_t^{(k)}}^{(k)} + B_{t-\sigma_t^{(k)}}^{(k)} - S_{\sigma_t^{(k)}}^{(k)} \\ &= \left( B_{\sigma_t^{(k)}}^{(k+)} + S_{\sigma_t^{(k)}}^{(k)} \right) + \left( B_{t-\sigma_t^{(k)}}^{(k)} - \theta(t - \sigma_t^{(k)}) \right). \end{aligned}$$

In view of the definition of  $\sigma_t^{(k)}$  we have that  $t - \sigma_t^{(k)}$ , and hence the second bracketed term above, is constant on each component of the set

$$\mathcal{S} = \left\{ t : B_{\sigma_t^{(k)}}^{(k+)} + S_{\sigma_t^{(k)}}^{(k)} > 0 \right\}.$$

Consequently

$$S_t^{(k-1)} = \sup_{h \leq t - \sigma_t^{(k)}} \left( -(B_h^{(k)} - \theta h) \right),$$



and hence

$$B_t^{((k-1)+)} + S_t^{(k-1)} \geq B_{\sigma_t^{(k)}}^{(k+)} + S_{\sigma_t^{(k)}}^{(k)}.$$

Moreover we only obtain equality here if  $t - \sigma_t^{(k)}$  is an instant at which  $h \mapsto B_h^{(k)} - \theta h$  attains a new minimum. Since  $B^{(k)}$  is independent of  $B^{(k+)}$  and  $\sigma^{(k)}$  the probability of this occurring for  $t$  belonging to  $\mathcal{S}$  is zero.

Let us prove the second assertion of the lemma: that  $\hat{N}_{s,t}$  is finite for all  $s < t$  with probability one. On applying Itô's formula we find that, for any  $\lambda > 0$ , and any  $k \geq 1$ ,

$$\begin{aligned} \exp\{-\lambda X_t^{(k)}\} - \exp\{-\lambda X_0^{(k)}\} &= \\ -\lambda \int_0^t \exp\{-\lambda X_s^{(k)}\} dX_s^{(k)} + \frac{1}{2}\lambda^2 \int_0^t \exp\{-\lambda X_s^{(k)}\} dX_s^{(k)} dX_s^{(k)} &= \\ -\lambda \int_0^t \exp\{-\lambda X_s^{(k)}\} 1_{\{X_s^{(k)} > 0, \dots, X_s^{(1)} > 0\}} dW_s & \\ -\theta\lambda \int_0^t \exp\{-\lambda X_s^{(k)}\} 1_{\{X_s^{(k)} = 0, X_s^{(k-1)} = 0, \dots, X_s^{(1)} > 0\}} ds & \\ + \frac{1}{2}\lambda^2 \int_0^t \exp\{-\lambda X_s^{(k)}\} 1_{\{X_s^{(k)} > 0, \dots, X_s^{(1)} > 0\}} ds. & \end{aligned}$$

Thus

$$\exp\{-\lambda X_t^{(k)} - \frac{1}{2}\lambda^2 t\} + \int_0^t e^{-\frac{1}{2}\lambda^2 s} \left\{ \theta\lambda dA_s^{(k)} + \frac{1}{2}\lambda^2 (dA_s^{(k)} + \dots + dA_s^{(1)}) \right\}$$

is a bounded martingale. Taking expectations, and letting  $t \rightarrow \infty$  we calculate that:

$$\mathbb{E} \left[ \int_0^\infty e^{-\frac{1}{2}\lambda^2 s} dA_s^{(k)} \right] = \frac{1}{\theta\lambda} \left( \frac{2\theta}{2\theta + \lambda} \right)^k.$$

Summing over  $k$ , since

$$\sum_{k=1}^\infty \frac{1}{\theta\lambda} \left( \frac{2\theta}{2\theta + \lambda} \right)^k = \frac{2}{\lambda^2},$$

we deduce that, with probability one, for all  $t$ ,

$$\sum_{k=1}^\infty A_t^{(k)} = t.$$

But the measure of the set  $\{t : \hat{N}_t = k - 1\}$  is equal to  $A_t^{(k)}$  for  $0 \leq k < \infty$ , and so the set  $\{t : \hat{N}_t = \infty\}$  has measure zero, and a complement dense in  $\mathbb{R}_+$ .  $\square$

It is not true that with probability one  $\hat{N}_t$  is finite for all  $t$ . For each  $k$ , the set  $\{t : \hat{N}_t > k\}$  is open and dense, and thus by virtue of Baire's Category Theorem the set  $\{t : \hat{N}_t = \infty\} = \bigcap_k \{t : \hat{N}_t > k\}$  is dense.

We now describe a result that will eventually lead to the independent increments property of the flow. If  $W(1)$  and  $W(2)$  are independent Brownian motions starting from zero and  $t_0$  some fixed time, then the process  $W$  defined by

$$W_t = \begin{cases} W_t(1) & \text{if } t \leq t_0, \\ W_{t_0}(1) + W_{t-t_0}(2) & \text{if } t \geq t_0, \end{cases}$$

and obtained by splicing  $(W_t(2); t \geq 0)$  onto the end of the path  $(W_t(1); 0 \leq t \leq t_0)$  is itself a Brownian motion. A more complicated procedure is given in the following lemma which generalizes this construction.

**Lemma 3.** *Suppose we are given two independent processes  $X(1)$  and  $X(2)$  each distributed according to the law given by Theorem 1. Fix some  $t_0$ , and construct a process  $X$  as follows. For each  $x \geq 0$  let*

$$T(x) = t_0 + \inf\{t : W_t(2) = -x\},$$

and let  $T_k = T(X_{t_0}^{(k)}(1))$  for  $k \geq 1$ , and  $T_0 = \infty$ . Define  $W$  via

$$W_t = \begin{cases} W_t(1) & \text{if } t \leq t_0, \\ W_{t_0}(1) + W_{t-t_0}(2) & \text{if } t \geq t_0. \end{cases}$$

Similarly for  $k \geq 1$ , take  $X_t^{(k)} = X_t^{(k)}(1)$  for  $t \leq t_0$  and then let

$$X_t^{(k)} = X_{t_0}^{(k)}(1) + W_t - W_{t_0} \quad \text{for } t_0 \leq t \leq T_k.$$

Continue by letting

$$X_t^{(k)} = X_{t-t_0}^{(r)}(2) \quad \text{when } t \in [T_{k-r+1}, T_{k-r}] \quad r = 1, 2, \dots, k.$$

Then the process  $X$  is also distributed according to the law determined by Theorem 1.

*Proof.* As remarked previously  $W$  is a Brownian motion, and indeed it is a Brownian motion with respect to the filtration generated by  $X$  since any increment  $W_u - W_t$  is independent of  $(X_s; 0 \leq s \leq t)$ . Note also that the times  $T_k$  are stopping times with respect to this filtration.

Observe that  $t \mapsto X_t^{(k)}$  is continuous- in fact at the times  $T_k, T_{k-1}, \dots, T_1$ , at which continuity is in doubt, it is zero. This is because, for any  $r$ , we have  $X_{T_r-t_0}^{(0)}(2) = 0$  and so by Lemma 2:

$$X_{T_r-t_0}^{(k)}(2) = 0 \quad k \geq 0.$$

Consider a time interval  $(T_{m+1}, T_m)$  (this may be empty-it does not matter). For  $s$  belonging to this interval we have  $X_s^{(1)} > X_s^{(2)} > \dots > X_s^{(m)} > 0$ , and so, for any  $k = m + r > m$ ,

$$\mathbf{1}_{\{X_s^{(k)} > 0, \dots, X_s^{(1)} > 0\}} = \mathbf{1}_{\{X_s^{(k)} > 0, \dots, X_s^{(m+1)} > 0\}} = \mathbf{1}_{\{X_{s-t_0}^{(r)}(2) > 0, \dots, X_{s-t_0}^{(1)}(2) > 0\}},$$

and likewise,

$$\mathbf{1}_{\{X_s^{(k)} = 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} = \mathbf{1}_{\{X_{s-t_0}^{(r)}(2) = 0, X_{s-t_0}^{(r-1)}(2) > 0, \dots, X_{s-t_0}^{(1)}(2) > 0\}}.$$

Now making use of the SDE satisfied by  $X^{(r)}(2)$  we obtain,

$$\begin{aligned} dX_s^{(k)} &= dX_{s-t_0}^{(r)}(2) = \\ & \mathbf{1}_{\{X_{s-t_0}^{(r)}(2) > 0, \dots, X_{s-t_0}^{(1)}(2) > 0\}} dW_{s-t_0}(2) + \theta \mathbf{1}_{\{X_{s-t_0}^{(r)}(2) = 0, X_{s-t_0}^{(r-1)}(2) > 0, \dots, X_{s-t_0}^{(1)}(2) > 0\}} ds \\ &= \mathbf{1}_{\{X_s^{(k)} > 0, \dots, X_s^{(1)} > 0\}} dW_s + \theta \mathbf{1}_{\{X_s^{(k)} = 0, X_s^{(k-1)} > 0, \dots, X_s^{(1)} > 0\}} ds. \end{aligned}$$

Putting together the intervals  $(T_{m+1}, T_m)$  for  $m = k-1, \dots, 0$  we find that  $X^{(k)}$  satisfies the appropriate SDE and the result follows from the uniqueness assertion of Theorem 1.  $\square$

Notice how to recover  $(X_t(1); 0 \leq t \leq t_0)$  and  $(X_t(2); t \geq 0)$  from  $X$ . The former is simply the restriction of  $X$  to times  $t \leq t_0$ . While  $W(2)$  is given by  $W_t(2) = W_{t+t_0} - W_{t_0}$  the recovery of  $X^{(r)}(2)$  for  $r \geq 1$  is more involved. Since  $T_m = \inf\{t \geq t_0 : X_t^{(m)} = 0\}$  if  $t_0 \leq t < T_m$  then  $X_t^{(m)} > 0$  and

$$\hat{N}_t = \inf\{k : X_t^{(k+1)} = 0\} \geq m.$$

Also  $\hat{N}_{T_{m+1}} \leq m$  since  $X_{T_{m+1}}^{(m+1)} = 0$ , and so for  $t \in [T_{m+1}, T_m)$  we have

$$\hat{N}_{t_0,t} = \inf\{\hat{N}_h : t_0 \leq h \leq t\} = m.$$

Consequently we have

$$X_t^{(\hat{N}_{t_0,t+r})} = X_{t-t_0}^{(r)}(2) \quad r \geq 1.$$

The special case  $r = 1$  deserves attention:

$$X_t^{(\hat{N}_{t_0,t+1})} = X_{t-t_0}^{(1)}(2)$$

which as  $t \geq t_0$  varies is a sticky Brownian motion starting from zero at time  $t = t_0$ . This is the key to the construction of the flow from  $X$ .

**Theorem 4.** Define for all  $0 \leq s < t < \infty$ , functions  $Z_{s,t} : \mathbb{R}_+ \mapsto \mathbb{R}_+$ , by

$$Z_{s,t}(x) = \begin{cases} X_t^{(\hat{N}_{s,t+1})} & \text{for } 0 \leq x \leq U_{s,t}, \\ W_t - W_s + x & \text{for } U_{s,t} < x. \end{cases}$$

where  $\sup_{s \leq h \leq t} (W_s - W_h)$  is denoted by  $U_{s,t}$ . We complete the definition of  $Z$  by taking  $Z_{t,t}$  to be the identity for all  $t$ .

Then with probability one  $Z$  possesses the properties of a flow: for all  $0 \leq s \leq t \leq u < \infty$ ,

$$Z_{s,u} = Z_{t,u} \circ Z_{s,t}.$$

For all  $x \in \mathbb{R}_+$  and any  $s \geq 0$ ,

$$t \mapsto Z_{s,t}(x) \text{ is continuous as } t \geq s \text{ varies.}$$

Moreover this trajectory satisfies the stochastic differential equation,

$$Z_{s,t}(x) = x + \int_s^t 1_{(Z_{s,u}(x) > 0)} dW_u + \theta \int_s^t 1_{(Z_{s,u}(x) = 0)} du.$$

The doubly indexed process  $(Z_{s,t}; 0 \leq s \leq t < \infty)$  has independent and stationary increments: for any  $0 \leq t_0 \leq t_1 \leq \dots \leq t_n$  the random maps  $Z_{t_0,t_1}, Z_{t_1,t_2}, \dots, Z_{t_{n-1},t_n}$  are independent and for  $s \leq t$  and  $h > 0$  the maps  $Z_{s,t}$  and  $Z_{s+h,t+h}$  have the same distribution.

*Proof.* Fix  $s < t < u$ . We evaluate the composition  $Z_{t,u}(Z_{s,t}(x))$  in four distinct and exhaustive cases. First note the following estimate,

$$(*) \quad X_t^{(\hat{N}_{s,t+1})} \leq W_t - W_s + U_{s,t}.$$

which we will verify below. Note also the following property of  $\hat{N}$  which is an easy consequence of its definition:

$$(**) \quad \hat{N}_{s,u} = \min\{\hat{N}_{s,t}, \hat{N}_{t,u}\}.$$

Let

$$T = \inf\{h \geq t : X_h^{(\hat{N}_{s,t}+1)} = 0\}.$$

CASE 1. Suppose that  $x > U_{s,t}$  and  $W_t - W_s + x > U_{t,u}$ . These conditions may be interpreted as saying that a particle started from  $x$  at time  $s$  does not reach zero by time  $u$ . Then,

$$Z_{t,u}(Z_{s,t}(x)) = Z_{t,u}(W_t - W_s + x) = W_u - W_s + x.$$

CASE 2. Suppose that  $x \leq U_{s,t}$  and  $X_t^{(\hat{N}_{s,t}+1)} > U_{t,u}$ . This time the particle starting from  $x$  at time  $s$  hits zero before time  $t$  but then does not visit zero during the interval  $[t, u]$ . We have  $T > u$ , and so for  $1 \leq k \leq \hat{N}_{s,t} + 1$ , and  $h \in [t, u]$ , we have  $X_h^{(k)} > 0$ , so (recall always Lemma 2)  $\hat{N}_{t,u} \geq \hat{N}_{s,t} + 1$  and whence by virtue of  $(**)$   $\hat{N}_{s,u} = \hat{N}_{s,t}$  and,

$$Z_{t,u}(Z_{s,t}(x)) = Z_{t,u}(X_t^{(\hat{N}_{s,t}+1)}) = X_t^{(\hat{N}_{s,t}+1)} + W_u - W_t = X_u^{(\hat{N}_{s,t}+1)} = X_u^{(\hat{N}_{s,u}+1)}.$$

CASE 3. Suppose that  $x \leq U_{s,t}$  and  $X_t^{(\hat{N}_{s,t}+1)} \leq U_{t,u}$ . The particle starting from  $x$  at time  $s$  visits 0 during both intervals  $[s, t]$  and  $[t, u]$ . Then,  $t \leq T \leq u$ , and so  $\hat{N}_{t,u} \leq \hat{N}_{s,t}$ , whence, using  $(**)$  again,  $\hat{N}_{t,u} = \hat{N}_{s,u}$ . Thus,

$$Z_{t,u}(Z_{s,t}(x)) = Z_{t,u}(X_t^{(\hat{N}_{s,t}+1)}) = X_u^{(\hat{N}_{t,u}+1)} = X_u^{(\hat{N}_{s,u}+1)}.$$

CASE 4. Suppose that  $x > U_{s,t}$  and  $W_t - W_s + x \leq U_{t,u}$ , so that the particle starting from  $x$  at time  $s$  visits 0 during the interval  $[t, u]$  but not during  $[s, t]$ . Combining these inequalities with that denoted by  $(*)$  above,  $X_t^{(\hat{N}_{s,t}+1)} \leq U_{t,u}$ , so  $t \leq T \leq u$  and we may argue as for the preceding case that  $\hat{N}_{t,u} = \hat{N}_{s,u}$ . Thus

$$Z_{t,u}(Z_{s,t}(x)) = Z_{t,u}(W_t - W_s + x) = X_u^{(\hat{N}_{t,u}+1)} = X_u^{(\hat{N}_{s,u}+1)}.$$

Finally note that  $x > U_{s,u}$  if and only if CASE 1 holds, and so we have demonstrated the required composition property.

Construct  $X$  as in the preceding lemma from two independent copies  $X(1)$  and  $X(2)$  taking the splicing time  $t_0$  to be equal to  $s$ , then we commented following Lemma 3 that

$$X_t^{(\hat{N}_{s,t}+1)} = X_{t-s}^{(1)}(2),$$

and it is also the case that

$$U_{s,t} = \sup_{0 \leq h \leq t-s} (-W_h(2)) = U_{0,t-s}(2).$$

By virtue of Lemma 2 we can now deduce equation  $(*)$  above since

$$X_{t-s}^{(1)}(2) \leq X_{t-s}^{(0)}(2) = W_t - W_s + U_{s,t}.$$

We can also prove that  $Z$  has stationary, independent increments from this splicing. The map  $Z_{s,t}$  can be written in terms of  $X(2)$  as

$$Z_{s,t}(x) = \begin{cases} X_{t-s}^{(1)}(2) & \text{for } 0 \leq x \leq U_{0,t-s}(2), \\ W_{t-s}(2) + x & \text{for } U_{0,t-s}(2) < x. \end{cases}$$

Thus  $Z_{s,t}$  is determined from  $X(2)$  in the same way that  $Z_{0,t-s}$  is determined from  $X$  and so has the same law as  $Z_{0,t-s}$ . The construction also shows that  $Z_{s,t}$  is independent of  $(Z_{u,v}; 0 \leq u \leq v \leq s)$ , this being determined by  $X(1)$ .

Finally note that  $t \mapsto X_{t-s}^{(1)}(2)$  is for  $t \geq s$  a sticky Brownian motion starting from zero and this is the same as  $t \mapsto Z_{s,t}(0)$ . More generally

$$Z_{s,t}(x) = \begin{cases} x + W_t - W_s & s \leq t \leq \tau, \\ Z_{s,t}(0) & t \geq \tau \end{cases}$$

where  $\tau = \inf\{t \geq s : x + W_t - W_s \leq 0\}$ , from which we may easily deduce that  $t \mapsto Z_{s,t}(x)$  is, for  $t \geq s$ , a sticky Brownian motion starting at time  $s$  from  $x$ , and that this trajectory satisfies the SDE as claimed.  $\square$

**Lemma 5.** *We may recover the process  $X$  from the flow: for each  $t$ ,*

$$\{Z_{s,t}(0) : 0 \leq s < t\} = \{X_t^{(k)} : k \geq 1\}.$$

*Proof.* To see this note first that the inclusion

$$\{Z_{s,t}(0) : 0 \leq s < t\} \subseteq \{X_t^{(k)} : k \geq 1\}.$$

follows from the very definition of  $Z_{s,t}(0)$  as  $X_t^{(k)}$  for some  $k$ . This inclusion is in fact an equality as the following argument shows. For each  $k \geq 1$  let

$$g_t^{(k)} = \sup\{s \leq t : X_t^{(k)} = 0\},$$

then  $X_h^{(k)} = W_h - W_{g_t^{(k)}}$  for  $h \in [g_t^{(k)}, t]$ . Now, using the fact that the  $X_t^{(k)}$  take distinct values for different  $k$  unless equal to zero, we see that:

$$\hat{N}_{g_t^{(k)}, t} = \inf\{r : X_h^{(r+1)} = 0 \text{ for some } h \in [g_t^{(k)}, t]\} = k - 1,$$

unless  $g_t^{(k)} = t$  and  $X_t^{(k)} = 0$ . It follows that  $Z_{g_t^{(k)}, t}(0) = X_t^{(k)}$  for all  $k$  such that  $X_t^{(k)} > 0$ .

Finally note that if, for some  $k$  we have  $X_t^{(k)} = 0$  then  $Z_{s,t}(0) = 0$  also for  $s$  sufficiently close to  $t$ . Thus regardless of the value of  $X_t^{(k)}$  there is always some  $s < t$  with  $Z_{s,t}(0) = X_t^{(k)}$ .  $\square$

The maps  $Z_{s,t}$  were expected to have a simple form on the basis of the verbal description given in the introduction. For an initial position  $x$  sufficiently large the motion is simply a ‘translate’ of  $W$ . Whereas for all values  $x$  less than or equal to some critical level, the corresponding particles have reached zero between times  $s$  and  $t$  and coalesced. We can identify the map  $Z_{s,t}$  with a point  $(X_{s,t}, U_{s,t}, V_{s,t}) \in \mathbb{R}_+^3$  where

$$U_{s,t} = \sup_{s \leq h \leq t} (W_s - W_h), \tag{2.1a}$$

$$V_{s,t} = \sup_{s \leq h \leq t} (W_t - W_h), \tag{2.1b}$$

$$X_{s,t} = Z_{s,t}(0). \tag{2.1c}$$

Notice that, equation (\*) from the proof of Theorem 4 can be written

$$X_{s,t} \leq V_{s,t}, \tag{2.2}$$

and in fact this holds simultaneously for all  $s \leq t$  with probability one. When we compose such maps we obtain via this identification a semigroup on  $\{(x, u, v) \in \mathbb{R}_+^3 : x \leq v\}$ ; the map  $z$  associated with  $(x, u, v)$  being

$$z(\xi) = \begin{cases} x & \text{if } \xi \leq u, \\ v - u + \xi & \text{if } \xi > u. \end{cases} \quad (2.3)$$

Thus when we compose the maps associated with  $z_1 = (x_1, u_1, v_1)$  and  $z_2 = (x_2, u_2, v_2)$  we obtain the map associated with  $z_3 = (x_3, u_3, v_3) = z_2 \circ z_1$  where

$$u_3 = u_1 + (u_2 - v_1)^+ \quad (2.4a)$$

$$v_3 = v_2 + (v_1 - u_2)^+ \quad (2.4b)$$

$$x_3 = x_2 1_{(u_2 \geq x_1)} + (x_1 + v_2 - u_2) 1_{(u_2 < x_1)}. \quad (2.4c)$$

Using this identification of each map  $Z_{s,t}$  with a point in  $\mathbb{R}_+^3$  we can investigate the continuity of  $(s, t) \mapsto Z_{s,t}$ . Discontinuities are caused by jumps in  $X_{s,t}$  which in turn are caused by jumps in  $\hat{N}_{s,t}$ . For a pair of times  $s_0 < t_0$  let  $k = \hat{N}_{s_0, t_0} + 1$ . Then if  $r < k$ ,  $X_h^{(r)} > 0$  for all  $h \in [s_0, t_0]$ , and in fact by continuity all  $h$  in some larger open interval containing  $[s_0, t_0]$ . If there exists  $h_0 \in (s_0, t_0)$  with  $X_{h_0}^{(k)} = 0$  then  $\hat{N}_{s,t} = \hat{N}_{s_0, t_0}$  for all pairs  $s, t$  sufficiently close to  $s_0, t_0$ . The alternative is that the only  $h \in [s_0, t_0]$  such that  $X_h^{(k)} = 0$  are  $h = s_0$  or  $h = t_0$  or both. In these cases  $(s, t) \mapsto \hat{N}_{s,t}$  has a simple discontinuity at  $(s_0, t_0)$ . However not all such discontinuities in  $\hat{N}$  cause discontinuities in  $(s, t) \mapsto X_{s,t} = Z_{s,t}(0)$ . It is only if the only  $h \in [s_0, t_0]$  such that  $X_h^{(\hat{N}_{s_0, t_0} + 1)} = 0$  is  $h = s_0$  that  $X_{s,t}$  has a point of discontinuity at  $(s_0, t_0)$  which is of the form:

$$\lim_{h, k \rightarrow 0, h \geq 0} X_{s-h, t+k} \text{ exists and equals } X_{s,t} \quad (2.5a)$$

$$\lim_{h, k \rightarrow 0, h > 0} X_{s+h, t+k} \text{ exists.} \quad (2.5b)$$

Another place to worry about the continuity of  $(s, t) \mapsto X_{s,t}$  is on the diagonal  $s = t$ . Here the behaviour of  $(s, t) \mapsto \hat{N}_{s,t}$  is particularly bad. But the inequality  $X_{s,t} \leq V_{s,t}$  assuages our concerns.

### 3 LOOKING AT THE FLOW BACKWARDS

In this section we study the flow  $(Z_{s,t}; s \leq t)$  by fixing  $t$  and letting  $s$  decrease. This reveals certain random variables which have a Poisson distribution and leads to a verification of the description of the  $\mathbb{R}^\infty$ -valued process  $X$  in terms of marked trees which was given in the introduction. Recently Watanabe [19] has also made a similar study of the flow  $Z$  in which the dual flow is described in terms of elastic Brownian motions. Let us define, for  $s < t$ , a random set  $\Xi_{s,t}$  via

$$\Xi_{s,t} = \{Z_{h,t}(0) : s \leq h < t\}. \quad (3.1)$$

Recall from Theorem 4 or Lemma 5 that the flow  $Z$  is constructed from the  $\mathbb{R}^\infty$ -valued process  $X$  in such a way that  $Z_{s,t}(0)$  is equal to  $X_t^{(k)}$  for a certain random  $k$ . Thus

$$\Xi_{s,t} \subseteq \{X_t^{(k)} : k = 1, 2, \dots\}, \quad (3.2)$$

and in view of Lemma 2 we deduce that except for an exceptional set of  $t$  having Lebesgue measure zero  $\Xi_{s,t}$  contains only a finite number of points. From this, and the stationarity of  $Z$ , it follows that for any fixed  $t$ , the probability that  $\Xi_{s,t}$  is finite for all  $s$  is one.

The principal result of this section is the description of the law of  $\Xi_{s,t}$  conditional on the driving Brownian motion  $(W_t; t \geq 0)$ .

**Theorem 6.** *For fixed  $s < t$ , the law of  $\Xi_{s,t}$  conditional on  $(W_t; t \geq 0)$  is given, by the following.*

- $0 \in \Xi_{s,t}$  with probability one;
- $\Xi_{s,t} \setminus \{0\}$  is distributed as a Poisson point process of intensity  $2\theta$  on  $(0, V_{s,t}]$ .

In proving this theorem it is helpful to extend the flow to negative time. Thus suppose  $(Z_{s,t}; -\infty < s \leq t < \infty)$  is a flow defined for all time with stationary independent increments, and whose restriction to positive time is the flow constructed in Theorem 4. Suppose also that  $Z$  possess the regularity properties (2.5). As in the introduction let  $\mathcal{F}_{s,t}$  be the  $\sigma$ -algebra generated by  $(Z_{u,v}; s \leq u \leq v \leq t)$  which we complete to contain all events with zero probability.

**Lemma 7.** *Suppose that  $T$  is a finite stopping time with respect to the filtration  $(\mathcal{F}_{-t,0})_{t \geq 0}$  then*

$$(Z_{s-T,t-T}; -\infty < s \leq t \leq 0) \stackrel{\text{law}}{=} (Z_{s,t}; -\infty < s \leq t \leq 0),$$

and is moreover independent of  $\mathcal{F}_{-T,0}$ .

*Proof.* We mimic the usual proof of the strong Markov property. Suppose the stopping time  $T$  only takes values in a discrete set of times, then, by virtue of the stationarity of the flow and the fact that  $\mathcal{F}_{-\infty,-t_i}$  and  $\mathcal{F}_{-t_i,0}$  are independent, the result holds for  $T$ . An arbitrary stopping time is the limit of a decreasing sequence of such elementary stopping times. Using this fact we deduce the result for an arbitrary finite  $T$  noting that as  $h \downarrow 0$ , for any  $s \leq t \leq 0$ ,

$$Z_{s-(T+h),t-(T+h)} \rightarrow Z_{s-T,t-T},$$

by equation (2.5). □

For a flow which is defined for all negative time we extend the driving Brownian motion  $W$  to negative time by the device of considering the doubly indexed process  $(W_{s,t}; -\infty < s \leq t < \infty)$  defined by

$$W_{s,t} = Z_{s,t}(x) \text{ for sufficiently large } x.$$

For non-negative  $s$  and  $t$  this is just the increment  $W_t - W_s$  of the Brownian motion  $W$ . The definitions of  $U_{s,t}$  and  $V_{s,t}$  can be extended to negative times via:

$$U_{s,t} = \sup_{h \in [s,t]} (-W_{s,h}) \text{ and } V_{s,t} = \sup_{h \in [s,t]} (W_{h,t}) \quad (3.3)$$

We may also define a random set  $\Xi_{-\infty,0}$  to be  $\{Z_{s,0}(0) : -\infty < s < 0\}$ .

**Lemma 8.**  *$\Xi_{-\infty,0} \setminus \{0\}$  is a Poisson point process on  $(0, \infty)$  with intensity  $2\theta$  and is independent of  $(W_{s,t}; -\infty < s \leq t < \infty)$ .*

*Proof.* For  $x \geq 0$  let  $-T_x = \sup\{s \leq 0 : V_{s,0} > x\}$ . Then  $x \mapsto T_x$  is right-continuous, and each  $T_x$  is a  $(\mathcal{F}_{-t,0})_{t \geq 0}$ -stopping time by virtue of  $\mathcal{F}_{-t,0}$  being complete.

We define a right continuous counting process  $(N(x); x \geq 0)$  via

$$N(x) = n(\Xi_{-T_x,0} \setminus \{0\}),$$

where  $nS$  counts the number of points belonging to the set  $S$ .

Now fix some  $x$  and notice that  $V_{-T_x,0} = x$ . If  $-T_x \leq s < 0$  then

$$Z_{s,0}(0) \leq V_{s,0} \leq V_{-T_x,0} = x.$$

On the other hand if  $s < -T_x$  then  $Z_{s,0}(0) = Z_{-T_x,0} \circ Z_{s,-T_x}(0)$ , which is either equal to  $Z_{-T_x,0}(0)$  or strictly greater than  $x$ . From this we deduce two facts. Firstly that  $(N(h); h \leq x)$  is determined by the set  $\{Z_{s,0}(0) : -T_x \leq s < 0\}$  and hence is measurable with respect to  $\mathcal{F}_{-T_x,0}$ . Secondly that for  $y \geq 0$ ,

$$\begin{aligned} N(x+y) - N(x) &= n\{Z_{s,0}(0) \in (x, x+y] : s < 0\} \\ &= n\{Z_{s,-T_x}(0) \in (0, y] : s < T_x\}. \end{aligned}$$

On applying the previous lemma we deduce that  $N$  is a process with independent identically distributed increments, and since it is constant except for positive jumps of size one, it must be a Poisson process of with some rate that remains to determine.

Before discussing the rate of  $N$  we first prove that it is independent of  $(W_{s,t}; -\infty < s \leq t < \infty)$ . In fact because of the independence of flow before and after time 0 it is enough to show that  $N$  is independent of  $(W_{s,t}; -\infty < s \leq t \leq 0)$ . It follows from Lemma 2 and stationarity that for each fixed  $s < 0$  with probability one, the strict inequality holds

$$Z_{s,0}(0) < V_{s,0}.$$

Nevertheless there exist random  $s$  at which equality holds: the times  $-T_x$  for  $x \in \Xi_{-\infty,0}$  in fact. By Fubini's Theorem, with probability one, the Lebesgue measure of the set of these exceptional  $s$  for which equality holds is zero. Now consider an interval of the form  $[-T_x, -T_{x-}]$  for which  $T_x \neq T_{x-} = \lim_{\epsilon \downarrow 0} T_{x-\epsilon}$ . For  $s$  belonging to such an interval

$$Z_{s,0}(0) = Z_{-T_x,0} \circ Z_{s,-T_x}(0) = Z_{-T_{x-},0}(0).$$

This latter value cannot be equal to  $x$ , for if it was then  $Z_{s,0}(0) = x = V_{s,0}$  for all  $s$  in the interval and the interval would then have to have length zero. Notice that if  $s < -T_x$  or  $-T_{x-} < s$  then  $Z_{s,0}(0)$  cannot be equal to 0 either, and so we have shown that any  $x$  such that  $T_{x-} < T_x$  cannot belong to  $\Xi_{-\infty,0}$  and is not the time of a jump of the process  $(N(x); x \geq 0)$ . But  $x$  such that  $T_{x-} < T_x$  are exactly the local times at which the reflecting Brownian motion  $(W_{-t,0} + V_{-t,0}; t \geq 0)$  makes excursions from zero. It follows from a well-known property of Poisson processes (see [9], Chapter XII, Proposition 1.7) that  $N$  is independent of the Poisson point process of excursions from zero made by this reflecting Brownian motion and hence from the reflecting Brownian motion itself. Finally we note that we can recover  $(W_{s,t}; -\infty < s \leq t \leq 0)$  from this reflecting Brownian motion and the proof of independence is complete.

To determine the rate of  $N$  we may calculate, using Lemma 2

$$\begin{aligned} \int_{-\infty}^0 e^{\frac{1}{2}\lambda^2 s} \mathbb{P}(Z_{s,0}(0) = 0) ds &= \int_0^{\infty} e^{-\frac{1}{2}\lambda^2 t} \mathbb{P}(Z_{0,t}(0) = 0) dt = \\ &= \int_0^{\infty} e^{-\frac{1}{2}\lambda^2 t} \mathbb{P}(X_t^{(1)} = 0) dt = \int_0^{\infty} e^{-\frac{1}{2}\lambda^2 t} dA_t^{(1)} = \frac{2}{\lambda(2\theta + \lambda)}. \end{aligned}$$

On the other hand,  $Z_{s,0}(0) = 0$  if and only if  $N(V_{s,0}) = 0$ , and so if  $N$  has rate  $\beta$  then

$$\begin{aligned} \int_{-\infty}^0 e^{\frac{1}{2}\lambda^2 s} \mathbb{P}(Z_{s,0}(0) = 0) ds &= \int_{-\infty}^0 e^{\frac{1}{2}\lambda^2 s} \mathbb{E}[e^{-\beta V_{s,0}}] ds = \\ &= \int_0^{\infty} e^{-\beta y} \int_0^{\infty} e^{-\frac{1}{2}\lambda^2 t} \mathbb{P}(|B_t| \in dy) dt = 2 \int_0^{\infty} e^{-\beta y} \frac{e^{-\lambda y}}{\lambda} dy = \frac{2}{\lambda(\beta + \lambda)}, \end{aligned}$$



where we have used the resolvent of Brownian motion and the fact that  $V_{0,t}$  has the same law as  $|B_t|$ . This shows that  $\beta = 2\theta$ .  $\square$

*Proof of Theorem 6.* We have  $0 \in \Xi_{s,t}$  unless  $\hat{N}_t = \infty$  which according to Lemma 2 occurs only for a set of  $t$  having Lebesgue measure zero. Thus by stationarity for any fixed  $s$  and  $t$  the probability that  $0 \in \Xi_{s,t}$  is one. By Fubini's Theorem it follows  $0 \in \Xi_{s,t}$  with probability one having conditioned on  $W$  (except for a null set of possible values for  $W$ .)

We consider the extended flow defined for negative time. By stationarity and Lemma 8

$$\Xi_{-\infty,t} \setminus \{0\} = \{Z_{s,t}(0) : -\infty < s < t\} \setminus \{0\}$$

is distributed as a Poisson point process with rate  $2\theta$  on  $(0, \infty)$  and is independent of the white noise  $(W_{s,t}; s \leq t)$  and hence of the Brownian motion  $(W_t; t \geq 0) = (W_{0,t}; t \geq 0)$ . Finally

$$\Xi_{s,t} = \Xi_{-\infty,t} \cap [0, V_{s,t}],$$

and  $V_{s,t}$  is measurable with respect to  $W$ .  $\square$

**Corollary 9.** *The law of  $Z_{s,t}$  depends only on  $h = t - s$  and is given by*

$$(Z_{s,t}(0), U_{s,t}, V_{s,t}) \stackrel{\text{law}}{=} ((UR_h - T)^+, (1 - U)R_h, UR_h) \quad (3.4)$$

where  $U$  is a variable uniformly distributed on  $[0, 1]$ ,  $T$  is an exponential random variable with mean  $1/(2\theta)$ ,  $R_h$  is distributed as  $\sqrt{h}R_1$  for  $R_1$  the modulus of a standard Gaussian variable in  $\mathbb{R}^3$ , and the three variables  $U$ ,  $T$  and  $R_h$  are independent.

*Proof.* The marginal law of  $(U_{s,t}, V_{s,t})$  which depends just on the BM  $W$  is very well known see for example [20]. The exponentially distributed contribution describing the conditional law of  $Z_{s,t}(0)$  is a consequence of Theorem 6 above, since  $Z_{s,t}(0) > x$  if and only if  $\Xi_{s,t}$  contains a point in  $(x, V_{s,t}]$ .  $\square$

We turn now to an examination of the joint law of  $\Xi_{0,t_1}$  and  $\Xi_{0,t_2}$  when  $t_1 < t_2$ . It is at this point that the tree structure discussed in the introduction becomes evident.

**Proposition 10.** *For fixed  $t_1 < t_2$ , the joint law of  $\Xi_{0,t_1} \setminus \{0\}$  and  $\Xi_{0,t_2} \setminus \{0\}$  conditional on  $(W_t; t \geq 0)$  is given by the following.*

1. *If  $V_{0,t_1} \leq U_{t_1,t_2}$  then  $\Xi_{0,t_1} \setminus \{0\}$  and  $\Xi_{0,t_2} \setminus \{0\}$  are independent Poisson point processes of intensity  $2\theta$  on  $(0, V_{0,t_1}]$  and  $(0, V_{0,t_2}]$  respectively.*
2. *If  $V_{0,t_1} > U_{t_1,t_2}$  then  $\Xi_{0,t_2}$  is composed of two pieces, one derived from  $\Xi_{0,t_1}$ , and the other independent of it:*
  - (a) *the restriction of  $\Xi_{0,t_2}$  to  $(V_{t_1,t_2}, V_{0,t_2}]$  is equal to the translation of the restriction of  $\Xi_{0,t_1}$  to  $(U_{t_1,t_2}, V_{0,t_1}]$  by  $W_{t_2} - W_{t_1}$ , that is,*

$$\Xi_{0,t_2}|_{(V_{t_1,t_2}, V_{0,t_2})} = (W_{t_2} - W_{t_1}) + \Xi_{0,t_1}|_{(U_{t_1,t_2}, V_{0,t_1})};$$

- (b) *the restriction of  $\Xi_{0,t_2} \setminus \{0\}$  to  $(0, V_{t_1,t_2}]$  and  $\Xi_{0,t_1} \setminus \{0\}$  are independent Poisson point processes of intensity  $2\theta$  on  $(0, V_{t_1,t_2}]$  and  $(0, V_{0,t_1}]$  respectively.*

*Proof.* First observe that  $\Xi_{0,t_1} \setminus \{0\}$  and  $\Xi_{t_1,t_2} \setminus \{0\}$  are by virtue of Theorem 6, conditionally on  $W$ , distributed as Poisson point processes of rate  $2\theta$  on  $(0, V_{0,t_1}]$  and  $(0, V_{t_1,t_2}]$  respectively. Moreover they are conditionally independent since they are measurable with respect to the independent  $\sigma$ -algebras:  $\sigma(Z_{u,v}; 0 \leq u \leq v \leq t_1)$  and  $\sigma(Z_{u,v}; t_1 \leq u \leq v \leq t_2)$  respectively.

Consider  $s < t_1$  and suppose first that  $Z_{s,t_1}(0) > U_{t_1,t_2}$  then

$$Z_{s,t_2}(0) = Z_{t_1,t_2}(Z_{s,t_1}(0)) = Z_{s,t_1}(0) + W_{t_2} - W_{t_1}.$$

On the other hand if  $Z_{s,t_1}(0) \leq U_{t_1,t_2}$  then

$$Z_{s,t_2}(0) = Z_{t_1,t_2}(Z_{s,t_1}(0)) = Z_{t_1,t_2}(0).$$

Now suppose  $V_{0,t_1} \leq U_{t_1,t_2}$  then since for  $0 \leq s < t_1$ ,  $Z_{s,t_1}(0) \leq V_{0,t_1}$ , we have  $Z_{s,t_2}(0) = Z_{t_1,t_2}(0)$ . Thus

$$\begin{aligned} \Xi_{0,t_2} \setminus \{0\} &= \{Z_{s,t_2}(0) : 0 \leq s < t_2\} \setminus \{0\} \\ &= \{Z_{s,t_2}(0) : t_1 \leq s < t_2\} \setminus \{0\} = \Xi_{t_1,t_2} \setminus \{0\}. \end{aligned}$$

In view of our opening observation that  $\Xi_{0,t_1}$  and  $\Xi_{t_1,t_2}$  are independent, statement 1 of the proposition is proved.

Suppose instead that  $V_{0,t_1} > U_{t_1,t_2}$  then we have

$$\begin{aligned} \Xi_{0,t_2} \setminus \{0\} &= \{Z_{s,t_2}(0) : 0 \leq s < t_2\} \setminus \{0\} \\ &= \{Z_{s,t_2}(0) : 0 \leq s < t_1 \text{ such that } Z_{s,t_1}(0) > U_{t_1,t_2}\} \cup \{Z_{s,t_2}(0) : t_1 \leq s < t_2\} \setminus \{0\} \\ &= \{Z_{s,t_1}(0) + W_{t_2} - W_{t_1} : 0 \leq s < t_1 \text{ such that } Z_{s,t_1}(0) > U_{t_1,t_2}\} \cup \Xi_{t_1,t_2} \setminus \{0\}, \end{aligned}$$

where the first set is supported on  $(V_{t_1,t_2}, V_{0,t_2}] = (U_{t_1,t_2} + W_{t_2} - W_{t_1}, V_{0,t_1} + W_{t_2} - W_{t_1}]$  and the second set is supported on  $(0, V_{t_1,t_2}]$ . Since  $\Xi_{0,t_1}$  and  $\Xi_{t_1,t_2}$  are independent, statement 2 of the proposition is proved.  $\square$

This proposition can be easily extended (by induction on  $n$ ) to a similar statement concerning the joint law of  $\Xi_{0,t_1}, \Xi_{0,t_2}, \dots, \Xi_{0,t_n}$  when  $t_1 < t_2 < \dots < t_n$ : each  $\Xi_{0,t_k}$  is made from two pieces, one derived from  $\Xi_{0,t_{k-1}}$ , which may be empty, and one from  $\Xi_{t_{k-1},t_k}$  which is independent of  $\Xi_{0,t_1}, \Xi_{0,t_2}, \dots, \Xi_{0,t_{k-1}}$ .

We will now re-interpret Proposition 10 in terms of the marked trees as described in the introduction. Recall that the excursions of the reflecting Brownian motion  $\xi$  from zero determine a family of rooted trees to which we add marks according to a Poisson point process of intensity  $2\theta$ . The random set  $\Xi_{0,t} \setminus \{0\}$  determines the position of the marks on the path corresponding to time  $t$ . Recall that this path has length  $\xi_t = W_t + \sup_{s \leq t} (-W_s)$ . If  $x \in \Xi_{0,t} \setminus \{0\}$  then there is a mark on the path at a distance  $\xi_t - x$  from the root. Part 1 of the proposition covers the case when the two times  $t_1$  and  $t_2$  are straddled by different excursions of  $\xi$ , and the paths corresponding to the times are disjoint. Part 2 of the proposition covers the case when the two times  $t_1$  and  $t_2$  are straddled by a single excursion of  $\xi$ . In this case the two corresponding paths co-incide from the root for a distance  $I_{t_1,t_2} = \inf_{h \in [t_1,t_2]} \xi_h$  and part 2(a) of the proposition guarantees that  $\Xi_{0,t_1}$  and  $\Xi_{0,t_2}$  determine marks in the same places along this common part of the paths. Then after a distance  $I_{t_1,t_2}$  from the root the paths separate and the location of marks on the two paths is independent as covered by part 2(b) of the proposition. Note that equation (1.5) in which the positions of the marks were determined from the  $\mathbb{R}^\infty$ -valued process  $X$  is equivalent to the rule we have just given for determining the positions from the sets  $\Xi_{0,t}$  by virtue of Lemma 5 with  $X_t^{(k)}$  determining the position of  $k$ th mark along the path counting from the origin.

We can re-cast these results in terms of a snake process of the type considered by Le Gall [8], with the motion process being a Poisson counting process. As a consequence the SDEs displayed in the statement of Theorem 1 specify the generator of the this snake process. Compare with the work of Dhersin and Serlet [4].

For each  $t$  we construct a path  $N_t : [0, \infty) \mapsto \mathbb{Z}^+ \cup \{\partial\}$  via

$$N_t(y) = \begin{cases} \partial & \text{if } y \geq \xi_t, \\ n(\Xi_{0,t} \cap [\xi_t, \xi_t - y)) & \text{if } y < \xi_t, \end{cases} \quad (3.5)$$

where  $nS$  counts the number of elements of a set  $S$ . The following proposition is a consequence of Proposition 10.

**Proposition 11.** *( $N_t; t \geq 0$ ) is a path-valued Markov process such that conditional on the lifetime process ( $\xi_t; t \geq 0$ ):*

- $y \mapsto N_t(y)$  is a Poisson counting process with rate  $2\theta$  on  $[0, \xi_t)$  for each  $t$ ;
- $N_{t_1}(y) = N_{t_2}(y)$  on  $[0, I_{t_1, t_2})$  for all  $t_1 < t_2$ ;
- the Poisson processes  $y \mapsto N_{t_2}(I_{t_1, t_2} + y) - N_{t_2}(I_{t_1, t_2})$  on  $[0, \xi_{t_2} - I_{t_1, t_2})$  and  $y \mapsto N_{t_1}(y)$  on  $[0, \xi_{t_1})$  are independent for all  $t_1 < t_2$ .

## 4 THE FLOW AS A NOISE

A probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  together with a doubly indexed family of sub- $\sigma$ -algebras  $(\mathcal{F}_{s,t}; -\infty \leq s \leq t \leq \infty)$  and a group of measure preserving shifts  $(\theta_t; t \in \mathbb{R})$  satisfying:

- For any  $h$  and any  $s < t$ , we have  $\theta_h^{-1} \mathcal{F}_{s,t} = \mathcal{F}_{s+h, t+h}$ .
- For any  $s < t$  and  $u < v$  with the intervals  $(s, t)$  and  $(u, v)$  disjoint, the two  $\sigma$ -algebras  $\mathcal{F}_{s,t}$  and  $\mathcal{F}_{u,v}$  are independent.
- For any  $s < t < u$  the  $\sigma$ -algebra  $\mathcal{F}_{s,u}$  is generated by  $\mathcal{F}_{s,t}$  and  $\mathcal{F}_{t,u}$ .

is called a *noise* by Tsirelson, [11] and [10]. The concept has been around in various forms for some time, for example see Feldman [5]. The most familiar example is that of white noise—obtained by considering the increments of a Brownian motion. Any stochastic flow with independent stationary increments gives rise to such an object, but often the corresponding noise is going to be the white noise associated with some driving Brownian motion.

We now consider the flow of sticky Brownian motions we have constructed as a noise. Recall that for each  $s \leq t$  the  $\sigma$ -algebra  $\mathcal{F}_{s,t}$  is generated by the random maps  $Z_{u,v}$  for  $s \leq u \leq v \leq t$ . We refer to a process  $B$  as a  $\mathcal{F}_{s,t}$ -Brownian motion if it is a Brownian motion in law, and if for all  $0 \leq s \leq t$  the increment  $B_t - B_s$  is measurable with respect to  $\mathcal{F}_{s,t}$ . Of course the driving Brownian motion  $W$  is an  $\mathcal{F}_{s,t}$ -Brownian motion. Now we define the linear component of the noise to be the family of  $\mathbb{P}$ -complete  $\sigma$ -algebras, denoted by  $(\mathcal{F}_{s,t}^{lin}, s \leq t)$ , generated by the increments of all  $\mathcal{F}_{s,t}$ -Brownian motions, thus

$$\mathcal{F}_{s,t}^{lin} = \sigma(B_v - B_u; s \leq u \leq v \leq t, B \text{ is a } \mathcal{F}_{s,t} \text{-Brownian motion}). \quad (4.1)$$

In particular  $\mathcal{F}_{s,t}^{lin}$  contains the  $\sigma$ -algebra  $\mathcal{F}_{s,t}^W$  generated by increments of the driving Brownian motion  $W$  between times  $s$  and  $t$ .

**Theorem 12.** For every  $0 \leq s \leq t$ ,

$$\mathcal{F}_{s,t}^{lin} = \mathcal{F}_{s,t}^W.$$

*Proof.* Suppose  $B$  is an  $\mathcal{F}_{s,t}$ -Brownian motion. Fix  $s < t$ , and then  $\mathcal{F}_{s,t}^W$  being contained in  $\mathcal{F}_{s,t}$  is independent of  $\mathcal{F}_{0,s}$  which contains  $\sigma(B_s)$  and  $\mathcal{F}_{0,s}^W$ . Thus

$$\mathbb{E}[B_s | \mathcal{F}_{0,t}^W] = \mathbb{E}[B_s | \mathcal{F}_{0,s}^W \vee \mathcal{F}_{s,t}^W] = \mathbb{E}[B_s | \mathcal{F}_{0,s}^W].$$

Similarly

$$\mathbb{E}[B_t - B_s | \mathcal{F}_{0,t}^W] = \mathbb{E}[B_t - B_s | \mathcal{F}_{0,s}^W \vee \mathcal{F}_{s,t}^W] = \mathbb{E}[B_t - B_s | \mathcal{F}_{s,t}^W].$$

It follows that

$$\mathbb{E}[B_t | \mathcal{F}_{0,t}^W] - \mathbb{E}[B_s | \mathcal{F}_{0,s}^W] = \mathbb{E}[B_t - B_s | \mathcal{F}_{s,t}^W].$$

Using this we deduce  $B_t - \mathbb{E}[B_t | \mathcal{F}_{0,t}^W]$  is an  $(\mathcal{F}_{0,t})_{t \geq 0}$ -martingale which is orthogonal to the Brownian motion  $W$ , and by virtue of the following proposition identically zero. But this means that the increment  $B_t - B_s$  is  $\mathcal{F}_{s,t}^W$ -measurable, and we are done.  $\square$

The following proposition is a consequence of the uniqueness result of Theorem 1 and the general theory of martingale representations, presented, for example, in Chapter V of [9].

**Proposition 13.**  $W$  has the  $\mathcal{F}_t$ -predictable representation property.

Thus we have obtained a noise with a nonlinear component: for we know from Theorem 6 that  $\mathcal{F}_{s,t}^W$  is strictly contained in  $\mathcal{F}_{s,t}$  and thus, in light of the above,  $\mathcal{F}_{s,t} \neq \mathcal{F}_{s,t}^{lin}$ .

This example of a noise also provides an affirmative answer to the question asked by Tsirelson in [10]: can a predictable noise generate a non-cosy filtration. This is because we know that the filtration  $(\mathcal{F}_{0,t}; t \geq 0)$  is non-cosy from [16].

Finally we note that this example of a noise is also time-asymmetric: in view of the martingale representation property just given all martingales in the filtration  $(\mathcal{F}_{0,t})_{t \geq 0}$  are continuous whereas in the reverse filtration  $(\mathcal{F}_{-t,0})_{t \geq 0}$  there are many discontinuous martingales.

## References

- [1] D.J. Aldous, *The continuum random tree III*. Annals of Probability. **21:1**, 248-289,(1993).
- [2] D.J. Aldous, *Recursive self-similarity for random trees, random triangulations and Brownian excursion*. Annals of Probability. **22:2**, 527-545,(1994).
- [3] D.J. Aldous and J. Pitman, *The standard additive coalescent*. Annals of Probability. **26:4**, 1703-1726,(1998).
- [4] J.S. Dhersin and L. Serlet, *A stochastic calculus approach for the Brownian snake*. Canadian Journal of Mathematics. **52:1**, 92-118,(2000).
- [5] J. Feldman, *Decomposable processes and continuous products of probability spaces*. Journal of Functional Analysis, **8**, 1-51,(1971).
- [6] W. Feller, *Generalized second order differential operators and their lateral conditions*. Illinois Journal of Mathematics **1**, 459-504 (1957).

- [7] F.B. Knight, *Essentials of Brownian motion and diffusion*. Volume 18 of Mathematical Surveys. American Mathematical Society, (1981).
- [8] J.F. Le Gall, *A class of path-valued Markov processes and its applications to superprocesses*. Probability Theory and Related Fields, **95:1**, 25-46, (1993).
- [9] D.Revuz and M.Yor, *Continuous martingales and Brownian motion*. Springer, Berlin, (1999).
- [10] B. Tsirelson, *Within and beyond the reach of Brownian innovation*. Documenta Mathematica, extra volume ICM 1998, III:311-320, (1998).
- [11] B. Tsirelson, *Unitary Brownian motions are linearizable*. MSRI preprint 1998-027, (1998).
- [12] B. Tsirelson, *Scaling limit, noise, stability*. Lecture Notes in Mathematics (to appear), (2002).
- [13] B. Tsirelson, *Triple points: From non-Brownian filtrations to harmonic measures*. Geom. Funct. Anal. **7**:1096-1142, (1997).
- [14] B. Tsirelson and A.V. Vershik, *Examples of nonlinear continuous tensor products of measure spaces and non-Fock factorizations*. Reviews in Mathematical Physics **10**:1, 81-145, (1998).
- [15] J. Warren, *Branching Processes, the Ray-Knight theorem, and sticky Brownian motion*. Sémin. de Probabilités XXXI, Lecture notes in Mathematics 1655, 1-15. Springer, (1997).
- [16] J. Warren, *On the joining of sticky Brownian motion*. Sémin. de Probabilités XXXIII Lecture notes in Mathematics 1709, 257-266. Springer, (1999).
- [17] S. Watanabe, *The existence of a multiple spider martingale in the natural filtration of a certain diffusion in the plane*. Sémin. de Probabilités XXXIII, Lecture notes in Mathematics 1709, 277-290. Springer, (1999).
- [18] S. Watanabe, *Killing operations in superdiffusions by Brownian snakes*. Trends in probability and related analysis, Taipei, 177-190. World Sci. Publishing, (1999).
- [19] S.Watanabe, *Stochastic flows in duality and noises*. Preprint, (2001).
- [20] M. Yor, *Some remarks on the joint law of Brownian motion and its supremum*. Sémin. de Probabilités XXXI, Lecture notes in Mathematics 1655, 306-314. Springer, (1997).